

Special topics in Quantitative Genetics using sommer

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The sommer package was developed to provide R users a powerful and reliable multivariate mixed model solver. The package is focused in problems of the type $p > n$ (more effects to estimate than observations) and its core algorithm is coded in C++ using the Armadillo library. This package allows the user to fit mixed models with the advantage of specifying the variance-covariance structure for the random effects, and specify heterogeneous variances, and obtain other parameters such as BLUPs, BLUEs, residuals, fitted values, variances for fixed and random effects, etc.

The purpose of this vignette is to show how to fit special models in quantitative genetics using the sommer package:

- 1) Partitioned model
- 2) UDU' decomposition
- 3) Mating designs
- 4) Dominance variance

1) Partitioned model

The partitioned model was popularized by () to show that marker effects can be obtained fitting a GBLUP model to reduce the computational burden and then recover them by creating some special matrices MM' for GBLUP and $M'(M'M)^{-1}$ to recover marker effects. Here we show a very easy example using the DT_cpdata:

```
library(sommer)
data("DT_cpdata")
DT <- DT_cpdata
M <- GT_cpdata

#####
# MARKER MODEL
#####
mix.marker <- mmer(color~1,
                  random=~Rowf+vs(M),
                  rcov=~units,data=DT,
                  verbose = FALSE)

## Version out of date. Please update sommer to the newest version using:
## install.packages('sommer') in a new session
## Use the 'date.warning' argument to disable the warning message.

me.marker <- mix.marker$U$`u:M`$color

#####
# PARTITIONED GBLUP MODEL
#####
```

```

MMT<-M%*%t(M) ## additive relationship matrix
MMTinv<-solve(MMT) ## inverse
MTMMTinv<-t(M)%*%MMTinv # M' %*% (M'M)-

mix.part <- mmer(color~1,
                 random=~Rowf+vs(id, Gu=MMT),
                 rcov=~units,data=DT,
                 verbose = FALSE)

```

```

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## Use the 'date.warning' argument to disable the warning message.

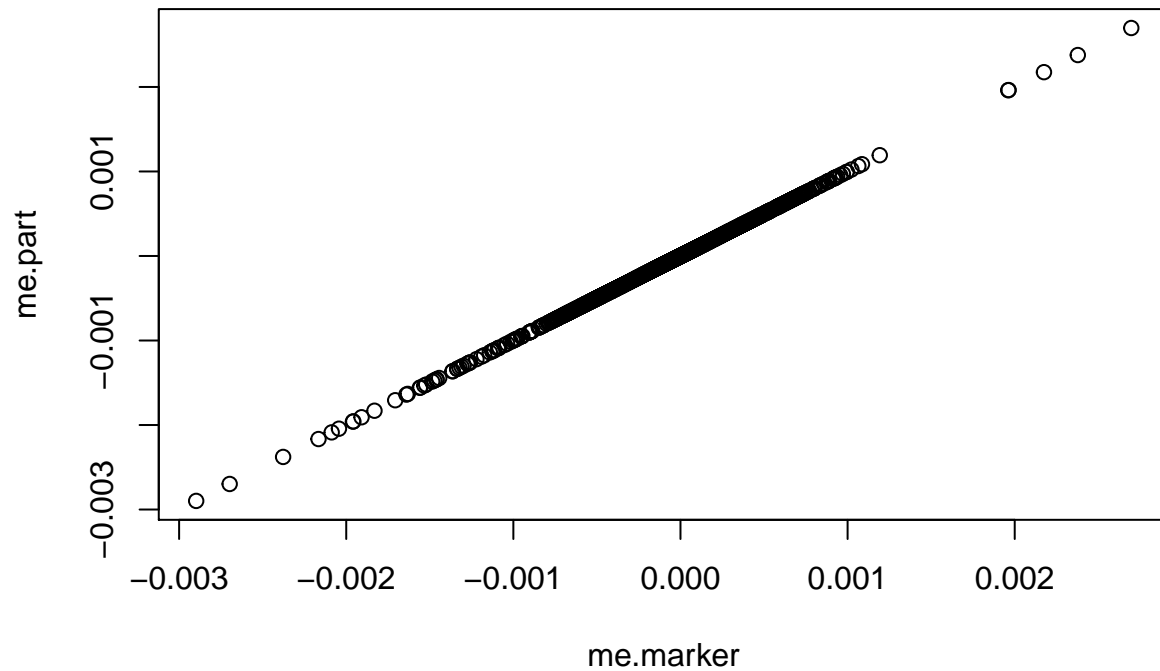
```

```

#convert BLUPs to marker effects me=M'(M'M)- u
me.part<-MTMMTinv%*%matrix(mix.part$U$`u:id`$color,ncol=1)

# compare marker effects between both models
plot(me.marker,me.part)

```



As can be seen this two models are equivalent with the exception that the partitioned model is more computationally efficient.

2) UDU' decomposition

Lee and Van der Warf (2015) proposed a decomposition of the relationship matrix $A=UDU'$ together with a transformation of the response and fixed effects $Uy = Ux + UZ + e$, to fit a model where the phenotypic variance matrix V is a diagonal because the relationship matrix is the diagonal matrix D from the decomposition that can be inverted easily and make multitrait models more feasible.

```

data("DT_wheat")
rownames(GT_wheat) <- rownames(DT_wheat)
G <- A.mat(GT_wheat)
Y <- data.frame(DT_wheat)

```

```

# make the decomposition
UD<-eigen(G) # get the decomposition: G = UDU'
U<-UD$vector
D<-diag(UD$values)# This will be our new 'relationship-matrix'
rownames(D) <- colnames(D) <- rownames(G)
X<-model.matrix(~1, data=Y) # here: only one fixed effect (intercept)
UX<-t(U)%*%X # premultiply X and y by U'
UY <- t(U) %*% as.matrix(Y) # multivariate

```

```

# dataset for decomposed model
DTd<-data.frame(id = rownames(G) ,UY, UX =UX[,1])
DTd$id<-as.character(DTd$id)

```

```

modeld <- mmer(cbind(X1,X2) ~ UX - 1,
              random = ~vs(id,Gu=D),
              rcov = ~vs(units),
              data=DTd, verbose = FALSE)

```

```

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```

```

# dataset for normal model
DTn<-data.frame(id = rownames(G) , DT_wheat)
DTn$id<-as.character(DTn$id)

```

```

modeln <- mmer(cbind(X1,X2) ~ 1,
              random = ~vs(id,Gu=G),
              rcov = ~vs(units),
              data=DTn, verbose = FALSE)

```

```

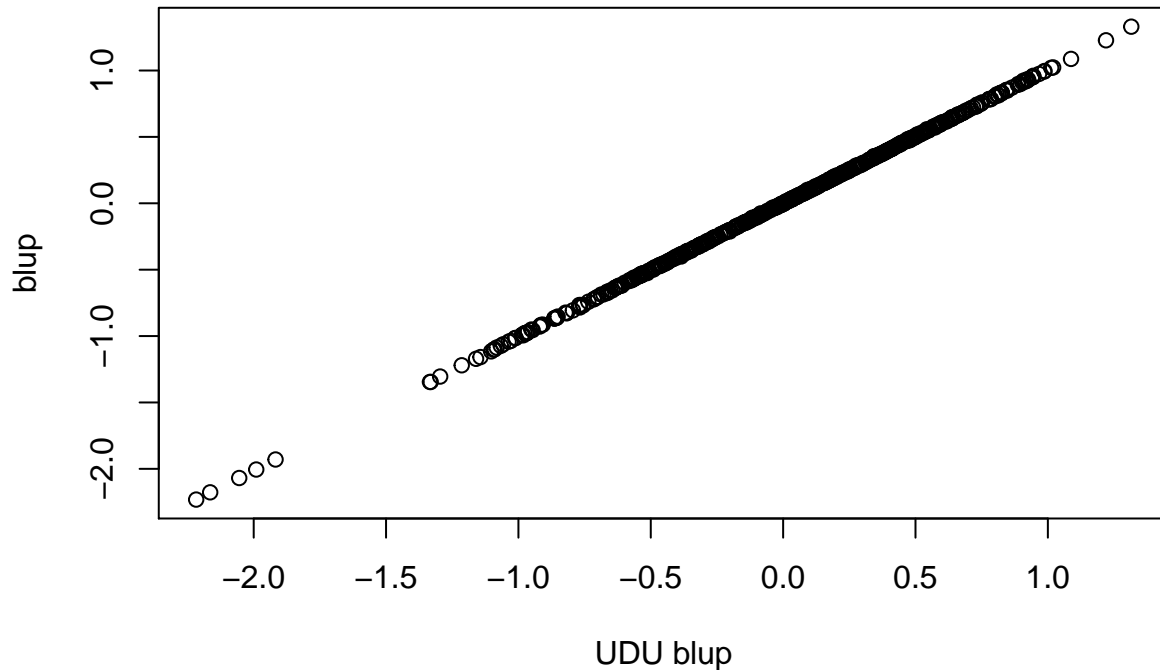
## Version out of date. Please update sommer to the newest version using:
## install.packages('sommer') in a new session
## Use the 'date.warning' argument to disable the warning message.

```

```

## compare regular and transformed blups
plot(x=(solve(t(U))%*%modeld$U$`u:id`$X2[colnames(D)]),
     y=modeln$U$`u:id`$X2[colnames(D)], xlab="UDU blup",
     ylab="blup")

```



As can be seen the two models are equivalent. Despite that sommer doesn't take a great advantage of this trick because it was built for dense matrices using the Armadillo library other software may be better using this trick.

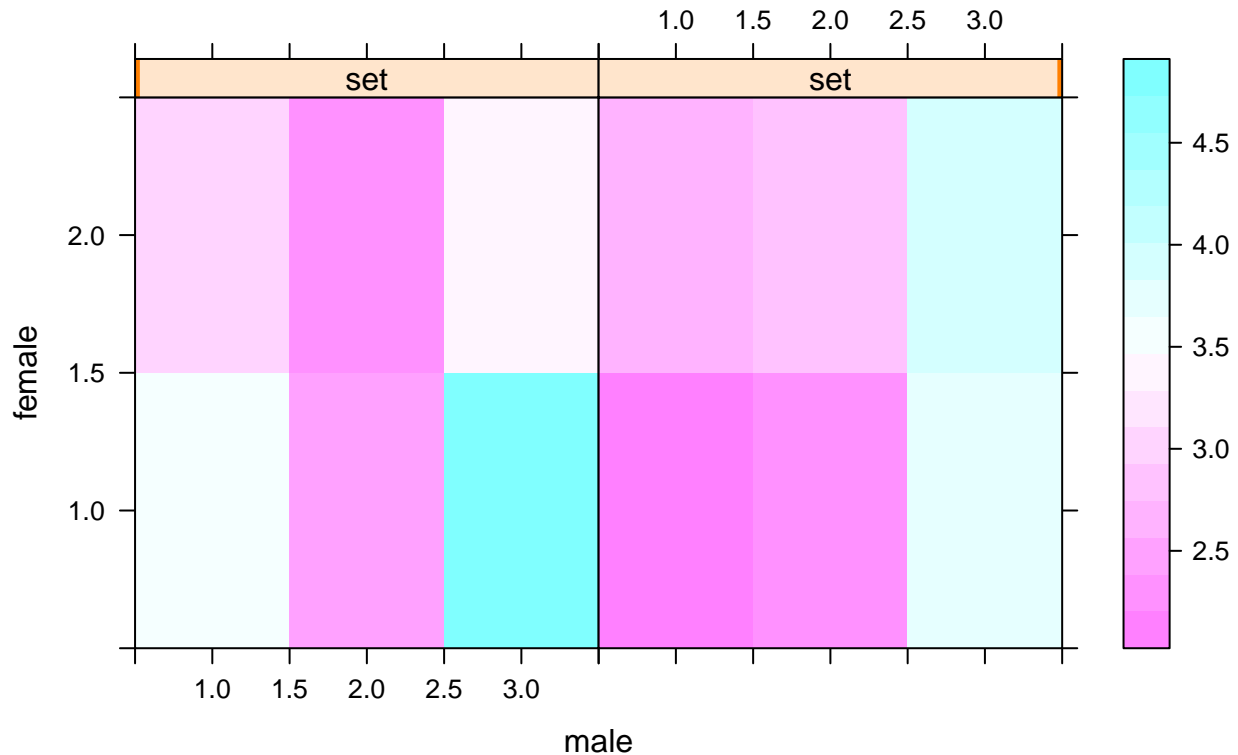
3) Mating designs

Estimating variance components has been a topic of interest for the breeding community for a long time. Here we show how to calculate additive and dominance variance using the North Carolina Design I (Nested design) and North Carolina Design II (Factorial design) using the classical Expected Mean Squares method and the REML methods from sommer and how these two are equivalent.

North Carolina Design I (Nested design)

```
data(DT_expdesigns)
DT <- DT_expdesigns$car1
DT <- aggregate(yield~set+male+female+rep, data=DT, FUN = mean)
DT$setf <- as.factor(DT$set)
DT$repf <- as.factor(DT$rep)
DT$malef <- as.factor(DT$male)
DT$femalef <- as.factor(DT$female)
levelplot(yield~male*female|set, data=DT, main="NC design I")
```

NC design I



```
#####
## Expected Mean Square method
#####
mix1 <- lm(yield~ setf + setf:repf + femalef:malef:setf + malef:setf, data=DT)
MS <- anova(mix1); MS
```

```
## Analysis of Variance Table
##
## Response: yield
##           Df Sum Sq Mean Sq F value  Pr(>F)
## setf      1  0.1780  0.17796   1.6646 0.226012
## setf:repf  2  0.9965  0.49824   4.6605 0.037141 *
## setf:malef  4  7.3904  1.84759  17.2822 0.000173 ***
## setf:femalef:malef  6  1.6083  0.26806   2.5074 0.095575 .
## Residuals 10  1.0691  0.10691
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
ms1 <- MS["malef:setf", "Mean Sq"]
ms2 <- MS["femalef:malef:setf", "Mean Sq"]
mse <- MS["Residuals", "Mean Sq"]
nrep=2
nfem=2
Vfm <- (ms2-mse)/nrep
Vm <- (ms1-ms2)/(nrep*nfem)

## Calculate Va and Vd
Va=4*Vm # assuming no inbreeding (4/(1+F))
```

```
Vd=4*(Vfm-Vm) # assuming no inbreeding (4/(1+F)^2)
Vg=c(Va,Vd); names(Vg) <- c("Va","Vd"); Vg
```

```
## Va Vd
## NA NA
```

```
#####
```

```
## REML method
```

```
#####
```

```
mix2 <- mmer(yield~ setf + setf:repf,
             random=~femalef:malef:setf + malef:setf,
             data=DT, verbose = FALSE)
```

```
## Version out of date. Please update sommer to the newest version using:
## install.packages('sommer') in a new session
## Use the 'date.warning' argument to disable the warning message.
```

```
vc <- summary(mix2)$varcomp; vc
```

```
##                               VarComp  VarCompSE   Zratio Constraint
## femalef:malef:setf.yield-yield 0.08056338 0.08096526 0.9950364   Positive
## malef:setf.yield-yield          0.39480593 0.32832346 1.2024908   Positive
## units.yield-yield               0.10691762 0.04785610 2.2341480   Positive
```

```
Vfm <- vc[1,"VarComp"]
```

```
Vm <- vc[2,"VarComp"]
```

```
## Calculate Va and Vd
```

```
Va=4*Vm # assuming no inbreeding (4/(1+F))
```

```
Vd=4*(Vfm-Vm) # assuming no inbreeding (4/(1+F)^2)
```

```
Vg=c(Va,Vd); names(Vg) <- c("Va","Vd"); Vg
```

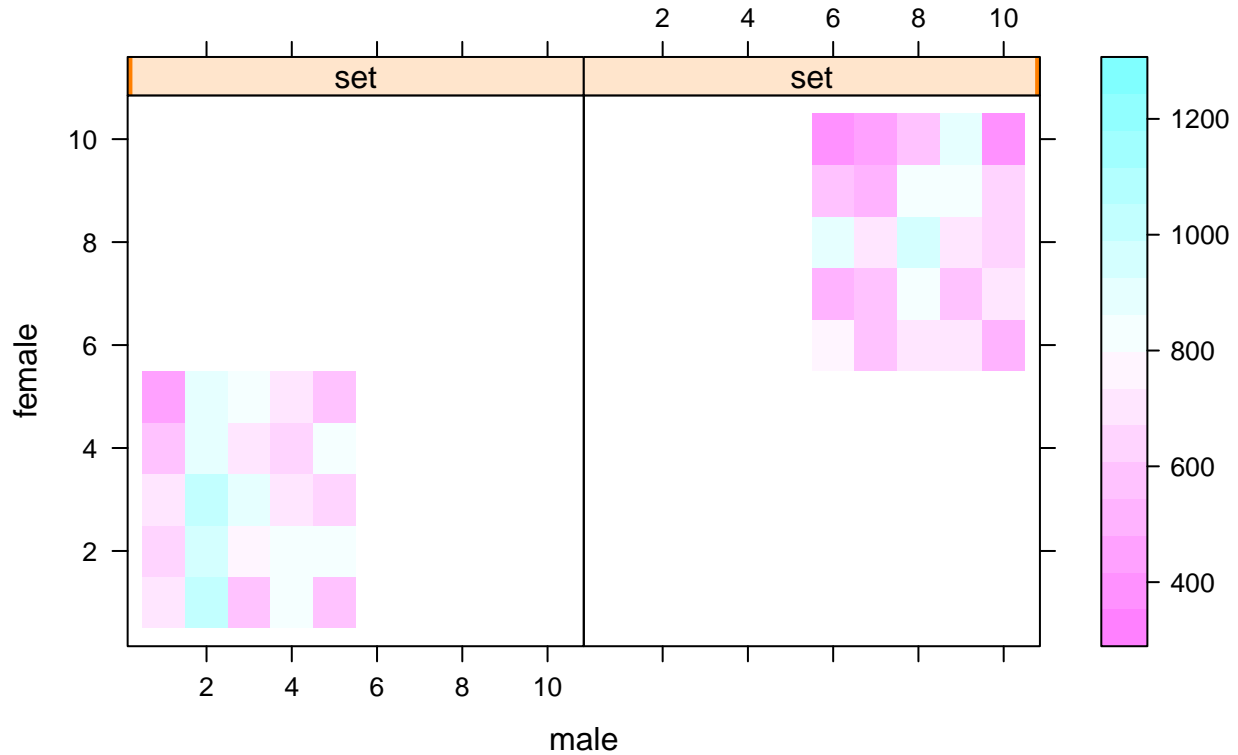
```
##           Va           Vd
## 1.579224 -1.256970
```

As can be seen the REML method is easier than manipulating the MS and we arrive to the same results.

North Carolina Design II (Factorial design)

```
DT <- DT_expdesigns$car2
DT <- aggregate(yield~set+male+female+rep, data=DT, FUN = mean)
DT$setf <- as.factor(DT$set)
DT$repf <- as.factor(DT$rep)
DT$malef <- as.factor(DT$male)
DT$femalef <- as.factor(DT$female)
levelplot(yield~male*female|set, data=DT, main="NC desing II")
```

NC desing II



head(DT)

```
##   set male female rep   yield setf repf malef femalef
## 1   1   1     1     1 831.03   1   1     1     1
## 2   1   2     1     1 1046.55  1   1     2     1
## 3   1   3     1     1  853.33  1   1     3     1
## 4   1   4     1     1  940.00  1   1     4     1
## 5   1   5     1     1  802.00  1   1     5     1
## 6   1   1     2     1  625.93  1   1     1     2
```

```
N=with(DT,table(female, male, set))
```

```
nmale=length(which(N[, ,1] > 0))
```

```
nfemale=length(which(N[, 1,1] > 0))
```

```
nrep=table(N[, ,1])
```

```
nrep=as.numeric(names(nrep[which(names(nrep) !=0)]))
```

```
#####
```

```
## Expected Mean Square method
```

```
#####
```

```
mix1 <- lm(yield~ setf + setf:repf +
```

```
femalef:malef:setf + malef:setf + femalef:setf, data=DT)
```

```
MS <- anova(mix1); MS
```

```
## Analysis of Variance Table
```

```
##
```

```
## Response: yield
```

```
##           Df Sum Sq Mean Sq F value    Pr(>F)
## setf      1  847836  847836  45.6296 1.097e-09 ***
```

```
## setf:repf          4 144345   36086  1.9421  0.109652
## setf:malef         8 861053  107632  5.7926  5.032e-06 ***
## setf:femalef       8 527023   65878  3.5455  0.001227 **
## setf:femalef:malef 32 807267   25227  1.3577  0.129527
## Residuals         96 1783762  18581
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
ms1 <- MS["setf:malef", "Mean Sq"]
ms2 <- MS["setf:femalef", "Mean Sq"]
ms3 <- MS["setf:femalef:malef", "Mean Sq"]
mse <- MS["Residuals", "Mean Sq"]
nrep=length(unique(DT$rep))
nfem=length(unique(DT$female))
nmale=length(unique(DT$male))
Vfm <- (ms3-mse)/nrep;
Vf <- (ms2-ms3)/(nrep*nmale);
Vm <- (ms1-ms3)/(nrep*nfem);

Va=4*Vm; # assuming no inbreeding (4/(1+F))
Va=4*Vf; # assuming no inbreeding (4/(1+F))
Vd=4*(Vfm); # assuming no inbreeding (4/(1+F)^2)
Vg=c(Va,Vd); names(Vg) <- c("Va","Vd"); Vg
```

```
##          Va          Vd
## 10840.192  8861.659
```

```
#####
## REML method
#####
```

```
mix2 <- mmer(yield~ setf + setf:repf ,
             random=-femalef:malef:setf + malef:setf + femalef:setf,
             data=DT)
```

```
## Version out of date. Please update sommer to the newest version using:
```

```
## install.packages('sommer') in a new session
```

```
## Use the 'date.warning' argument to disable the warning message.iteration    LogLik    wall    cpu(
```

```
##      1      -47.2423  21:10:27      0      0
##      2      -46.9375  21:10:27      0      0
##      3      -46.8452  21:10:27      0      0
##      4      -46.8361  21:10:27      0      0
##      5      -46.836  21:10:27      0      0
```

```
vc <- summary(mix2)$varcomp; vc
```

	VarComp	VarCompSE	Zratio	Constraint
## femalef:malef:setf.yield-yield	2215.618	2284.794	0.9697231	Positive
## malef:setf.yield-yield	5493.338	3610.989	1.5212836	Positive
## femalef:setf.yield-yield	2710.176	2236.621	1.2117280	Positive
## units.yield-yield	18580.739	2681.742	6.9286068	Positive

```
Vfm <- vc[1,"VarComp"]
Vm <- vc[2,"VarComp"]
Vf <- vc[3,"VarComp"]
```

```
Va=4*Vm; # assuming no inbreeding (4/(1+F))
```



```
Va=4*Vf; # assuming no inbreeding (4/(1+F))
Vd=4*(Vfm); # assuming no inbreeding(4/(1+F)^2)
Vg=c(Va,Vd); names(Vg) <- c("Va","Vd"); Vg
```

```
##           Va           Vd
## 10840.704  8862.471
```

As can be seen the REML method is easier than manipulating the MS and we arrive to the same results.

4) Dominance variance

The estimation of non-additive variance has been proposed to be a challenge since the additive and dominance relationship matrices are not orthogonal. In recent literature it has been proposed that the best practice to fit the dominance component is to fit the additive component first and then fix the value of that variance c

```
data(DT_cpdata)
DT <- DT_cpdata
GT <- GT_cpdata
MP <- MP_cpdata
#### create the variance-covariance matrix
A <- A.mat(GT) # additive relationship matrix
#### look at the data and fit the model
mix1 <- mmer(Yield~1,
             random=~vs(id,Gu=A),
             rcov=~units,
             data=DT, verbose = FALSE)
```

```
## Version out of date. Please update sommer to the newest version using:
## install.packages('sommer') in a new session
## Use the 'date.warning' argument to disable the warning message.
```

```
####=====####
#### adding dominance and forcing the other VC's
####=====####
DT$idd <- DT$id;
D <- D.mat(GT) # dominance relationship matrix
mm <- matrix(3,1,1) ## matrix to fix the var comp

mix2 <- mmer(Yield~1,
             random=~vs(id, Gu=A, Gt=mix1$sigma_scaled$`u:id`, Gtc=mm)
             + vs(idd, Gu=D, Gtc=unsm(1)),
             rcov=~vs(units,Gt=mix1$sigma_scaled$units, Gtc=mm),
             data=DT, verbose = FALSE)
```

```
## Version out of date. Please update sommer to the newest version using:
## install.packages('sommer') in a new session
## Use the 'date.warning' argument to disable the warning message.
```

```
# analyze variance components
summary(mix1)$varcomp
```

```
##           VarComp VarCompSE   Zratio Constraint
## u:id.Yield-Yield  650.4145  325.5562  1.997856   Positive
## units.Yield-Yield 4031.0153  344.6051 11.697493   Positive
```

```
summary(mix2)$varcomp
```

##	VarComp	VarCompSE	Zratio	Constraint
## u:id.Yield-Yield	650.4145	504.0361	1.2904126	Fixed
## u:idd.Yield-Yield	220.6311	410.7679	0.5371186	Positive
## u:units.Yield-Yield	4031.0153	360.7322	11.1745357	Fixed

Literature

Covarrubias-Pazaran G. 2016. Genome assisted prediction of quantitative traits using the R package sommer. PLoS ONE 11(6):1-15.

Covarrubias-Pazaran G. 2018. Software update: Moving the R package sommer to multivariate mixed models for genome-assisted prediction. doi: <https://doi.org/10.1101/354639>

Bernardo Rex. 2010. Breeding for quantitative traits in plants. Second edition. Stemma Press. 390 pp.

Gilmour et al. 1995. Average Information REML: An efficient algorithm for variance parameter estimation in linear mixed models. Biometrics 51(4):1440-1450.

Henderson C.R. 1975. Best Linear Unbiased Estimation and Prediction under a Selection Model. Biometrics vol. 31(2):423-447.

Kang et al. 2008. Efficient control of population structure in model organism association mapping. Genetics 178:1709-1723.

Lee, D.-J., Durban, M., and Eilers, P.H.C. (2013). Efficient two-dimensional smoothing with P-spline ANOVA mixed models and nested bases. Computational Statistics and Data Analysis, 61, 22 - 37.

Lee et al. 2015. MTG2: An efficient algorithm for multivariate linear mixed model analysis based on genomic information. Cold Spring Harbor. doi: <http://dx.doi.org/10.1101/027201>.

Maier et al. 2015. Joint analysis of psychiatric disorders increases accuracy of risk prediction for schizophrenia, bipolar disorder, and major depressive disorder. Am J Hum Genet; 96(2):283-294.

Rodriguez-Alvarez, Maria Xose, et al. Correcting for spatial heterogeneity in plant breeding experiments with P-splines. Spatial Statistics 23 (2018): 52-71.

Searle. 1993. Applying the EM algorithm to calculating ML and REML estimates of variance components. Paper invited for the 1993 American Statistical Association Meeting, San Francisco.

Yu et al. 2006. A unified mixed-model method for association mapping that accounts for multiple levels of relatedness. Genetics 38:203-208.

Abdollahi Arpanahi R, Morota G, Valente BD, Kranis A, Rosa GJM, Gianola D. 2015. Assessment of bagging GBLUP for whole genome prediction of broiler chicken traits. Journal of Animal Breeding and Genetics 132:218-228.

Tunncliffe W. 1989. On the use of marginal likelihood in time series model estimation. JRSS 51(1):15-27.