Package 'mets'

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Type Package

Title Analysis of Multivariate Event Times

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Description Implementation of various statistical models for multivariate event history data <doi:10.1007/s10985-013-9244-x>. Including multivariate cumulative incidence models <doi:10.1002/sim.6016>, and bivariate random effects probit models (Liability models) <doi:10.1016/j.csda.2015.01.014>. Also contains two-stage binomial modelling that can do pairwise odds-ratio dependence modelling based marginal logistic regression models. This is an alternative to the alternating logistic regression approach (ALR).

License GPL (>= 2)

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BugReports https://github.com/kkholst/mets/issues

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`mets-package` *Analysis of Multivariate Events*

Description

Implementation of various statistical models for multivariate event history data. Including multivariate cumulative incidence models, and bivariate random effects probit models (Liability models)

Author(s)

Klaus K. Holst and Thomas Scheike

Examples

```
## To appear
```

`aalenfrailty` *Aalen frailty model*

Description

Additive hazards model with (gamma) frailty

Usage

```
aalenfrailty(time, status, X, id, theta, B = NULL, ...)
```

Arguments

time	Time variable
status	Status variable (0,1)
X	Covariate design matrix
id	cluster variable
theta	list of thetas (returns score evaluated here), or starting point for optimization (defaults to magic number 0.1)
B	(optional) Cumulative coefficients (update theta by fixing B)
...	Additional arguments to lower level functions

Details

Aalen frailty model

Value

Parameter estimates

Author(s)

Klaus K. Holst

Examples

```
library("timereg")
dd <- simAalenFrailty(5000)
f <- ~1##+x
X <- model.matrix(f,dd) ## design matrix for non-parametric terms
system.time(out<-aalen(update(f,Surv(time,status)~.),dd,n.sim=0,robust=0))
dix <- which(dd$status==1)
t1 <- system.time(bb <- .Call("Bhat",as.integer(dd$status),
                              X,0.2,as.integer(dd$id),NULL,NULL,
                              PACKAGE="mets"))
spec <- 1
##plot(out,spec=spec)
## plot(dd$time[dix],bb$B2[,spec],col="red",type="s",
##      ylim=c(0,max(dd$time)*c(beta0,beta)[spec]))
## abline(a=0,b=c(beta0,beta)[spec])
##'

## Not run:
thetas <- seq(0.1,2,length.out=10)
Us <- unlist(aalenfrailty(dd$time,dd$status,X,dd$id,as.list(thetas)))
##plot(thetas,Us,type="l",ylim=c(-.5,1)); abline(h=0,lty=2); abline(v=theta,lty=2)
op <- aalenfrailty(dd$time,dd$status,X,dd$id)
op

## End(Not run)
```

back2timereg *Convert to timereg object*

Description

convert to timereg object

Usage

```
back2timereg(obj)
```

Arguments

obj	no use

Author(s)

Thomas Scheike

base1cumhaz *rate of CRBSI for HPN patients of Copenhagen*

Description

rate of CRBSI for HPN patients of Copenhagen

Source

Estimated data

base44cumhaz *rate of Occlusion/Thrombosis complication for catheter of HPN patients of Copenhagen*

Description

rate of Occlusion/Thrombosis complication for catheter of HPN patients of Copenhagen

Source

Estimated data

base4cumhaz — *rate of Mechanical (hole/defect) complication for catheter of HPN patients of Copenhagen*

Description

rate of Mechanical (hole/defect) complication for catheter of HPN patients of Copenhagen

Source

Estimated data

basehazplot.phreg — *Plotting the baslines of stratified Cox*

Description

Plotting the baslines of stratified Cox

Usage

```
basehazplot.phreg(x, se = FALSE, time = NULL, add = FALSE,
  ylim = NULL, xlim = NULL, lty = NULL, col = NULL,
  legend = TRUE, ylab = NULL, xlab = NULL, polygon = TRUE,
  level = 0.95, stratas = NULL, robust = FALSE, ...)
```

Arguments

x	phreg object
se	to include standard errors
time	to plot for specific time variables
add	to add to previous plot
ylim	to give ylim
xlim	to give xlim
lty	to specify lty of components
col	to specify col of components
legend	to specify col of components
ylab	to specify ylab
xlab	to specify xlab
polygon	to get standard error in shaded form
level	of standard errors
stratas	wich strata to plot
robust	to use robust standard errors if possible
...	Additional arguments to lower level funtions

Author(s)

Klaus K. Holst, Thomas Scheike

Examples

```
data(TRACE)
dcut(TRACE) <- ~.
out1 <- phreg(Surv(time,status==9)~vf+chf+strata(wmicat.4),data=TRACE)

par(mfrow=c(2,2))
bplot(out1)
bplot(out1,stratas=c(0,3))
bplot(out1,stratas=c(0,3),col=2:3,lty=1:2,se=TRUE)
bplot(out1,stratas=c(0),col=2,lty=2,se=TRUE,polygon=FALSE)
bplot(out1,stratas=c(0),col=matrix(c(2,1,3),1,3),
            lty=matrix(c(1,2,3),1,3),se=TRUE,polygon=FALSE)
```

---

bicomprisk *Estimation of concordance in bivariate competing risks data*

Description

Estimation of concordance in bivariate competing risks data

Usage

```
bicomprisk(formula, data, cause = c(1, 1), cens = 0, causes, indiv,
  strata = NULL, id, num, max.clust = 1000, marg = NULL,
  se.clusters = NULL, wname = NULL, prodlim = FALSE,
  messages = TRUE, model, return.data = 0, uniform = 0,
  conservative = 1, resample.iid = 1, ...)
```

Arguments

formula	Formula with left-hand-side being a Event object (see example below) and the left-hand-side specying the covariate structure
data	Data frame
cause	Causes (default (1,1)) for which to estimate the bivariate cumulative incidence
cens	The censoring code
causes	causes
indiv	indiv
strata	Strata
id	Clustering variable
num	num

max.clust	max number of clusters in comp.risk call for iid decompostion, max.clust=NULL uses all clusters otherwise rougher grouping.
marg	marginal cumulative incidence to make stanard errors for same clusters for subsequent use in casewise.test()
se.clusters	to specify clusters for standard errors. Either a vector of cluster indices or a column name in data. Defaults to the id variable.
wname	name of additonal weight used for paired competing risks data.
prodlim	prodlim to use prodlim estimator (Aalen-Johansen) rather than IPCW weighted estimator based on comp.risk function.These are equivalent in the case of no covariates. These esimators are the same in the case of stratified fitting.
messages	Control amount of output
model	Type of competing risk model (default is Fine-Gray model "fg", see comp.risk).
return.data	Should data be returned (skipping modeling)
uniform	to compute uniform standard errors for concordance estimates based on resampling.
conservative	for conservative standard errors, recommended for larger data-sets.
resample.iid	to return iid residual processes for further computations such as tests.
...	Additional arguments to comp.risk function

Author(s)

Thomas Scheike, Klaus K. Holst

References

Scheike, T. H.; Holst, K. K. & Hjelmborg, J. B. Estimating twin concordance for bivariate competing risks twin data Statistics in Medicine, Wiley Online Library, 2014 , 33 , 1193-204

Examples

```
library("timereg")

## Simulated data example
prt <- simnordic.random(2000,delayed=TRUE,ptrunc=0.7,
      cordz=0.5,cormz=2,lam0=0.3)
## Bivariate competing risk, concordance estimates
p11 <- bicomprisk(Event(time,cause)~strata(zyg)+id(id),data=prt,cause=c(1,1))

p11mz <- p11$model$"MZ"
p11dz <- p11$model$"DZ"
par(mfrow=c(1,2))
## Concordance
plot(p11mz,ylim=c(0,0.1));
plot(p11dz,ylim=c(0,0.1));

## entry time, truncation weighting
### other weighting procedure
```

```
prtl <-  prt[!prt$truncated,]
prt2 <- ipw2(prtl,cluster="id",same.cens=TRUE,
     time="time",cause="cause",entrytime="entry",
     pairs=TRUE,strata="zyg",obs.only=TRUE)

prt22 <- fast.reshape(prt2,id="id")

prt22$event <- (prt22$cause1==1)*(prt22$cause2==1)*1
prt22$timel <- pmax(prt22$time1,prt22$time2)
ipwc <- comp.risk(Event(timel,event)~-1+factor(zyg1),
  data=prt22,cause=1,n.sim=0,model="rcif2",times=50:90,
  weights=prt22$weights1,cens.weights=rep(1,nrow(prt22)))

p11wmz <- ipwc$cum[,2]
p11wdz <- ipwc$cum[,3]
lines(ipwc$cum[,1],p11wmz,col=3)
lines(ipwc$cum[,1],p11wdz,col=3)
```

`binomial.twostage` *Fits Clayton-Oakes or bivariate Plackett (OR) models for binary data using marginals that are on logistic form. If clusters contain more than two times, the algoritm uses a compososite likelihood based on all pairwise bivariate models.*

Description

The pairwise pairwise odds ratio model provides an alternative to the alternating logistic regression (ALR).

Usage

```
binomial.twostage(margbin, data = sys.parent(),
  score.method = "fisher.scoring", Nit = 60, detail = 0,
  clusters = NULL, silent = 1, weights = NULL, control = list(),
  theta = NULL, theta.des = NULL, var.link = 0, var.par = 1,
  var.func = NULL, iid = 1, step = 1, notaylor = 1,
  model = "plackett", marginal.p = NULL, beta.iid = NULL,
  Dbeta.iid = NULL, strata = NULL, max.clust = NULL,
  se.clusters = NULL, numDeriv = 0, random.design = NULL,
  pairs = NULL, pairs.rvs = NULL, additive.gamma.sum = NULL,
  pair.ascertained = 0, case.control = 0, twostage = 1,
  beta = NULL)
```

Arguments

`margbin`	Marginal binomial model
`data`	data frame

score.method	Scoring method default is "fisher.scoring" among "fisher.scoring","nlminb","optimize","nlm"
Nit	Number of iterations
detail	Detail
clusters	Cluster variable
silent	Debug information
weights	Weights for log-likelihood, can be used for each type of outcome in 2x2 tables.
control	Optimization arguments
theta	Starting values for variance components
theta.des	design for dependence parameters, when pairs are given this is could be a (pairs) x (numer of parameters) x (max number random effects) matrix
var.link	Link function for variance
var.par	parametrization
var.func	when alternative parametrizations are used this function can specify how the paramters are related to the λ_j's.
iid	Calculate i.i.d. decomposition when iid>=1, when iid=2 then avoids adding the uncertainty for marginal paramters for additive gamma model (default).
step	Step size
notaylor	Taylor expansion
model	model
marginal.p	vector of marginal probabilities
beta.iid	iid decomposition of marginal probability estimates for each subject, if based on GLM model this is computed.
Dbeta.iid	derivatives of marginal model wrt marginal parameters, if based on GLM model this is computed.
strata	strata for fitting: considers only pairs where both are from same strata
max.clust	max clusters
se.clusters	clusters for iid decomposition for roubst standard errors
numDeriv	uses Fisher scoring aprox of second derivative if 0, otherwise numerical derivatives
random.design	random effect design for additive gamma model, when pairs are given this is a (pairs) x (2) x (max number random effects) matrix, see pairs.rvs below
pairs	matrix with rows of indeces (two-columns) for the pairs considered in the pairwise composite score, useful for case-control sampling when marginal is known.
pairs.rvs	for additive gamma model and random.design and theta.des are given as arrays, this specifice number of random effects for each pair.
additive.gamma.sum	this is specification of the lamtot in the models via a matrix that is multiplied onto the parameters theta (dimensions=(number random effects x number of theta parameters), when null then sums all parameters. Default is a matrix of 1's
pair.ascertained	if pairs are sampled only when there are events in the pair i.e. Y1+Y2>=1.

case.control	if data is case control data for pair call, and here 2nd column of pairs are probands (cases or controls)
twostage	default twostage=1, to fit MLE use twostage=0
beta	is starting value for beta for MLE version

Details

The reported standard errors are based on a cluster corrected score equations from the pairwise likelihoods assuming that the marginals are known. This gives correct standard errors in the case of the Odds-Ratio model (Plackett distribution) for dependence, but incorrect standard errors for the Clayton-Oakes types model (that is also called "gamma"-frailty). For the additive gamma version of the standard errors are adjusted for the uncertainty in the marginal models via an iid deomposition using the iid() function of lava. For the clayton oakes model that is not speicifed via the random effects these can be fixed subsequently using the iid influence functions for the marginal model, but typically this does not change much.

For the Clayton-Oakes version of the model, given the gamma distributed random effects it is assumed that the probabilities are indpendent, and that the marginal survival functions are on logistic form

$$logit(P(Y=1|X)) = \alpha + x^T\beta$$

therefore conditional on the random effect the probability of the event is

$$logit(P(Y=1|X,Z)) = exp(-Z \cdot Laplace^{-1}(lamtot, lamtot, P(Y=1|x)))$$

Can also fit a structured additive gamma random effects model, such the ACE, ADE model for survival data:

Now random.design specificies the random effects for each subject within a cluster. This is a matrix of 1's and 0's with dimension n x d. With d random effects. For a cluster with two subjects, we let the random.design rows be v_1 and v_2. Such that the random effects for subject 1 is

$$v_1^T(Z_1, ..., Z_d)$$

, for d random effects. Each random effect has an associated parameter $(\lambda_1, ..., \lambda_d)$. By construction subjects 1's random effect are Gamma distributed with mean $\lambda_j/v_1^T\lambda$ and variance $\lambda_j/(v_1^T\lambda)^2$. Note that the random effect $v_1^T(Z_1, ..., Z_d)$ has mean 1 and variance $1/(v_1^T\lambda)$. It is here asssumed that $lamtot = v_1^T\lambda$ is fixed over all clusters as it would be for the ACE model below.

The DEFAULT parametrization uses the variances of the random effecs (var.par=1)

$$\theta_j = \lambda_j/(v_1^T\lambda)^2$$

For alternative parametrizations (var.par=0) one can specify how the parameters relate to λ_j with the function

Based on these parameters the relative contribution (the heritability, h) is equivalent to the expected values of the random effects $\lambda_j/v_1^T\lambda$

Given the random effects the probabilities are independent and on the form

$$logit(P(Y=1|X)) = exp(-Laplace^{-1}(lamtot, lamtot, P(Y=1|x)))$$

with the inverse laplace of the gamma distribution with mean 1 and variance lamtot.

The parameters $(\lambda_1, ..., \lambda_d)$ are related to the parameters of the model by a regression construction $pard$ (d x k), that links the d λ parameters with the (k) underlying θ parameters

$$\lambda = theta.des\theta$$

here using theta.des to specify these low-dimension association. Default is a diagonal matrix.

Author(s)

Thomas Scheike

References

Two-stage binomial modelling

Examples

```
library("timereg")
data("twinstut",package="mets")
twinstut0 <- subset(twinstut, tvparnr<2300000)
twinstut <- twinstut0
twinstut$binstut <- (twinstut$stutter=="yes")*1
theta.des <- model.matrix( ~-1+factor(zyg),data=twinstut)
margbin <- glm(binstut~factor(sex)+age,data=twinstut,family=binomial())
bin <- binomial.twostage(margbin,data=twinstut,var.link=1,
         clusters=twinstut$tvparnr,theta.des=theta.des,detail=0,
                 score.method="fisher.scoring")
summary(bin)

twinstut$cage <- scale(twinstut$age)
theta.des <- model.matrix( ~-1+factor(zyg)+cage,data=twinstut)
bina <- binomial.twostage(margbin,data=twinstut,var.link=1,
         clusters=twinstut$tvparnr,theta.des=theta.des)
summary(bina)

theta.des <- model.matrix( ~-1+factor(zyg)+factor(zyg)*cage,data=twinstut)
bina <- binomial.twostage(margbin,data=twinstut,var.link=1,
         clusters=twinstut$tvparnr,theta.des=theta.des)
summary(bina)

## refers to zygosity of first subject in eash pair : zyg1
## could also use zyg2 (since zyg2=zyg1 within twinpair's))
out <- easy.binomial.twostage(stutter~factor(sex)+age,data=twinstut,
                          response="binstut",id="tvparnr",var.link=1,
                     theta.formula=~-1+factor(zyg1))
summary(out)

## refers to zygosity of first subject in eash pair : zyg1
## could also use zyg2 (since zyg2=zyg1 within twinpair's))
desfs<-function(x,num1="zyg1",num2="zyg2")
    c(x[num1]=="dz",x[num1]=="mz",x[num1]=="os")*1

out3 <- easy.binomial.twostage(binstut~factor(sex)+age,
```

```
      data=twinstut,response="binstut",id="tvparnr",var.link=1,
      theta.formula=desfs,desnames=c("mz","dz","os"))
summary(out3)

### use of clayton oakes binomial additive gamma model
############################################################
 ## Reduce Ex.Timings
data <- simbinClaytonOakes.family.ace(10000,2,1,beta=NULL,alpha=NULL)
margbin <- glm(ybin~x,data=data,family=binomial())
margbin

head(data)
data$number <- c(1,2,3,4)
data$child <- 1*(data$number==3)

### make ace random effects design
out <- ace.family.design(data,member="type",id="cluster")
out$pardes
head(out$des.rv)

bints <- binomial.twostage(margbin,data=data,
     clusters=data$cluster,detail=0,var.par=1,
     theta=c(2,1),var.link=0,
     random.design=out$des.rv,theta.des=out$pardes)
summary(bints)

data <- simbinClaytonOakes.twin.ace(10000,2,1,beta=NULL,alpha=NULL)
out  <- twin.polygen.design(data,id="cluster",zygname="zygosity")
out$pardes
head(out$des.rv)
margbin <- glm(ybin~x,data=data,family=binomial())

bintwin <- binomial.twostage(margbin,data=data,
     clusters=data$cluster,detail=1,var.par=1,
     theta=c(2,1),random.design=out$des.rv,theta.des=out$pardes)
summary(bintwin)
concordanceTwinACE(bintwin)
```

---

binreg *Binomial Regression for censored competing risks data*

Description

Simple version of comp.risk function of timereg for just one time-point thus fitting the model

$$P(T \leq t, \epsilon = 1|X) = expit(X^T beta)$$

Usage

```
binreg(formula, data, cause = 1, time = NULL, beta = NULL,
  offset = NULL, weights = NULL, cens.weights = NULL,
  cens.model = ~+1, se = TRUE, kaplan.meier = TRUE, cens.code = 0,
  no.opt = FALSE, method = "nr", ...)
```

Arguments

formula	formula with outcome (see coxph)
data	data frame
cause	cause of interest
time	time of interest
beta	starting values
offset	offsets for partial likelihood
weights	for score equations
cens.weights	censoring weights
cens.model	stratified cox model
se	to compute se's based on IPCW
kaplan.meier	uses Kaplan-Meier for baseline than standard Cox
cens.code	gives censoring code
no.opt	to not optimize
method	for optimization
...	Additional arguments to lower level funtions

Details

Based on binomial regresion IPCW response estimating equation:

$$X(\Delta I(T \leq t, \epsilon = 1)/G_c(T_i-) - expit(X^T beta)) = 0$$

for IPCW adjusted responses.

Author(s)

Thomas Scheike

Examples

```
data(bmt)
# logistic regresion with IPCW binomial regression
out <- binreg(Event(time,cause)~tcell+platelet,bmt,time=50)
summary(out)
predict(out,data.frame(tcell=c(0,1),platelet=c(1,1)),se=TRUE)

outs <- binreg(Event(time,cause)~tcell+platelet,bmt,time=50,cens.model=~strata(tcell,platelet))
summary(outs)
```

biprobit *Bivariate Probit model*

Description

Bivariate Probit model

Usage

```
biprobit(x, data, id, rho = ~1, num = NULL, strata = NULL,
  eqmarg = TRUE, indep = FALSE, weights = NULL, biweight,
  samecens = TRUE, randomeffect = FALSE, vcov = "robust",
  pairs.only = FALSE, allmarg = samecens & !is.null(weights),
  control = list(trace = 0), messages = 1, constrain = NULL,
  table = pairs.only, p = NULL, ...)
```

Arguments

x	formula (or vector)
data	data.frame
id	The name of the column in the dataset containing the cluster id-variable.
rho	Formula specifying the regression model for the dependence parameter
num	Optional name of order variable
strata	Strata
eqmarg	If TRUE same marginals are assumed (exchangeable)
indep	Independence
weights	Weights
biweight	Function defining the bivariate weight in each cluster
samecens	Same censoring
randomeffect	If TRUE a random effect model is used (otherwise correlation parameter is estimated allowing for both negative and positive dependence)
vcov	Type of standard errors to be calculated
pairs.only	Include complete pairs only?
allmarg	Should all marginal terms be included
control	Control argument parsed on to the optimization routine. Starting values may be parsed as 'start'.
messages	Control amount of messages shown
constrain	Vector of parameter constraints (NA where free). Use this to set an offset.
table	Type of estimation procedure
p	Parameter vector p in which to evaluate log-Likelihood and score function
...	Optional arguments

Examples

```
data(prt)
prt0 <- subset(prt,country=="Denmark")
a <- biprobit(cancer~1+zyg, ~1+zyg, data=prt0, id="id")
b <- biprobit(cancer~1+zyg, ~1+zyg, data=prt0, id="id",pairs.only=TRUE)
predict(b,newdata=lava::Expand(prt,zyg=c("MZ")))
predict(b,newdata=lava::Expand(prt,zyg=c("MZ","DZ")))

 ## Reduce Ex.Timings
library(lava)
m <- lvm(c(y1,y2)~x)
covariance(m,y1~y2) <- "r"
constrain(m,r~x+a+b) <- function(x) tanh(x[2]+x[3]*x[1])
distribution(m,~x) <- uniform.lvm(a=-1,b=1)
ordinal(m) <- ~y1+y2
d <- sim(m,1000,p=c(a=0,b=-1)); d <- d[order(d$x),]
dd <- fast.reshape(d)

a <- biprobit(y~1+x,rho=~1+x,data=dd,id="id")
summary(a, mean.contrast=c(1,.5), cor.contrast=c(1,.5))
with(predict(a,data.frame(x=seq(-1,1,by=.1))), plot(p00~x,type="l"))

pp <- predict(a,data.frame(x=seq(-1,1,by=.1)),which=c(1))
plot(pp[,1]~pp$x, type="l", xlab="x", ylab="Concordance", lwd=2, xaxs="i")
confband(pp$x,pp[,2],pp[,3],polygon=TRUE,lty=0,col=Col(1))

pp <- predict(a,data.frame(x=seq(-1,1,by=.1)),which=c(9)) ## rho
plot(pp[,1]~pp$x, type="l", xlab="x", ylab="Correlation", lwd=2, xaxs="i")
confband(pp$x,pp[,2],pp[,3],polygon=TRUE,lty=0,col=Col(1))
with(pp, lines(x,tanh(-x),lwd=2,lty=2))

xp <- seq(-1,1,length.out=6); delta <- mean(diff(xp))
a2 <- biprobit(y~1+x,rho=~1+I(cut(x,breaks=xp)),data=dd,id="id")
pp2 <- predict(a2,data.frame(x=xp[-1]-delta/2),which=c(9)) ## rho
confband(pp2$x,pp2[,2],pp2[,3],center=pp2[,1])




## Time
## Not run:
    a <- biprobit.time(cancer~1, rho=~1+zyg, id="id", data=prt, eqmarg=TRUE,
                       cens.formula=Surv(time,status==0)~1,
                       breaks=seq(75,100,by=3),fix.censweights=TRUE)

    a <- biprobit.time2(cancer~1+zyg, rho=~1+zyg, id="id", data=prt0, eqmarg=TRUE,
                       cens.formula=Surv(time,status==0)~zyg,
                       breaks=100)

   a1 <- biprobit.time2(cancer~1, rho=~1, id="id", data=subset(prt0,zyg=="MZ"), eqmarg=TRUE,
                       cens.formula=Surv(time,status==0)~1,
                       breaks=100,pairs.only=TRUE)
```

```
a2 <- biprobit.time2(cancer~1, rho=~1, id="id", data=subset(prt0,zyg=="DZ"), eqmarg=TRUE,
                     cens.formula=Surv(time,status==0)~1,
                     breaks=100,pairs.only=TRUE)

 prt0$trunc <- prt0$time*runif(nrow(prt0))*rbinom(nrow(prt0),1,0.5)
a3 <- biprobit.time(cancer~1, rho=~1, id="id", data=subset(prt0,zyg=="DZ"), eqmarg=TRUE,
                     cens.formula=Surv(trunc,time,status==0)~1,
                     breaks=100,pairs.only=TRUE)

 plot(a,which=3,ylim=c(0,0.1))

## End(Not run)
```

`blocksample` *Block sampling*

Description

Sample blockwise from clustered data

Usage

```
blocksample(data, size, idvar = NULL, replace = TRUE, ...)
```

Arguments

`data`	Data frame
`size`	Size of samples
`idvar`	Column defining the clusters
`replace`	Logical indicating wether to sample with replacement
`...`	additional arguments to lower level functions

Details

Original id is stored in the attribute 'id'

Value

`data.frame`

Author(s)

Klaus K. Holst

Examples

```
d <- data.frame(x=rnorm(5), z=rnorm(5), id=c(4,10,10,5,5), v=rnorm(5))
(dd <- blocksample(d,size=20,~id))
attributes(dd)$id

## Not run:
blocksample(data.table::data.table(d),1e6,~id)

## End(Not run)


d <- data.frame(x=c(1,rnorm(9)),
                z=rnorm(10),
                id=c(4,10,10,5,5,4,4,5,10,5),
                id2=c(1,1,2,1,2,1,1,1,1,2),
                v=rnorm(10))
dsample(d,~id, size=2)
dsample(d,.~id+id2)
dsample(d,x+z~id|x>0,size=5)
```

Bootphreg *Wild bootstrap for Cox PH regression*

Description

wild bootstrap for uniform bands for Cox models

Usage

```
Bootphreg(formula, data, offset = NULL, weights = NULL, B = 1000,
  type = c("exp", "poisson", "normal"), ...)
```

Arguments

formula	formula with 'Surv' outcome (see coxph)
data	data frame
offset	offsets for cox model
weights	weights for Cox score equations
B	bootstraps
type	distribution for multiplier
...	Additional arguments to lower level funtions

Author(s)

Klaus K. Holst, Thomas Scheike

References

Wild bootstrap based confidence intervals for multiplicative hazards models, Dobler, Pauly, and Scheike (2018),

Examples

```
n <- 100
x <- 4*rnorm(n)
time1 <- 2*rexp(n)/exp(x*0.3)
time2 <- 2*rexp(n)/exp(x*(-0.3))
status <- ifelse(time1<time2,1,2)
time <- pmin(time1,time2)
rbin <- rbinom(n,1,0.5)
cc <-rexp(n)*(rbin==1)+(rbin==0)*rep(3,n)
status <- ifelse(time < cc,status,0)
time  <- ifelse(time < cc,time,cc)
data <- data.frame(time=time,status=status,x=x)

b1 <- Bootphreg(Surv(time,status==1)~x,data,B=1000)
b2 <- Bootphreg(Surv(time,status==2)~x,data,B=1000)
c1 <- phreg(Surv(time,status==1)~x,data)
c2 <- phreg(Surv(time,status==2)~x,data)

### exp to make all bootstraps positive
out <- pred.cif.boot(b1,b2,c1,c2,gplot=0)

cif.true <- (1-exp(-out$time))*.5
with(out,plot(time,cif,ylim=c(0,1),type="l"))
lines(out$time,cif.true,col=3)
with(out,plotConfRegion(time,band.EE,col=1))
with(out,plotConfRegion(time,band.EE.log,col=3))
with(out,plotConfRegion(time,band.EE.log.o,col=2))
```

---

bptwin *Liability model for twin data*

Description

Liability-threshold model for twin data

Usage

```
bptwin(x, data, id, zyg, DZ, group = NULL, num = NULL,
  weights = NULL, biweight = function(x) 1/min(x), strata = NULL,
  messages = 1, control = list(trace = 0), type = "ace",
  eqmean = TRUE, pairs.only = FALSE, samecens = TRUE,
  allmarg = samecens & !is.null(weights), stderr = TRUE,
```

```
robustvar = TRUE, p, indiv = FALSE, constrain, bound = FALSE,
varlink, ...)
```

Arguments

x	Formula specifying effects of covariates on the response.
data	data.frame with one observation pr row. In addition a column with the zygosity (DZ or MZ given as a factor) of each individual much be specified as well as a twin id variable giving a unique pair of numbers/factors to each twin pair.
id	The name of the column in the dataset containing the twin-id variable.
zyg	The name of the column in the dataset containing the zygosity variable.
DZ	Character defining the level in the zyg variable corresponding to the dyzogitic twins.
group	Optional. Variable name defining group for interaction analysis (e.g., gender)
num	Optional twin number variable
weights	Weight matrix if needed by the chosen estimator (IPCW)
biweight	Function defining the bivariate weight in each cluster
strata	Strata
messages	Control amount of messages shown
control	Control argument parsed on to the optimization routine. Starting values may be parsed as 'start'.
type	Character defining the type of analysis to be performed. Should be a subset of "acde" (additive genetic factors, common environmental factors, dominant genetic factors, unique environmental factors).
eqmean	Equal means (with type="cor")?
pairs.only	Include complete pairs only?
samecens	Same censoring
allmarg	Should all marginal terms be included
stderr	Should standard errors be calculated?
robustvar	If TRUE robust (sandwich) variance estimates of the variance are used
p	Parameter vector p in which to evaluate log-Likelihood and score function
indiv	If TRUE the score and log-Likelihood contribution of each twin-pair
constrain	Development argument
bound	Development argument
varlink	Link function for variance parameters
...	Additional arguments to lower level functions

Author(s)

Klaus K. Holst

See Also

twinlm, twinlm.time, twinlm.strata, twinsim

Examples

```
data(twinstut)
b0 <- bptwin(stutter~sex,
             data=droplevels(subset(twinstut,zyg%in%c("mz","dz"))),
             id="tvparnr",zyg="zyg",DZ="dz",type="ae")
summary(b0)
```

`casewise` *Estimates the casewise concordance based on Concordance and marginal estimate using prodlim but no testing*

Description

.. content for description (no empty lines) ..

Usage

```
casewise(conc, marg, cause.marg)
```

Arguments

`conc`	Concordance
`marg`	Marginal estimate
`cause.marg`	specifiefes which cause that should be used for marginal cif based on prodlim

Author(s)

Thomas Scheike

Examples

```
 ## Reduce Ex.Timings
library(prodlim)
data(prt);

### marginal cumulative incidence of prostate cancer##'
outm <- prodlim(Hist(time,status)~+1,data=prt)

times <- 60:100
cifmz <- predict(outm,cause=2,time=times,newdata=data.frame(zyg="MZ")) ## cause is 2 (second cause)
cifdz <- predict(outm,cause=2,time=times,newdata=data.frame(zyg="DZ"))

### concordance for MZ and DZ twins
cc <- bicomprisk(Event(time,status)~strata(zyg)+id(id),data=prt,cause=c(2,2),prodlim=TRUE)
```

```
cdz <- cc$model$"DZ"
cmz <- cc$model$"MZ"

cdz <- casewise(cdz,outm,cause.marg=2)
cmz <- casewise(cmz,outm,cause.marg=2)

plot(cmz,ci=NULL,ylim=c(0,0.5),xlim=c(60,100),legend=TRUE,col=c(3,2,1))
par(new=TRUE)
plot(cdz,ci=NULL,ylim=c(0,0.5),xlim=c(60,100),legend=TRUE)
summary(cdz)
summary(cmz)
```

`casewise.test` *Estimates the casewise concordance based on Concordance and marginal estimate using timereg and performs test for independence*

Description

Estimates the casewise concordance based on Concordance and marginal estimate using timereg and performs test for independence

Usage

```
casewise.test(conc, marg, test = "no-test", p = 0.01)
```

Arguments

`conc`	Concordance
`marg`	Marginal estimate
`test`	Type of test for independence assumption. "conc" makes test on concordance scale and "case" means a test on the casewise concordance
`p`	check that marginal probability is greater at some point than p

Details

Uses cluster based conservative standard errors for marginal

Author(s)

Thomas Scheike

Examples

```
 ## Reduce Ex.Timings
library("timereg")
data("prt",package="mets");

prt <- prt[which(prt$id %in% sample(unique(prt$id),7500)),]
### marginal cumulative incidence of prostate cancer
times <- seq(60,100,by=2)
outm <- comp.risk(Event(time,status)~+1,data=prt,cause=2,times=times)

cifmz <- predict(outm,X=1,uniform=0,resample.iid=1)
cifdz <- predict(outm,X=1,uniform=0,resample.iid=1)

### concordance for MZ and DZ twins
cc <- bicomprisk(Event(time,status)~strata(zyg)+id(id),
                 data=prt,cause=c(2,2))
cdz <- cc$model$"DZ"
cmz <- cc$model$"MZ"

### To compute casewise cluster argument must be passed on,
###  here with a max of 100 to limit comp-time
outm <-comp.risk(Event(time,status)~+1,data=prt,
                 cause=2,times=times,max.clust=100)
cifmz <-predict(outm,X=1,uniform=0,resample.iid=1)
cc <-bicomprisk(Event(time,status)~strata(zyg)+id(id),data=prt,
                cause=c(2,2),se.clusters=outm$clusters)
cdz <- cc$model$"DZ"
cmz <- cc$model$"MZ"

cdz <- casewise.test(cdz,cifmz,test="case") ## test based on casewise
cmz <- casewise.test(cmz,cifmz,test="conc") ## based on concordance

plot(cmz,ylim=c(0,0.7),xlim=c(60,100))
par(new=TRUE)
plot(cdz,ylim=c(0,0.7),xlim=c(60,100))

slope.process(cdz$casewise[,1],cdz$casewise[,2],iid=cdz$casewise.iid)

slope.process(cmz$casewise[,1],cmz$casewise[,2],iid=cmz$casewise.iid)
```

`cif` *Cumulative incidence with robust standard errors*

Description

Cumulative incidence with robust standard errors

Usage

```
cif(formula, data = data, cause = 1, cens.code = 0, ...)
```

Arguments

formula	formula with 'Surv' outcome (see coxph)
data	data frame
cause	NULL looks at all, otherwise specify which cause to consider
cens.code	censoring code "0" is default
...	Additional arguments to lower level funtions

Author(s)

Thomas Scheike

Examples

```
data(TRACE)
TRACE$cluster <- sample(1:100,1878,replace=TRUE)
out1 <- cif(Event(time,status)~+1,data=TRACE,cause=9)
out2 <- cif(Event(time,status)~+1+cluster(cluster),data=TRACE,cause=9)

out1 <- cif(Event(time,status)~strata(vf,chf),data=TRACE,cause=9)
out2 <- cif(Event(time,status)~strata(vf,chf)+cluster(cluster),data=TRACE,cause=9)

par(mfrow=c(1,2))
bplot(out1,se=TRUE)
bplot(out2,se=TRUE)
```

---

cifreg *CIF regression*

Description

CIF logistic for propodds=1 default CIF Fine-Gray (cloglog) regression for propodds=NULL

Usage

```
cifreg(formula, data = data, cause = 1, cens.code = 0,
  weights = NULL, offset = NULL, Gc = NULL, propodds = 1, ...)
```

Arguments

formula	formula with 'Event' outcome
data	data frame
cause	of interest
cens.code	code of censoring
weights	weights for Cox score equations
offset	offsets for cox model
Gc	censoring weights for time argument, default is to calculate these with a Kaplan-Meier estimator, should then give G_c(T_i-)
propodds	1 is logistic model, NULL is fine-gray model
...	Additional arguments to lower level funtions

Details

For FG model:

$$\int (X - E) Y_1(t) w(t) dM_1$$

is computed and summed over clusters and returned multiplied with inverse of second derivative as iid.naive

The iid decomposition of the beta's, however, also have a censoring term that is also is computed and added to UUiid (still scaled with inverse second derivative)

$$\int (X - E) Y_1(t) w(t) dM_1 + \int q(s)/p(s) dM_c$$

and returned as iid

Author(s)

Thomas Scheike

Examples

```
## data with no ties
data(bmt,package="timereg")
bmt$time <- bmt$time+runif(nrow(bmt))*0.01
bmt$id <- 1:nrow(bmt)

## logistic link  OR interpretation
ll=cifreg(Event(time,cause)~tcell+platelet+age,data=bmt,cause=1)
bplot(ll)
nd <- data.frame(tcell=c(1,0),platelet=0,age=0)
pll <- predict(ll,nd)
plot(pll)

## Fine-Gray model
llfg=cifreg(Event(time,cause)~tcell+platelet+age,data=bmt,cause=1,propodds=NULL)
bplot(ll)
```

```
nd <- data.frame(tcell=c(1,0),platelet=0,age=0)
pll <- predict(ll,nd)
plot(pll)
```

ClaytonOakes *Clayton-Oakes model with piece-wise constant hazards*

Description

Clayton-Oakes frailty model

Usage

```
ClaytonOakes(formula, data = parent.frame(), cluster, var.formula = ~1,
  cuts = NULL, type = "piecewise", start, control = list(),
  var.invlink = exp, ...)
```

Arguments

formula	formula specifying the marginal proportional (piecewise constant) hazard structure with the right-hand-side being a survival object (Surv) specifying the entry time (optional), the follow-up time, and event/censoring status at follow-up. The clustering can be specified using the special function `cluster` (see example below).
data	Data frame
cluster	Variable defining the clustering (if not given in the formula)
var.formula	Formula specifying the variance component structure (if not given via the cluster special function in the formula) using a linear model with log-link.
cuts	Cut points defining the piecewise constant hazard
type	when equal to `two.stage`, the Clayton-Oakes-Glidden estimator will be calculated via the `timereg` package
start	Optional starting values
control	Control parameters to the optimization routine
var.invlink	Inverse link function for variance structure model
...	Additional arguments

Author(s)

Klaus K. Holst

Examples

```
set.seed(1)
d <- subset(simClaytonOakes(500,4,2,1,stoptime=2,left=2),truncated)
e <- ClaytonOakes(survival::Surv(lefttime,time,status)~x+cluster(~1,cluster),
                  cuts=c(0,0.5,1,2),data=d)
e


d2 <- simClaytonOakes(500,4,2,1,stoptime=2,left=0)
d2$z <- rep(1,nrow(d2)); d2$z[d2$cluster%in%sample(d2$cluster,100)] <- 0
## Marginal=Cox Proportional Hazards model:
ts <- ClaytonOakes(survival::Surv(time,status)~timereg::prop(x)+cluster(~1,cluster),
                   data=d2,type="two.stage")
## Marginal=Aalens additive model:
ts2 <- ClaytonOakes(survival::Surv(time,status)~x+cluster(~1,cluster),
                    data=d2,type="two.stage")
## Marginal=Piecewise constant:
e2 <- ClaytonOakes(survival::Surv(time,status)~x+cluster(~-1+factor(z),cluster),
                   cuts=c(0,0.5,1,2),data=d2)
e2
plot(ts)
plot(e2,add=TRUE)

e3 <- ClaytonOakes(survival::Surv(time,status)~x+cluster(~1,cluster),cuts=c(0,0.5,1,2),
                   data=d,var.invlink=identity)
e3
```

`cluster.index` *Finds subjects related to same cluster*

Description

Finds subjects related to same cluster

Usage

```
cluster.index(clusters, index.type = FALSE, num = NULL, Rindex = 0,
  mat = NULL, return.all = FALSE, code.na = NA)
```

Arguments

`clusters`	list of indeces
`index.type`	if TRUE then already list of integers of index.type
`num`	to get numbering according to num-type in separate columns
`Rindex`	index starts with 1, in C is it is 0
`mat`	to return matrix of indeces
`return.all`	return all arguments
`code.na`	how to code missing values

Author(s)

Klaus Holst, Thomas Scheike

References

Cluster indeces

See Also

familycluster.index familyclusterWithProbands.index

Examples

```
i<-c(1,1,2,2,1,3)
d<- cluster.index(i)
print(d)

type<-c("m","f","m","c","c","c")
d<- cluster.index(i,num=type,Rindex=1)
print(d)
```

---

concordanceCor *Concordance Computes concordance and casewise concordance*

Description

Concordance for Twins

Usage

```
concordanceCor(object, cif1, cif2 = NULL, messages = TRUE,
  model = NULL, coefs = NULL, ...)
```

Arguments

object	Output from the cor.cif, rr.cif or or.cif function
cif1	Marginal cumulative incidence
cif2	Marginal cumulative incidence of other cause (cause2) if it is different from cause1
messages	To print messages
model	Specfifies wich model that is considered if object not given.
coefs	Specfifies dependence parameters if object is not given.
...	Extra arguments, not used.

Details

The concordance is the probability that both twins have experienced the event of interest and is defined as

$$cor(t) = P(T_1 \leq t, \epsilon_1 = 1, T_2 \leq t, \epsilon_2 = 1)$$

Similarly, the casewise concordance is

$$casewise(t) = \frac{cor(t)}{P(T_1 \leq t, \epsilon_1 = 1)}$$

that is the probability that twin "2" has the event given that twins "1" has.

Author(s)

Thomas Scheike

References

Estimating twin concordance for bivariate competing risks twin data Thomas H. Scheike, Klaus K. Holst and Jacob B. Hjelmborg, Statistics in Medicine 2014, 1193-1204

Estimating Twin Pair Concordance for Age of Onset. Thomas H. Scheike, Jacob V B Hjelmborg, Klaus K. Holst, 2015 in Behavior genetics DOI:10.1007/s10519-015-9729-3

`cor.cif` *Cross-odds-ratio, OR or RR risk regression for competing risks*

Description

Fits a parametric model for the log-cross-odds-ratio for the predictive effect of for the cumulative incidence curves for T_1 experiencing cause i given that T_2 has experienced a cause k :

$$\log(COR(i|k)) = h(\theta, z_1, i, z_2, k, t) =_{default} \theta^T z =$$

with the log cross odds ratio being

$$COR(i|k) = \frac{O(T_1 \leq t, cause_1 = i | T_2 \leq t, cause_2 = k)}{O(T_1 \leq t, cause_1 = i)}$$

the conditional odds divided by the unconditional odds, with the odds being, respectively

$$O(T_1 \leq t, cause_1 = i | T_2 \leq t, cause_1 = k) = \frac{P_x(T_1 \leq t, cause_1 = i | T_2 \leq t, cause_2 = k)}{P_x((T_1 \leq t, cause_1 = i)^c | T_2 \leq t, cause_2 = k)}$$

and

$$O(T_1 \leq t, cause_1 = i) = \frac{P_x(T_1 \leq t, cause_1 = i)}{P_x((T_1 \leq t, cause_1 = i)^c)}.$$

Here B^c is the complement event of B, P_x is the distribution given covariates (x are subject specific and z are cluster specific covariates), and $h()$ is a function that is the simple identity $\theta^T z$ by default.

Usage

```
cor.cif(cif, data, cause = NULL, times = NULL, cause1 = 1,
  cause2 = 1, cens.code = NULL, cens.model = "KM", Nit = 40,
  detail = 0, clusters = NULL, theta = NULL, theta.des = NULL,
  step = 1, sym = 0, weights = NULL, par.func = NULL,
  dpar.func = NULL, dimpar = NULL, score.method = "nlminb",
  same.cens = FALSE, censoring.weights = NULL, silent = 1, ...)
```

Arguments

cif	a model object from the comp.risk function with the marginal cumulative incidence of cause1, i.e., the event of interest, and whose odds the comparision is compared to the conditional odds given cause2
data	a data.frame with the variables.
cause	specifies the causes related to the death times, the value cens.code is the censoring value. When missing it comes from marginal cif.
times	time-vector that specifies the times used for the estimating euqations for the cross-odds-ratio estimation.
cause1	specificies the cause considered.
cause2	specificies the cause that is conditioned on.
cens.code	specificies the code for the censoring if NULL then uses the one from the marginal cif model.
cens.model	specified which model to use for the ICPW, KM is Kaplan-Meier alternatively it may be "cox"
Nit	number of iterations for Newton-Raphson algorithm.
detail	if 0 no details are printed during iterations, if 1 details are given.
clusters	specifies the cluster structure.
theta	specifies starting values for the cross-odds-ratio parameters of the model.
theta.des	specifies a regression design for the cross-odds-ratio parameters.
step	specifies the step size for the Newton-Raphson algorithm.
sym	specifies if symmetry is used in the model.
weights	weights for estimating equations.
par.func	parfunc
dpar.func	dparfunc
dimpar	dimpar
score.method	"nlminb", can also use "fisher-scoring".
same.cens	if true then censoring within clusters are assumed to be the same variable, default is independent censoring.
censoring.weights	these probabilities are used for the bivariate censoring dist.
silent	1 to suppress output about convergence related issues.
...	Not used.

Details

The OR dependence measure is given by

$$OR(i,k) = \log(\frac{O(T_1 \leq t, cause_1 = i | T_2 \leq t, cause_2 = k)}{O(T_1 \leq t, cause_1 = i) | T_2 \leq t, cause_2 = k)}$$

This measure is numerically more stabile than the COR measure, and is symetric in i,k.

The RR dependence measure is given by

$$RR(i,k) = \log(\frac{P(T_1 \leq t, cause_1 = i, T_2 \leq t, cause_2 = k)}{P(T_1 \leq t, cause_1 = i) P(T_2 \leq t, cause_2 = k)}$$

This measure is numerically more stabile than the COR measure, and is symetric in i,k.

The model is fitted under symmetry (sym=1), i.e., such that it is assumed that T_1 and T_2 can be interchanged and leads to the same cross-odd-ratio (i.e. $COR(i|k) = COR(k|i)$), as would be expected for twins or without symmetry as might be the case with mothers and daughters (sym=0).

$h()$ may be specified as an R-function of the parameters, see example below, but the default is that it is simply $\theta^T z$.

Value

returns an object of type 'cor'. With the following arguments:

`theta`	estimate of proportional odds parameters of model.
`var.theta`	variance for gamma.
`hess`	the derivative of the used score.
`score`	scores at final stage.
`score`	scores at final stage.
`theta.iid`	matrix of iid decomposition of parametric effects.

Author(s)

Thomas Scheike

References

Cross odds ratio Modelling of dependence for Multivariate Competing Risks Data, Scheike and Sun (2012), Biostatistics.

A Semiparametric Random Effects Model for Multivariate Competing Risks Data, Scheike, Zhang, Sun, Jensen (2010), Biometrika.

Examples

```
library("timereg")
data(multcif);
multcif$cause[multcif$cause==0] <- 2
zyg <- rep(rbinom(200,1,0.5),each=2)
theta.des <- model.matrix(~-1+factor(zyg))
```

```
times=seq(0.05,1,by=0.05) # to speed up computations use only these time-points
add<-comp.risk(Event(time,cause)~+1+cluster(id),data=multcif,cause=1,
               n.sim=0,times=times,model="fg",max.clust=NULL)
add2<-comp.risk(Event(time,cause)~+1+cluster(id),data=multcif,cause=2,
               n.sim=0,times=times,model="fg",max.clust=NULL)

out1<-cor.cif(add,data=multcif,cause1=1,cause2=1)
summary(out1)

out2<-cor.cif(add,data=multcif,cause1=1,cause2=1,theta.des=theta.des)
summary(out2)

##out3<-cor.cif(add,data=multcif,cause1=1,cause2=2,cif2=add2)
##summary(out3)
############################################################
# investigating further models using parfunc and dparfunc
############################################################
 ## Reduce Ex.Timings
set.seed(100)
prt<-simnordic.random(2000,cordz=2,cormz=5)
prt$status <-prt$cause
table(prt$status)

times <- seq(40,100,by=10)
cifmod <- comp.risk(Event(time,cause)~+1+cluster(id),data=prt,
                    cause=1,n.sim=0,
                    times=times,conservative=1,max.clust=NULL,model="fg")
theta.des <- model.matrix(~-1+factor(zyg),data=prt)

parfunc <- function(par,t,pardes)
{
par <- pardes %*% c(par[1],par[2]) +
       pardes %*% c( par[3]*(t-60)/12,par[4]*(t-60)/12)
par
}
head(parfunc(c(0.1,1,0.1,1),50,theta.des))

dparfunc <- function(par,t,pardes)
{
dpar <- cbind(pardes, t(t(pardes) * c( (t-60)/12,(t-60)/12)) )
dpar
}
head(dparfunc(c(0.1,1,0.1,1),50,theta.des))

names(prt)
or1 <- or.cif(cifmod,data=prt,cause1=1,cause2=1,theta.des=theta.des,
              same.cens=TRUE,theta=c(0.6,1.1,0.1,0.1),
              par.func=parfunc,dpar.func=dparfunc,dimpar=4,
              score.method="fisher.scoring",detail=1)
summary(or1)

 cor1 <- cor.cif(cifmod,data=prt,cause1=1,cause2=1,theta.des=theta.des,
```

```
                same.cens=TRUE,theta=c(0.5,1.0,0.1,0.1),
                par.func=parfunc,dpar.func=dparfunc,dimpar=4,
                control=list(trace=TRUE),detail=1)
summary(cor1)

### piecewise contant OR model
gparfunc <- function(par,t,pardes)
{
cuts <- c(0,80,90,120)
grop <- diff(t<cuts)
paru  <- (pardes[,1]==1) * sum(grop*par[1:3]) +
    (pardes[,2]==1) * sum(grop*par[4:6])
paru
}

dgparfunc <- function(par,t,pardes)
{
cuts <- c(0,80,90,120)
grop <- diff(t<cuts)
par1 <- matrix(c(grop),nrow(pardes),length(grop),byrow=TRUE)
parmz <- par1* (pardes[,1]==1)
pardz <- (pardes[,2]==1) * par1
dpar <- cbind( parmz,pardz)
dpar
}
head(dgparfunc(rep(0.1,6),50,theta.des))
head(gparfunc(rep(0.1,6),50,theta.des))

or1g <- or.cif(cifmod,data=prt,cause1=1,cause2=1,
               theta.des=theta.des, same.cens=TRUE,
               par.func=gparfunc,dpar.func=dgparfunc,
               dimpar=6,score.method="fisher.scoring",detail=1)
summary(or1g)
names(or1g)
head(or1g$theta.iid)
```

`count.history` *Counts the number of previous events of two types for recurrent events processes*

Description

Counts the number of previous events of two types for recurrent events processes

Usage

```
count.history(data, status = "status", id = "id", types = 1:2,
  names.count = "Count", lag = TRUE)
```

Arguments

data	data-frame
status	name of status
id	id
types	types of the events (code) related to status
names.count	name of Counts, for example Count1 Count2 when types=c(1,2)
lag	if true counts previously observed, and if lag=FALSE counts up to know

Author(s)

Thomas Scheike

Examples

```
########################################
## getting some rates to mimick
########################################

data(base1cumhaz)
data(base4cumhaz)
data(drcumhaz)
dr <- drcumhaz
base1 <- base1cumhaz
base4 <- base4cumhaz

#########################################################################
### simulating simple model that mimicks data
### now with two event types and second type has same rate as death rate
#########################################################################

rr <- simRecurrentII(1000,base1,base4,death.cumhaz=dr)
rr <-  count.history(rr)
dtable(rr,~"Count*"+status,level=1)
```

---

covarianceRecurrent *Estimation of covariance for bivariate recurrent events with terminal event*

Description

Estimation of probability of more that k events for recurrent events process where there is terminal event

Usage

```
covarianceRecurrent(data, type1, type2, status = "status",
  death = "death", start = "start", stop = "stop", id = "id",
  names.count = "Count")
```

Arguments

data	data-frame
type1	type of first event (code) related to status
type2	type of second event (code) related to status
status	name of status
death	name of death indicator
start	start stop call of Hist() of prodlim
stop	start stop call of Hist() of prodlim
id	id
names.count	name of count for number of previous event of different types, here generated by count.history()

Author(s)

Thomas Scheike

References

Scheike, Eriksson, Tribler (2019) The mean, variance and correlation for bivariate recurrent events with a terminal event, JRSS-C

Examples

```
########################################
## getting some data to work on
########################################
data(base1cumhaz)
data(base4cumhaz)
data(drcumhaz)
dr <- drcumhaz
base1 <- base1cumhaz
base4 <- base4cumhaz
rr <- simRecurrent(1000,base1,cumhaz2=base4,death.cumhaz=dr)
rr <- count.history(rr)
rr$strata <- 1
dtable(rr,~death+status)

covrp <- covarianceRecurrent(rr,1,2,status="status",death="death",
                        start="entry",stop="time",id="id",names.count="Count")
par(mfrow=c(1,3))
plot(covrp)

### with strata, each strata in matrix column, provides basis for fast Bootstrap
covrpS <- covarianceRecurrentS(rr,1,2,status="status",death="death",
        start="entry",stop="time",strata="strata",id="id",names.count="Count")
```

daggregate *aggregating for for data frames*

Description

aggregating for for data frames

Usage

```
daggregate(data, y = NULL, x = NULL, subset, ..., fun = "summary",
  regex = mets.options()$regex, missing = FALSE,
  remove.empty = FALSE, matrix = FALSE, silent = FALSE,
  na.action = na.pass, convert = NULL)
```

Arguments

data	data.frame
y	name of variable, or formula, or names of variables on data frame.
x	name of variable, or formula, or names of variables on data frame.
subset	subset expression
...	additional arguments to lower level functions
fun	function defining aggregation
regex	interpret x,y as regular expressions
missing	Missing used in groups (x)
remove.empty	remove empty groups from output
matrix	if TRUE a matrix is returned instead of an array
silent	suppress messages
na.action	How model.frame deals with 'NA's
convert	if TRUE try to coerce result into matrix. Can also be a user-defined function

Examples

```
data("sTRACE",package="timereg")
daggregate(iris, "^.e.al", x="Species", fun=cor, regex=TRUE)
daggregate(iris, Sepal.Length+Petal.Length ~Species, fun=summary)
daggregate(iris, log(Sepal.Length)+I(Petal.Length>1.5) ~ Species,
                fun=summary)
daggregate(iris, "*Length*", x="Species", fun=head)
daggregate(iris, "^.e.al", x="Species", fun=tail, regex=TRUE)
daggregate(sTRACE, status~ diabetes, fun=table)
daggregate(sTRACE, status~ diabetes+sex, fun=table)
daggregate(sTRACE, status + diabetes+sex ~ vf+I(wmi>1.4), fun=table)
daggregate(iris, "^.e.al", x="Species",regex=TRUE)
dlist(iris,Petal.Length+Sepal.Length ~ Species |Petal.Length>1.3 & Sepal.Length>5,
```

```
            n=list(1:3,-(3:1)))
daggregate(iris, I(Sepal.Length>7)~Species | I(Petal.Length>1.5))
daggregate(iris, I(Sepal.Length>7)~Species | I(Petal.Length>1.5),
                 fun=table)

dsum(iris, .~Species, matrix=TRUE, missing=TRUE)

par(mfrow=c(1,2))
data(iris)
drename(iris) <- ~.
daggregate(iris,'sepal*'~species|species!="virginica",fun=plot)
daggregate(iris,'sepal*'~I(as.numeric(species))|I(as.numeric(species))!=1,fun=summary)

dnumeric(iris) <- ~species
daggregate(iris,'sepal*'~species.n|species.n!=1,fun=summary)
```

Dbvn *Derivatives of the bivariate normal cumulative distribution function*

Description

Derivatives of the bivariate normal cumulative distribution function

Usage

```
Dbvn(p,design=function(p,...) {
     return(list(mu=cbind(p[1],p[1]),
                 dmu=cbind(1,1),
                 S=matrix(c(p[2],p[3],p[3],p[4]),ncol=2),
                 dS=rbind(c(1,0,0,0),c(0,1,1,0),c(0,0,0,1)))  )},
     Y=cbind(0,0))
```

Arguments

p	Parameter vector
design	Design function with defines mean, derivative of mean, variance, and derivative of variance with respect to the parameter p
Y	column vector where the CDF is evaluated

Author(s)

Klaus K. Holst

dby *Calculate summary statistics grouped by*

Description

Calculate summary statistics grouped by variable

Usage

```
dby(data, INPUT, ..., ID = NULL, ORDER = NULL, SUBSET = NULL,
  SORT = 0, COMBINE = !REDUCE, NOCHECK = FALSE, ARGS = NULL, NAMES,
  COLUMN = FALSE, REDUCE = FALSE, REGEX = mets.options()$regex,
  ALL = TRUE)
```

Arguments

data	Data.frame
INPUT	Input variables (character or formula)
...	functions
ID	id variable
ORDER	(optional) order variable
SUBSET	(optional) subset expression
SORT	sort order (id+order variable)
COMBINE	If TRUE result is appended to data
NOCHECK	No sorting or check for missing data
ARGS	Optional list of arguments to functions (...)
NAMES	Optional vector of column names
COLUMN	If TRUE do the calculations for each column
REDUCE	Reduce number of redundant rows
REGEX	Allow regular expressions
ALL	if FALSE only the subset will be returned

Details

Calculate summary statistics grouped by

dby2 for column-wise calculations

Author(s)

Klaus K. Holst and Thomas Scheike

Examples

```
n <- 4
k <- c(3,rbinom(n-1,3,0.5)+1)
N <- sum(k)
d <- data.frame(y=rnorm(N),x=rnorm(N),id=rep(seq(n),k),num=unlist(sapply(k,seq)))
d2 <- d[sample(nrow(d)),]

dby(d2, y~id, mean)
dby(d2, y~id + order(num), cumsum)

dby(d,y ~ id + order(num), dlag)
dby(d,y ~ id + order(num), dlag, ARGS=list(k=1:2))
dby(d,y ~ id + order(num), dlag, ARGS=list(k=1:2), NAMES=c("l1","l2"))

dby(d, y~id + order(num), mean=mean, csum=cumsum, n=length)
dby(d2, y~id + order(num), a=cumsum, b=mean, N=length, l1=function(x) c(NA,x)[-length(x)])

dby(d, y~id + order(num), nn=seq_along, n=length)
dby(d, y~id + order(num), nn=seq_along, n=length)

d <- d[,1:4]
dby(d, x<0) <- list(z=mean)
d <- dby(d, is.na(z), z=1)

f <- function(x) apply(x,1,min)
dby(d, y+x~id, min=f)

dby(d,y+x~id+order(num), function(x) x)

f <- function(x) { cbind(cumsum(x[,1]),cumsum(x[,2]))/sum(x)}
dby(d, y+x~id, f)

## column-wise
a <- d
dby2(a, mean, median, REGEX=TRUE) <- '^[y|x]'~id
a
## wildcards
dby2(a,'y*'+'x*'~id,mean)


## subset
dby(d, x<0) <- list(z=NA)
d
dby(d, y~id|x>-1, v=mean,z=1)
dby(d, y+x~id|x>-1, mean, median, COLUMN=TRUE)

dby2(d, y+x~id|x>0, mean, REDUCE=TRUE)

dby(d,y~id|x<0,mean,ALL=FALSE)

a <- iris
a <- dby(a,y=1)
```

```
dby(a,Species=="versicolor") <- list(y=2)
```

---

dcor *summary, tables, and correlations for data frames*

Description

summary, tables, and correlations for data frames

Usage

```
dcor(data, y = NULL, x = NULL, use = "pairwise.complete.obs", ...)
```

Arguments

data	if x is formula or names for data frame then data frame is needed.
y	name of variable, or fomula, or names of variables on data frame.
x	possible group variable
use	how to handle missing values
...	Optional additional arguments

Author(s)

Klaus K. Holst and Thomas Scheike

Examples

```
data("sTRACE",package="timereg")
dt<- sTRACE
dt$time2 <- dt$time^2
dt$wmi2 <- dt$wmi^2
head(dt)

dcor(dt)

dcor(dt,~time+wmi)
dcor(dt,~time+wmi,~vf+chf)
dcor(dt,time+wmi~vf+chf)

dcor(dt,c("time*","wmi*"),~vf+chf)
```

dcut *Cutting, sorting, rm (removing), rename for data frames*

Description

Cut variables, if breaks are given these are used, otherwise cuts into using group size given by probs, or equispace groups on range. Default is equally sized groups if possible

Usage

```
dcut(data, y = NULL, x = NULL, breaks = 4, probs = NULL,
  equi = FALSE, regex = mets.options()$regex, sep = NULL,
  na.rm = TRUE, labels = NULL, all = FALSE, ...)
```

Arguments

data	if x is formula or names for data frame then data frame is needed.
y	name of variable, or fomula, or names of variables on data frame.
x	name of variable, or fomula, or names of variables on data frame.
breaks	number of breaks, for variables or vector of break points,
probs	groups defined from quantiles
equi	for equi-spaced breaks
regex	for regular expressions.
sep	seperator for naming of cut names.
na.rm	to remove NA for grouping variables.
labels	to use for cut groups
all	to do all variables, even when breaks are not unique
...	Optional additional arguments

Author(s)

Klaus K. Holst and Thomas Scheike

Examples

```
data("sTRACE",package="timereg")
sTRACE$age2 <- sTRACE$age^2
sTRACE$age3 <- sTRACE$age^3

mm <- dcut(sTRACE,~age+wmi)
head(mm)

mm <- dcut(sTRACE,catage4+wmi4~age+wmi)
head(mm)
```

```
mm <- dcut(sTRACE,~age+wmi,breaks=c(2,4))
head(mm)

mm <- dcut(sTRACE,c("age","wmi"))
head(mm)

mm <- dcut(sTRACE,~.)
head(mm)

mm <- dcut(sTRACE,c("age","wmi"),breaks=c(2,4))
head(mm)

gx <- dcut(sTRACE$age)
head(gx)


## Removes all cuts variables with these names wildcards
mm1 <- drm(mm,c("*.2","*.4"))
head(mm1)

## wildcards, for age, age2, age4 and wmi
head(dcut(mm,c("a*","?m*")))

## with direct asignment
drm(mm) <- c("*.2","*.4")
head(mm)

dcut(mm) <- c("age","*m*")
dcut(mm) <- ageg1+wmig1~age+wmi
head(mm)

############################
## renaming
############################

head(mm)
drename(mm, ~Age+Wmi) <- c("wmi","age")
head(mm)
mm1 <- mm

## all names to lower
drename(mm1) <- ~.
head(mm1)

## A* to lower
mm2 <-  drename(mm,c("A*","W*"))
head(mm2)
drename(mm) <- "A*"
head(mm)

dd <- data.frame(A_1=1:2,B_1=1:2)
funn <- function(x) gsub("_",".",x)
drename(dd) <- ~.
```

```
drename(dd,fun=funn) <- ~.
names(dd)
```

`dermalridges` *Dermal ridges data (families)*

Description

Data on dermal ridge counts in left and right hand in (nuclear) families

Format

Data on 50 families with ridge counts in left and right hand for moter, father and each child. Family id in 'family' and gender and child number in 'sex' and 'child'.

Source

Sarah B. Holt (1952). Genetics of dermal ridges: bilateral asymmetry in finger ridge-counts. Annals of Eugenics 17 (1), pp.211–231. DOI: 10.1111/j.1469-1809.1952.tb02513.x

Examples

```
data(dermalridges)
fast.reshape(dermalridges,id="family",varying=c("child.left","child.right","sex"))
```

`dermalridgesMZ` *Dermal ridges data (monozygotic twins)*

Description

Data on dermal ridge counts in left and right hand in (nuclear) families

Format

Data on dermal ridge counts (left and right hand) in 18 monozygotic twin pairs.

Source

Sarah B. Holt (1952). Genetics of dermal ridges: bilateral asymmetry in finger ridge-counts. Annals of Eugenics 17 (1), pp.211–231. DOI: 10.1111/j.1469-1809.1952.tb02513.x

Examples

```
data(dermalridgesMZ)
fast.reshape(dermalridgesMZ,id="id",varying=c("left","right"))
```

divide.conquer *Split a data set and run function*

Description

Split a data set and run function

Usage

```
divide.conquer(func = NULL, data, size, splits, id = NULL, ...)
```

Arguments

func	called function
data	data-frame
size	size of splits
splits	number of splits (ignored if size is given)
id	optional cluster variable
...	Additional arguments to lower level functions

Author(s)

Thomas Scheike, Klaus K. Holst

Examples

```
library(timereg)
data(TRACE)
res <- divide.conquer(prop.odds,TRACE,
     formula=Event(time,status==9)~chf+vf+age,n.sim=0,size=200)
```

divide.conquer.timereg *Split a data set and run function from timereg and aggregate*

Description

Split a data set and run function of cox-aalen type and aggregate results

Usage

```
divide.conquer.timereg(func = NULL, data, size, ...)
```

Arguments

func	called function
data	data-frame
size	size of splits
...	Additional arguments to lower level functions

Author(s)

Thomas Scheike, Klaus K. Holst

Examples

```
library(timereg)
data(TRACE)
a <- divide.conquer.timereg(prop.odds,TRACE,
                            formula=Event(time,status==9)~chf+vf+age,n.sim=0,size=200)
coef(a)
a2 <- divide.conquer.timereg(prop.odds,TRACE,
                             formula=Event(time,status==9)~chf+vf+age,n.sim=0,size=500)
coef(a2)

if (interactive()) {
par(mfrow=c(1,1))
plot(a,xlim=c(0,8),ylim=c(0,0.01))
par(new=TRUE)
plot(a2,xlim=c(0,8),ylim=c(0,0.01))
}
```

---

dlag *Lag operator*

Description

Lag operator

Usage

```
dlag(data, x, k = 1, combine = TRUE, simplify = TRUE, names, ...)
```

Arguments

data	data.frame or vector
x	optional column names or formula
k	lag (vector of integers)
combine	combine results with original data.frame
simplify	Return vector if possible
names	optional new column names
...	additional arguments to lower level functions

Examples

```
d <- data.frame(y=1:10,x=c(10:1))
dlag(d,k=1:2)
dlag(d,~x,k=0:1)
dlag(d$x,k=1)
dlag(d$x,k=-1:2, names=letters[1:4])
```

---

dprint *list, head, print, tail*

Description

listing for data frames

Usage

```
dprint(data, y = NULL, n = 0, ..., x = NULL)
```

Arguments

data	if x is formula or names for data frame then data frame is needed.
y	name of variable, or fomula, or names of variables on data frame.
n	Index of observations to print (default c(1:nfirst, n-nlast:nlast)
...	Optional additional arguments (nfirst,nlast, and print options)
x	possible group variable

Author(s)

Klaus K. Holst and Thomas Scheike

Examples

```
n <- 20
m <- lava::lvm(letters)
d <- lava::sim(m,n)

dlist(d,~a+b+c)
dlist(d,~a+b+c|a<0 & b>0)
## listing all :
dlist(d,~a+b+c|a<0 & b>0,n=0)
dlist(d,a+b+c~I(d>0)|a<0 & b>0)
dlist(d,.~I(d>0)|a<0 & b>0)
dlist(d,~a+b+c|a<0 & b>0, nlast=0)
dlist(d,~a+b+c|a<0 & b>0, nfirst=3, nlast=3)
dlist(d,~a+b+c|a<0 & b>0, 1:5)
dlist(d,~a+b+c|a<0 & b>0, -(5:1))
dlist(d,~a+b+c|a<0 & b>0, list(1:5,50:55,-(5:1)))
dprint(d,a+b+c ~ I(d>0) |a<0 & b>0, list(1:5,50:55,-(5:1)))
```

drcumhaz *Rate for leaving HPN program for patients of Copenhagen*

Description

Rate for leaving HPN program for patients of Copenhagen

Source

Estimated data

dreg *Regression for data frames with dutility call*

Description

Regression for data frames with dutility call

Usage

```
dreg(data, y, x = NULL, z = NULL, x.oneatatime = TRUE,
  x.base.names = NULL, z.arg = c("clever", "base", "group",
  "condition"), fun. = lm, summary. = summary, regex = FALSE,
  convert = NULL, doSummary = TRUE, special = NULL, equal = TRUE,
  test = 1, ...)
```

Arguments

data	data frame
y	name of variable, or fomula, or names of variables on data frame.
x	name of variable, or fomula, or names of variables on data frame.
z	name of variable, or fomula, or names of variables on data frame.
x.oneatatime	x's one at a time
x.base.names	base covarirates
z.arg	what is Z, c("clever","base","group","condition"), clever decides based on type of Z, base means that Z is used as fixed baseline covaraites for all X, group means the analyses is done based on groups of Z, and condition means that Z specifies a condition on the data
fun.	function lm is default
summary.	summary to use
regex	regex
convert	convert

doSummary	doSummary or not
special	special's
equal	to do pairwise stuff
test	development argument
...	Additional arguments for fun

Author(s)

Klaus K. Holst, Thomas Scheike

Examples

```
##'
data(iris)
data <- iris
drename(iris) <- ~.
names(iris)
set.seed(1)
iris$time <- runif(nrow(iris))
iris$time1 <- runif(nrow(iris))
iris$status <- rbinom(nrow(iris),1,0.5)
iris$S1 <- with(iris,Surv(time,status))
iris$S2 <- with(iris,Surv(time1,status))
iris$id <- 1:nrow(iris)

mm <- dreg(iris,"*.length"~"*.width"|I(species=="setosa" & status==1))
mm <- dreg(iris,"*.length"~"*.width"|species+status)
mm <- dreg(iris,"*.length"~"*.width"|species)
mm <- dreg(iris,"*.length"~"*.width"|species+status,z.arg="group")

 ## Reduce Ex.Timings
y <- "S*"~"*.width"
xs <- dreg(iris,y,fun.=phreg)
xs <- dreg(iris,y,fun.=survdiff)

y <- "S*"~"*.width"
xs <- dreg(iris,y,x.oneatatime=FALSE,fun.=phreg)

## under condition
y <- S1~"*.width"|I(species=="setosa" & sepal.width>3)
xs <- dreg(iris,y,z.arg="condition",fun.=phreg)
xs <- dreg(iris,y,fun.=phreg)

## under condition
y <- S1~"*.width"|species=="setosa"
xs <- dreg(iris,y,z.arg="condition",fun.=phreg)
xs <- dreg(iris,y,fun.=phreg)

## with baseline  after |
y <- S1~"*.width"|sepal.length
xs <- dreg(iris,y,fun.=phreg)
```

```
## by group by species, not working
y <- S1~"*.width"|species
ss <- split(iris,paste(iris$species,iris$status))

xs <- dreg(iris,y,fun.=phreg)

## species as base, species is factor so assumes that this is grouping
y <- S1~"*.width"|species
xs <- dreg(iris,y,z.arg="base",fun.=phreg)

##  background var after | and then one of x's at at time
y <- S1~"*.width"|status+"sepal*"
xs <- dreg(iris,y,fun.=phreg)

##  background var after | and then one of x's at at time
##y <- S1~"*.width"|status+"sepal*"
##xs <- dreg(iris,y,x.oneatatime=FALSE,fun.=phreg)
##xs <- dreg(iris,y,fun.=phreg)

##  background var after | and then one of x's at at time
##y <- S1~"*.width"+factor(species)
##xs <- dreg(iris,y,fun.=phreg)
##xs <- dreg(iris,y,fun.=phreg,x.oneatatime=FALSE)

y <- S1~"*.width"|factor(species)
xs <- dreg(iris,y,z.arg="base",fun.=phreg)

y <- S1~"*.width"|cluster(id)+factor(species)
xs <- dreg(iris,y,z.arg="base",fun.=phreg)
xs <- dreg(iris,y,z.arg="base",fun.=coxph)

## under condition with groups
y <- S1~"*.width"|I(sepal.length>4)
xs <- dreg(subset(iris,species=="setosa"),y,z.arg="group",fun.=phreg)

## under condition with groups
y <- S1~"*.width"+I(log(sepal.length))|I(sepal.length>4)
xs <- dreg(subset(iris,species=="setosa"),y,z.arg="group",fun.=phreg)

y <- S1~"*.width"+I(dcut(sepal.length))|I(sepal.length>4)
xs <- dreg(subset(iris,species=="setosa"),y,z.arg="group",fun.=phreg)

ff <- function(formula,data,...) {
 ss <- survfit(formula,data,...)
 kmplot(ss,...)
 return(ss)
}

if (interactive()) {
dcut(iris) <- ~"*.width"
y <- S1~"*.4"|I(sepal.length>4)
par(mfrow=c(1,2))
```

```
xs <- dreg(iris,y,fun.=ff)
}
```

---

drelevel *relev levels for data frames*

Description

levels shows levels for variables in data frame, relevel relevels a factor in data.frame

Usage

```
drelevel(data, y = NULL, x = NULL, ref = NULL, newlevels = NULL,
  regex = mets.options()$regex, sep = NULL, overwrite = FALSE, ...)
```

Arguments

data	if x is formula or names for data frame then data frame is needed.
y	name of variable, or fomula, or names of variables on data frame.
x	name of variable, or fomula, or names of variables on data frame.
ref	new reference variable
newlevels	to combine levels of factor in data frame
regex	for regular expressions.
sep	seperator for naming of cut names.
overwrite	to overwrite variable
...	Optional additional arguments

Author(s)

Klaus K. Holst and Thomas Scheike

Examples

```
data(mena)
dstr(mena)
dfactor(mena)  <- ~twinnum
dnumeric(mena) <- ~twinnum.f

dstr(mena)

mena2 <- drelevel(mena,"cohort",ref="(1980,1982]")
mena2 <- drelevel(mena,~cohort,ref="(1980,1982]")
mena2 <- drelevel(mena,cohortII~cohort,ref="(1980,1982]")
```

```
dlevels(mena)
dlevels(mena2)
drelevel(mena,ref="(1975,1977]")  <-  ~cohort
drelevel(mena,ref="(1980,1982]")  <-  ~cohort
dlevels(mena,"coh*")
dtable(mena,"coh*",level=1)

### level 1 of zyg as baseline for new variable
drelevel(mena,ref=1) <- ~zyg
drelevel(mena,ref=c("DZ","[1973,1975]")) <- ~ zyg+cohort
drelevel(mena,ref=c("DZ","[1973,1975]")) <- zygdz+cohort.early~ zyg+cohort
### level 2 of zyg and cohort as baseline for new variables
drelevel(mena,ref=2) <- ~ zyg+cohort
dlevels(mena)

##################### combining factor levels with newlevels argument

dcut(mena,labels=c("I","II","III","IV")) <- cat4~agemena
dlevels(drelevel(mena,~cat4,newlevels=1:3))
dlevels(drelevel(mena,ncat4~cat4,newlevels=3:2))
drelevel(mena,newlevels=3:2) <- ncat4~cat4
dlevels(mena)

dlevels(drelevel(mena,nca4~cat4,newlevels=list(c(1,4),2:3)))

drelevel(mena,newlevels=list(c(1,4),2:3)) <- nca4..2 ~ cat4
dlevels(mena)

drelevel(mena,newlevels=list("I-III"=c("I","II","III"),"IV"="IV")) <- nca4..3 ~ cat4
dlevels(mena)

drelevel(mena,newlevels=list("I-III"=c("I","II","III"))) <- nca4..4 ~ cat4
dlevels(mena)

drelevel(mena,newlevels=list(group1=c("I","II","III"))) <- nca4..5 ~ cat4
dlevels(mena)

drelevel(mena,newlevels=list(g1=c("I","II","III"),g2="IV")) <- nca4..6 ~ cat4
dlevels(mena)
```

---

dsort *Sort data frame*

Description

Sort data according to columns in data frame

Usage

```
dsort(data, x, ..., decreasing = FALSE, return.order = FALSE)
```

Arguments

data	Data frame
x	variable to order by
...	additional variables to order by
decreasing	sort order (vector of length x)
return.order	return order

Value

data.frame

Examples

```
data(data="hubble",package="lava")
dsort(hubble, "sigma")
dsort(hubble, hubble$sigma,"v")
dsort(hubble,~sigma+v)
dsort(hubble,~sigma-v)

## with direct asignment
dsort(hubble) <- ~sigma-v
```

---

dspline *Simple linear spline*

Description

Constructs simple linear spline on a data frame using the formula syntax of dutils that is adds (x-cuti)* (x>cuti) to the data-set for each knot of the spline. The full spline is thus given by x and spline variables added to the data-set.

Usage

```
dspline(data, y = NULL, x = NULL, breaks = 4, probs = NULL,
  equi = FALSE, regex = mets.options()$regex, sep = NULL,
  na.rm = TRUE, labels = NULL, all = FALSE, ...)
```

Arguments

data	if x is formula or names for data frame then data frame is needed.
y	name of variable, or fomula, or names of variables on data frame.
x	name of variable, or fomula, or names of variables on data frame.
breaks	number of breaks, for variables or vector of break points,
probs	groups defined from quantiles
equi	for equi-spaced breaks

regex	for regular expressions.
sep	seperator for naming of cut names.
na.rm	to remove NA for grouping variables.
labels	to use for cut groups
all	to do all variables, even when breaks are not unique
...	Optional additional arguments

Author(s)

Thomas Scheike

Examples

```
data(TRACE)
TRACE <- dspline(TRACE,~wmi,breaks=c(1,1.3,1.7))
cca <- coxph(Surv(time,status==9)~age+vf+chf+wmi,data=TRACE)
cca2 <- coxph(Surv(time,status==9)~age+wmi+vf+chf+wmi.spline1+wmi.spline2+wmi.spline3,data=TRACE)
anova(cca,cca2)

nd=data.frame(age=50,vf=0,chf=0,wmi=seq(0.4,3,by=0.01))
nd <- dspline(nd,~wmi,breaks=c(1,1.3,1.7))
pl <- predict(cca2,newdata=nd)
plot(nd$wmi,pl,type="l")
```

---

dtable *tables for data frames*

Description

tables for data frames

Usage

```
dtable(data, y = NULL, x = NULL, ..., level = -1, response = NULL,
  flat = TRUE, total = FALSE, prop = FALSE, summary = NULL)
```

Arguments

data	if x is formula or names for data frame then data frame is needed.
y	name of variable, or fomula, or names of variables on data frame.
x	name of variable, or fomula, or names of variables on data frame.
...	Optional additional arguments
level	1 for all marginal tables, 2 for all 2 by 2 tables, and null for the full table, possible versus group variable

response	For level=2, only produce tables with columns given by 'response' (index)
flat	produce flat tables
total	add total counts/proportions
prop	Proportions instead of counts (vector of margins)
summary	summary function

Author(s)

Klaus K. Holst and Thomas Scheike

Examples

```
data("sTRACE",package="timereg")

dtable(sTRACE,~status)
dtable(sTRACE,~status+vf)
dtable(sTRACE,~status+vf,level=1)
dtable(sTRACE,~status+vf,~chf+diabetes)

dtable(sTRACE,c("*f*","status"),~diabetes)
dtable(sTRACE,c("*f*","status"),~diabetes, level=2)
dtable(sTRACE,c("*f*","status"),level=1)

dtable(sTRACE,~"*f*"+status,level=1)
dtable(sTRACE,~"*f*"+status+I(wmi>1.4)|age>60,level=2)
dtable(sTRACE,"*f*"+status~I(wmi>0.5)|age>60,level=1)
dtable(sTRACE,status~dcut(age))

dtable(sTRACE,~status+vf+sex|age>60)
dtable(sTRACE,status+vf+sex~+1|age>60, level=2)
dtable(sTRACE,.~status+vf+sex|age>60,level=1)
dtable(sTRACE,status+vf+sex~diabetes|age>60)
dtable(sTRACE,status+vf+sex~diabetes|age>60, flat=FALSE)

dtable(sTRACE,status+vf+sex~diabetes|age>60, level=1)
dtable(sTRACE,status+vf+sex~diabetes|age>60, level=2)

dtable(sTRACE,status+vf+sex~diabetes|age>60, level=2, prop=1, total=TRUE)
dtable(sTRACE,status+vf+sex~diabetes|age>60, level=2, prop=2, total=TRUE)
dtable(sTRACE,status+vf+sex~diabetes|age>60, level=2, prop=1:2, summary=summary)
```

---

dtransform *Transform that allows condition*

Description

Defines new variables under condition for data frame

Usage

```
dtransform(data, ...)
```

Arguments

data	is data frame
...	new variable definitions including possible if condition

Examples

```
data(mena)

xx <- dtransform(mena,ll=log(agemena)+twinnum)

xx <- dtransform(mena,ll=log(agemena)+twinnum,agemena<15)
xx <- dtransform(xx  ,ll=100+agemena,ll2=1000,agemena>15)
dsummary(xx,ll+ll2~I(agemena>15))
```

`easy.binomial.twostage`

Fits two-stage binomial for describing depdendence in binomial data using marginals that are on logistic form using the binomial.twostage funcion, but call is different and easier and the data manipulation is build into the function. Useful in particular for family design data.

Description

If clusters contain more than two times, the algoritm uses a compososite likelihood based on the pairwise bivariate models.

Usage

```
easy.binomial.twostage(margbin = NULL, data = sys.parent(),
  score.method = "fisher.scoring", response = "response", id = "id",
  Nit = 60, detail = 0, silent = 1, weights = NULL,
  control = list(), theta = NULL, theta.formula = NULL,
  desnames = NULL, deshelp = 0, var.link = 1, iid = 1, step = 1,
  model = "plackett", marginal.p = NULL, strata = NULL,
  max.clust = NULL, se.clusters = NULL)
```

Arguments

margbin	Marginal binomial model
data	data frame
score.method	Scoring method
response	name of response variable in data frame

id	name of cluster variable in data frame
Nit	Number of iterations
detail	Detail for more output for iterations
silent	Debug information
weights	Weights for log-likelihood, can be used for each type of outcome in 2x2 tables.
control	Optimization arguments
theta	Starting values for variance components
theta.formula	design for depedence, either formula or design function
desnames	names for dependence parameters
deshelp	if 1 then prints out some data sets that are used, on on which the design function operates
var.link	Link function for variance
iid	Calculate i.i.d. decomposition
step	Step size
model	model
marginal.p	vector of marginal probabilities
strata	strata for fitting
max.clust	max clusters used for i.i.d. decompostion
se.clusters	clusters for iid decomposition for roubst standard errors

Details

The reported standard errors are based on the estimated information from the likelihood assuming that the marginals are known. This gives correct standard errors in the case of the plackett distribution (OR model for dependence), but incorrect for the clayton-oakes types model. The OR model is often known as the ALR model. Our fitting procedures gives correct standard errors due to the ortogonality and is fast.

Examples

```
data(twinstut)
twinstut0 <- subset(twinstut, tvparnr<2300000)
twinstut <- twinstut0
twinstut$binstut <- (twinstut$stutter=="yes")*1
theta.des <- model.matrix( ~-1+factor(zyg),data=twinstut)
margbin <- glm(binstut~factor(sex)+age,data=twinstut,family=binomial())
bin <- binomial.twostage(margbin,data=twinstut,var.link=1,
         clusters=twinstut$tvparnr,theta.des=theta.des,detail=0,
                 score.method="fisher.scoring")
summary(bin)
lava::estimate(coef=bin$theta,vcov=bin$var.theta,f=function(p) exp(p))

twinstut$cage <- scale(twinstut$age)
theta.des <- model.matrix( ~-1+factor(zyg)+cage,data=twinstut)
bina <- binomial.twostage(margbin,data=twinstut,var.link=1,
```

```
         clusters=twinstut$tvparnr,theta.des=theta.des,detail=0)
summary(bina)

theta.des <- model.matrix( ~-1+factor(zyg)+factor(zyg)*cage,data=twinstut)
bina <- binomial.twostage(margbin,data=twinstut,var.link=1,
         clusters=twinstut$tvparnr,theta.des=theta.des)
summary(bina)

out <- easy.binomial.twostage(stutter~factor(sex)+age,data=twinstut,
                              response="binstut",id="tvparnr",var.link=1,
          theta.formula=~-1+factor(zyg1))
summary(out)

## refers to zygosity of first subject in eash pair : zyg1
## could also use zyg2 (since zyg2=zyg1 within twinpair's))
desfs <- function(x,num1="zyg1",namesdes=c("mz","dz","os"))
    c(x[num1]=="mz",x[num1]=="dz",x[num1]=="os")*1

out3 <- easy.binomial.twostage(binstut~factor(sex)+age,
                               data=twinstut, response="binstut",id="tvparnr",
                               var.link=1,theta.formula=desfs,
                               desnames=c("mz","dz","os"))
summary(out3)

 ## Reduce Ex.Timings
n <- 10000
set.seed(100)
dd <- simBinFam(n,beta=0.3)
binfam <- fast.reshape(dd,varying=c("age","x","y"))
## mother, father, children  (ordered)
head(binfam)

########### ########### ########### ########### ########### ###########
####  simple analyses of binomial family data
########### ########### ########### ########### ########### ###########
desfs <- function(x,num1="num1",num2="num2")
{
     pp <- 1*(((x[num1]=="m")*(x[num2]=="f"))|(x[num1]=="f")*(x[num2]=="m"))
     pc <- (x[num1]=="m" | x[num1]=="f")*(x[num2]=="b1" | x[num2]=="b2")*1
     cc <- (x[num1]=="b1")*(x[num2]=="b1" | x[num2]=="b2")*1
     c(pp,pc,cc)
}

ud <- easy.binomial.twostage(y~+1,data=binfam,
     response="y",id="id",
     theta.formula=desfs,desnames=c("pp","pc","cc"))
summary(ud)

udx <- easy.binomial.twostage(y~+x,data=binfam,
     response="y",id="id",
     theta.formula=desfs,desnames=c("pp","pc","cc"))
summary(udx)
```

```
########### ########### ########### ########### ########### ###########
####  now allowing parent child POR to be different for mother and father
########### ########### ########### ########### ########### ###########

desfsi <- function(x,num1="num1",num2="num2")
{
    pp <- (x[num1]=="m")*(x[num2]=="f")*1
    mc <- (x[num1]=="m")*(x[num2]=="b1" | x[num2]=="b2")*1
    fc <- (x[num1]=="f")*(x[num2]=="b1" | x[num2]=="b2")*1
    cc <- (x[num1]=="b1")*(x[num2]=="b1" | x[num2]=="b2")*1
    c(pp,mc,fc,cc)
}

udi <- easy.binomial.twostage(y~+1,data=binfam,
     response="y",id="id",
     theta.formula=desfsi,desnames=c("pp","mother-child","father-child","cc"))
summary(udi)

##now looking to see if interactions with age or age influences marginal models
##converting factors to numeric to make all involved covariates numeric
##to use desfai2 rather then desfai that works on binfam

nbinfam <- binfam
nbinfam$num <- as.numeric(binfam$num)
head(nbinfam)

desfsai <- function(x,num1="num1",num2="num2")
{
    pp <- (x[num1]=="m")*(x[num2]=="f")*1
### av age for pp=1 i.e parent pairs
    agepp <- ((as.numeric(x["age1"])+as.numeric(x["age2"]))/2-30)*pp
    mc <- (x[num1]=="m")*(x[num2]=="b1" | x[num2]=="b2")*1
    fc <- (x[num1]=="f")*(x[num2]=="b1" | x[num2]=="b2")*1
    cc <- (x[num1]=="b1")*(x[num2]=="b1" | x[num2]=="b2")*1
    agecc <- ((as.numeric(x["age1"])+as.numeric(x["age2"]))/2-12)*cc
    c(pp,agepp,mc,fc,cc,agecc)
}

desfsai2 <- function(x,num1="num1",num2="num2")
{
    pp <- (x[num1]==1)*(x[num2]==2)*1
    agepp <- (((x["age1"]+x["age2"]))/2-30)*pp ### av age for pp=1 i.e parent pairs
    mc <- (x[num1]==1)*(x[num2]==3 | x[num2]==4)*1
    fc <- (x[num1]==2)*(x[num2]==3 | x[num2]==4)*1
    cc <- (x[num1]==3)*(x[num2]==3 | x[num2]==4)*1
    agecc <- ((x["age1"]+x["age2"])/2-12)*cc ### av age for children
    c(pp,agepp,mc,fc,cc,agecc)
}

udxai2 <- easy.binomial.twostage(y~+x+age,data=binfam,
     response="y",id="id",
     theta.formula=desfsai,
     desnames=c("pp","pp-age","mother-child","father-child","cc","cc-age"))
```

```
summary(udxai2)
```

`easy.survival.twostage` *Wrapper for easy fitting of Clayton-Oakes or bivariate Plackett models for bivariate survival data*

Description

Fits two-stage model for describing depdendence in survival data using marginals that are on cox or aalen form using the twostage funcion, but call is different and easier and the data manipulation build into the function. Useful in particular for family design data.

Usage

```
easy.survival.twostage(margsurv = NULL, data = sys.parent(),
  score.method = "nlminb", status = "status", time = "time",
  entry = NULL, id = "id", Nit = 60, detail = 0, silent = 1,
  weights = NULL, control = list(), theta = NULL,
  theta.formula = NULL, desnames = NULL, deshelp = 0, var.link = 1,
  iid = 1, step = 0.5, model = "plackett", marginal.surv = NULL,
  strata = NULL, se.clusters = NULL)
```

Arguments

`margsurv`	model
`data`	data frame
`score.method`	Scoring method
`status`	Status at exit time
`time`	Exit time
`entry`	Entry time
`id`	name of cluster variable in data frame
`Nit`	Number of iterations
`detail`	Detail for more output for iterations
`silent`	Debug information
`weights`	Weights for log-likelihood, can be used for each type of outcome in 2x2 tables.
`control`	Optimization arguments
`theta`	Starting values for variance components
`theta.formula`	design for depedence, either formula or design function
`desnames`	names for dependence parameters
`deshelp`	if 1 then prints out some data sets that are used, on on which the design function operates

var.link	Link function for variance (exp link)
iid	Calculate i.i.d. decomposition
step	Step size for newton-raphson
model	plackett or clayton-oakes model
marginal.surv	vector of marginal survival probabilities
strata	strata for fitting
se.clusters	clusters for iid decomposition for roubst standard errors

Details

If clusters contain more than two times, the algoritm uses a composite likelihood based on the pairwise bivariate models.

The reported standard errors are based on the estimated information from the likelihood assuming that the marginals are known.

Examples

```
library(mets)
data("prt",package="mets")
prtsam <- blocksample(prt,idvar="id",1e3,replace=FALSE)
margp <- coxph(Surv(time,status==1)~factor(country),data=prtsam)
fitco <- survival.twostage(margp,data=prtsam,clusters=prtsam$id)
summary(fitco)

des <- model.matrix(~-1+factor(zyg),data=prtsam);
fitco <- survival.twostage(margp,data=prtsam,theta.des=des,clusters=prtsam$id)
summary(fitco)
rm(prtsam)

dfam <- simSurvFam(1000)
dfam <- fast.reshape(dfam,var=c("x","time","status"))

desfs <- function(x,num1="num1",num2="num2")
{
pp <- (x[num1]=="m")*(x[num2]=="f")*1   ## mother-father
pc <- (x[num1]=="m" | x[num1]=="f")*(x[num2]=="b1" | x[num2]=="b2")*1 ## mother-child
cc <- (x[num1]=="b1")*(x[num2]=="b1" | x[num2]=="b2")*1               ## child-child
c(pp,pc,cc)
}

marg <- coxph(Surv(time,status)~factor(num),data=dfam)
out3 <- easy.survival.twostage(marg,data=dfam,time="time",status="status",id="id",
             deshelp=0,
             score.method="fisher.scoring",theta.formula=desfs,
             model="plackett",
             desnames=c("parent-parent","parent-child","child-cild"),iid=1)
summary(out3)
```

EVaddGam *Relative risk for additive gamma model*

Description

Computes the relative risk for additive gamma model at time 0

Usage

```
EVaddGam(theta, x1, x2, thetades, ags)
```

Arguments

theta	theta
x1	x1
x2	x2
thetades	thetades
ags	ags

Author(s)

Thomas Scheike

References

Eriksson and Scheike (2015), Additive Gamma frailty models for competing risks data, Biometrics (2015)

Examples

```
lam0 <- c(0.5,0.3)
pars <- c(1,1,1,1,0,1)
## genetic random effects, cause1, cause2 and overall
parg <- pars[c(1,3,5)]
## environmental random effects, cause1, cause2 and overall
parc <- pars[c(2,4,6)]

## simulate competing risks with two causes with hazards 0.5 and 0.3
## ace for each cause, and overall ace
out <- simCompete.twin.ace(10000,parg,parc,0,2,lam0=lam0,overall=1,all.sum=1)

## setting up design for running the model
mm <- familycluster.index(out$cluster)
head(mm$familypairindex,n=10)
pairs <- matrix(mm$familypairindex,ncol=2,byrow=TRUE)
tail(pairs,n=12)
#
kinship <- (out[pairs[,1],"zyg"]=="MZ")+ (out[pairs[,1],"zyg"]=="DZ")*0.5
```

```
# dout <- make.pairwise.design.competing(pairs,kinship,
#          type="ace",compete=length(lam0),overall=1)
# head(dout$ant.rvs)
## MZ
# dim(dout$theta.des)
# dout$random.design[,,1]
## DZ
# dout$theta.des[,,nrow(pairs)]
# dout$random.design[,,nrow(pairs)]
#
# thetades <- dout$theta.des[,,1]
# x <- dout$random.design[,,1]
# x
##EVaddGam(rep(1,6),x[1,],x[3,],thetades,matrix(1,18,6))

# thetades <- dout$theta.des[,,nrow(out)/2]
# x <- dout$random.design[,,nrow(out)/2]
##EVaddGam(rep(1,6),x[1,],x[4,],thetades,matrix(1,18,6))
```

`eventpois` *Extract survival estimates from lifetable analysis*

Description

Summary for survival analyses via the 'lifetable' function

Usage

```
eventpois(object, ..., timevar, time, int.len, confint = FALSE,
  level = 0.95, individual = FALSE, length.out = 25)
```

Arguments

`object`	glm object (poisson regression)
`...`	Contrast arguments
`timevar`	Name of time variable
`time`	Time points (optional)
`int.len`	Time interval length (optional)
`confint`	If TRUE confidence limits are supplied
`level`	Level of confidence limits
`individual`	Individual predictions
`length.out`	Length of time vector

Details

Summary for survival analyses via the 'lifetable' function

Author(s)

Klaus K. Holst

`familycluster.index` *Finds all pairs within a cluster (family)*

Description

Finds all pairs within a cluster (family)

Usage

```
familycluster.index(clusters, index.type = FALSE, num = NULL,
  Rindex = 1)
```

Arguments

`clusters`	list of indeces
`index.type`	argument of cluster index
`num`	num
`Rindex`	index starts with 1 in R, and 0 in C

Author(s)

Klaus Holst, Thomas Scheike

References

Cluster indeces

See Also

cluster.index familyclusterWithProbands.index

Examples

```
i<-c(1,1,2,2,1,3)
d<- familycluster.index(i)
print(d)
```

`familyclusterWithProbands.index` *Finds all pairs within a cluster (famly) with the proband (case/control)*

Description

second column of pairs are the probands and the first column the related subjects

Usage

```
familyclusterWithProbands.index(clusters, probands, index.type = FALSE,
  num = NULL, Rindex = 1)
```

Arguments

`clusters`	list of indeces giving the clusters (families)
`probands`	list of 0,1 where 1 specifices which of the subjects that are probands
`index.type`	argument passed to other functions
`num`	argument passed to other functions
`Rindex`	index starts with 1, in C is it is 0

Author(s)

Klaus Holst, Thomas Scheike

References

Cluster indeces

See Also

familycluster.index cluster.index

Examples

```
i<-c(1,1,2,2,1,3)
p<-c(1,0,0,1,0,1)
d<- familyclusterWithProbands.index(i,p)
print(d)
```

fast.approx *Fast approximation*

Description

Fast approximation

Usage

```
fast.approx(time, new.time, equal = FALSE, type = c("nearest", "right",
  "left"), sorted = FALSE, ...)
```

Arguments

time	Original ordered time points
new.time	New time points
equal	If TRUE a list is returned with additional element
type	Type of matching, nearest index, nearest greater than or equal (right), number of elements smaller than y otherwise the closest value above new.time is returned.
sorted	Set to true if new.time is already sorted
...	Optional additional arguments

Author(s)

Klaus K. Holst

Examples

```
id <- c(1,1,2,2,7,7,10,10)
fast.approx(unique(id),id)

t <- 0:6
n <- c(-1,0,0.1,0.9,1,1.1,1.2,6,6.5)
fast.approx(t,n,type="left")
```

fast.pattern *Fast pattern*

Description

Fast pattern

Usage

```
fast.pattern(x, y, categories = 2, ...)
```

Arguments

x	Matrix (binary) of patterns. Optionally if y is also passed as argument, then the pattern matrix is defined as the elements agreeing in the two matrices.
y	Optional matrix argument with same dimensions as x (see above)
categories	Default 2 (binary)
...	Optional additional arguments

Author(s)

Klaus K. Holst

Examples

```
X <- matrix(rbinom(100,1,0.5),ncol=4)
fast.pattern(X)

X <- matrix(rbinom(100,3,0.5),ncol=4)
fast.pattern(X,categories=4)
```

---

fast.reshape *Fast reshape*

Description

Fast reshape/tranpose of data

Usage

```
fast.reshape(data, varying, id, num, sep = "", keep, idname = "id",
  numname = "num", factor = FALSE, idcombine = TRUE,
  labelnum = FALSE, labels, regex = mets.options()$regex,
  dropid = FALSE, ...)
```

Arguments

data	data.frame or matrix
varying	Vector of prefix-names of the time varying variables. Optional for Long->Wide reshaping.
id	id-variable. If omitted then reshape Wide->Long.
num	Optional number/time variable
sep	String seperating prefix-name with number/time
keep	Vector of column names to keep
idname	Name of id-variable (Wide->Long)
numname	Name of number-variable (Wide->Long)

factor	If true all factors are kept (otherwise treated as character)
idcombine	If TRUE and id is vector of several variables, the unique id is combined from all the variables. Otherwise the first variable is only used as identifier.
labelnum	If TRUE varying variables in wide format (going from long->wide) are labeled 1,2,3,... otherwise use 'num' variable. In long-format (going from wide->long) varying variables matching 'varying' prefix are only selected if their postfix is a number.
labels	Optional labels for the number variable
regex	Use regular expressions
dropid	Drop id in long format (default FALSE)
...	Optional additional arguments

Author(s)

Thomas Scheike, Klaus K. Holst

Examples

```
library("lava")
m <- lvm(c(y1,y2,y3,y4)~x)
d <- sim(m,5)
d
fast.reshape(d,"y")
fast.reshape(fast.reshape(d,"y"),id="id")

##### From wide-format
(dd <- fast.reshape(d,"y"))
## Same with explicit setting new id and number variable/column names
## and seperator "" (default) and dropping x
fast.reshape(d,"y",idname="a",timevar="b",sep="",keep=c())
## Same with 'reshape' list-syntax
fast.reshape(d,list(c("y1","y2","y3","y4")),labelnum=TRUE)

##### From long-format
fast.reshape(dd,id="id")
## Restrict set up within-cluster varying variables
fast.reshape(dd,"y",id="id")
fast.reshape(dd,"y",id="id",keep="x",sep=".")

#####
x <- data.frame(id=c(5,5,6,6,7),y=1:5,x=1:5,tv=c(1,2,2,1,2))
x
(xw <- fast.reshape(x,id="id"))
(xl <- fast.reshape(xw,c("y","x"),idname="id2",keep=c()))
(xl <- fast.reshape(xw,c("y","x","tv")))
(xw2 <- fast.reshape(xl,id="id",num="num"))
fast.reshape(xw2,c("y","x"),idname="id")

### more generally:
### varying=list(c("ym","yf","yb1","yb2"), c("zm","zf","zb1","zb2"))
```

```
### varying=list(c("ym","yf","yb1","yb2")))

##### Family cluster example
d <- mets:::simBinFam(3)
d
fast.reshape(d,var="y")
fast.reshape(d,varying=list(c("ym","yf","yb1","yb2")))

d <- sim(lvm(~y1+y2+ya),10)
d
(dd <- fast.reshape(d,"y"))
fast.reshape(d,"y",labelnum=TRUE)
fast.reshape(dd,id="id",num="num")
fast.reshape(dd,id="id",num="num",labelnum=TRUE)
fast.reshape(d,c(a="y"),labelnum=TRUE) ## New column name


##### Unbalanced data
m <- lvm(c(y1,y2,y3,y4)~ x+z1+z3+z5)
d <- sim(m,3)
d
fast.reshape(d,c("y","z"))

##### not-varying syntax:
fast.reshape(d,-c("x"))

##### Automatically define varying variables from trailing digits
fast.reshape(d)

##### Prostate cancer example
data(prt)
head(prtw <- fast.reshape(prt,"cancer",id="id"))
ftable(cancer1~cancer2,data=prtw)
rm(prtw)
```

---

ghaplos *ghaplos haplo-types for subjects of haploX data*

Description

ghaplos haplo-types for subjects of haploX data

Source

Simulated data

gof.phreg *GOF for Cox PH regression*

Description

Cumulative score process residuals for Cox PH regression p-values based on Lin, Wei, Ying resampling.

Usage

```
## S3 method for class 'phreg'
gof(object, n.sim = 1000, silent = 1, robust = NULL,
  ...)
```

Arguments

object	is phreg object
n.sim	number of simulations for score processes
silent	to show timing estimate will be produced for longer jobs
robust	to control wether robust dM_i(t) or dN_i are used for simulations
...	Additional arguments to lower level funtions

Author(s)

Thomas Scheike and Klaus K. Holst

Examples

```
data(TRACE)

m1 <- phreg(Surv(time,status==9)~vf+chf+diabetes,data=TRACE)
gg <- gof(m1)
par(mfrow=c(1,3))
plot(gg)

m1 <- phreg(Surv(time,status==9)~strata(vf)+chf+diabetes,data=TRACE)
## to get Martingale ~ dN based simulations
gg <- gof(m1)

## to get Martingale robust simulations, specify cluster in  call
m1 <- phreg(Surv(time,status==9)~chf+diabetes+cluster(id),data=TRACE)
gg <- gof(m1)
```

gofG.phreg *Stratified baseline graphical GOF test for Cox covariates in PH regression*

Description

Looks at stratified baseline in Cox model and plots all baselines versus each other to see if lines are straight, with 50 resample versions under the assumptiosn that the stratified Cox is correct

Usage

```
gofG.phreg(x, sim = 0, silent = 1, lm = TRUE, ...)
```

Arguments

x	phreg object
sim	to simulate som variation from cox model to put on graph
silent	to keep it absolutely silent
lm	addd line to plot, regressing the cumulatives on each other
...	Additional arguments to lower level funtions

Author(s)

Thomas Scheike and Klaus K. Holst

Examples

```
data(TRACE)

m1 <- phreg(Surv(time,status==9)~strata(vf)+chf+wmi,data=TRACE)
m2 <- phreg(Surv(time,status==9)~vf+strata(chf)+wmi,data=TRACE)
par(mfrow=c(2,2))

gofG.phreg(m1)
gofG.phreg(m2)

bplot(m1,log="y")
bplot(m2,log="y")
```

gofM.phreg *GOF for Cox covariates in PH regression*

Description

Cumulative residuals after model matrix for Cox PH regression p-values based on Lin, Wei, Ying resampling.

Usage

```
gofM.phreg(formula, data, offset = NULL, weights = NULL,
  modelmatrix = NULL, n.sim = 1000, silent = 1, ...)
```

Arguments

formula	formula for cox regression
data	data for model
offset	offset
weights	weights
modelmatrix	matrix for cumulating residuals
n.sim	number of simulations for score processes
silent	to keep it absolutely silent, otherwise timing estimate will be prduced for longer jobs.
...	Additional arguments to lower level funtions

Details

That is, computes

$$U(t) = \int_0^t M^t d\hat{M}$$

and resamples its asymptotic distribution.

This will show if the residuals are consistent with the model. Typically, M will be a design matrix for the continous covariates that gives for example the quartiles, and then the plot will show if for the different quartiles of the covariate the risk prediction is consistent over time (time x covariate interaction).

Author(s)

Thomas Scheike and Klaus K. Holst

Examples

```
library(mets)
data(TRACE)
set.seed(1)
TRACEsam <- blocksample(TRACE,idvar="id",replace=FALSE,100)

dcut(TRACEsam)  <- ~.
mm <- model.matrix(~-1+factor(wmicat.4),data=TRACEsam)
m1 <- gofM.phreg(Surv(time,status==9)~vf+chf+wmi,data=TRACEsam,modelmatrix=mm)
summary(m1)
if (interactive()) {
par(mfrow=c(2,2))
plot(m1)
}

m1 <- gofM.phreg(Surv(time,status==9)~strata(vf)+chf+wmi,data=TRACEsam,modelmatrix=mm)
summary(m1)

## cumulative sums in covariates, via design matrix mm
mm <- cumContr(TRACEsam$wmi,breaks=10,equi=TRUE)
m1 <- gofM.phreg(Surv(time,status==9)~strata(vf)+chf+wmi,data=TRACEsam,
  modelmatrix=mm,silent=0)
summary(m1)
```

---

gofZ.phreg *GOF for Cox covariates in PH regression*

Description

That is, computes

$$U(z,\tau) = \int_0^\tau M(z)^t d\hat{M}$$

and resamples its asymptotic distribution.

Usage

```
gofZ.phreg(formula, data, vars = NULL, offset = NULL, weights = NULL,
  breaks = 50, equi = FALSE, n.sim = 1000, silent = 1, ...)
```

Arguments

formula	formula for cox regression
data	data for model
vars	which variables to test for linearity
offset	offset
weights	weights

breaks	number of breaks for cumulatives in covarirate direction
equi	equidistant breaks or not
n.sim	number of simulations for score processes
silent	to keep it absolutely silent, otherwise timing estimate will be prduced for longer jobs.
...	Additional arguments to lower level funtions

Details

This will show if the residuals are consistent with the model evaulated in the z covariate. M is here chosen based on a grid (z_1, ..., z_m) and the different columns are $I(Z_i \leq z_l)$. for $l = 1, ..., m$. The process in z is resampled to find extreme values. The time-points of evuluation is by default 50 points, chosen as 2

The p-value is valid but depends on the chosen grid. When the number of break points are high this will give the orginal test of Lin, Wei and Ying for linearity, that is also computed in the timereg package.

Author(s)

Thomas Scheike and Klaus K. Holst

Examples

```
library(mets)
data(TRACE)
set.seed(1)
TRACEsam <- blocksample(TRACE,idvar="id",replace=FALSE,100)

## cumulative sums in covariates, via design matrix mm
 ## Reduce Ex.Timings
m1 <- gofZ.phreg(Surv(time,status==9)~strata(vf)+chf+wmi+age,data=TRACEsam)
summary(m1)
plot(m1,type="z")
```

Grandom.cif *Additive Random effects model for competing risks data for polygenetic modelling*

Description

Fits a random effects model describing the dependence in the cumulative incidence curves for subjects within a cluster. Given the gamma distributed random effects it is assumed that the cumulative incidence curves are indpendent, and that the marginal cumulative incidence curves are on additive form

$$P(T \leq t, cause = 1|x, z) = P_1(t, x, z) = 1 - exp(-x^T A(t) - tz^T \beta)$$

Usage

```
Grandom.cif(cif, data, cause = NULL, cif2 = NULL, times = NULL,
  cause1 = 1, cause2 = 1, cens.code = NULL, cens.model = "KM",
  Nit = 40, detail = 0, clusters = NULL, theta = NULL,
  theta.des = NULL, weights = NULL, step = 1, sym = 0,
  same.cens = FALSE, censoring.weights = NULL, silent = 1,
  var.link = 0, score.method = "fisher.scoring", entry = NULL,
  estimator = 1, trunkp = 1, admin.cens = NULL,
  random.design = NULL, ...)
```

Arguments

cif	a model object from the comp.risk function with the marginal cumulative incidence of cause2, i.e., the event that is conditioned on, and whose odds the comparision is made with respect to
data	a data.frame with the variables.
cause	specifies the causes related to the death times, the value cens.code is the censoring value.
cif2	specificies model for cause2 if different from cause1.
times	time points
cause1	cause of first coordinate.
cause2	cause of second coordinate.
cens.code	specificies the code for the censoring if NULL then uses the one from the marginal cif model.
cens.model	specified which model to use for the ICPW, KM is Kaplan-Meier alternatively it may be "cox"
Nit	number of iterations for Newton-Raphson algorithm.
detail	if 0 no details are printed during iterations, if 1 details are given.
clusters	specifies the cluster structure.
theta	specifies starting values for the cross-odds-ratio parameters of the model.
theta.des	specifies a regression design for the cross-odds-ratio parameters.
weights	weights for score equations.
step	specifies the step size for the Newton-Raphson algorith.m
sym	1 for symmetri and 0 otherwise
same.cens	if true then censoring within clusters are assumed to be the same variable, default is independent censoring.
censoring.weights	Censoring probabilities
silent	debug information
var.link	if var.link=1 then var is on log-scale.
score.method	default uses "nlminb" optimzer, alternatively, use the "fisher-scoring" algorithm.

entry	entry-age in case of delayed entry. Then two causes must be given.
estimator	estimator
trunkp	gives probability of survival for delayed entry, and related to entry-ages given above.
admin.cens	Administrative censoring
random.design	specifies a regression design of 0/1's for the random effects.
...	extra arguments.

Details

We allow a regression structure for the indenpendent gamma distributed random effects and their variances that may depend on cluster covariates.

random.design specificies the random effects for each subject within a cluster. This is a matrix of 1's and 0's with dimension n x d. With d random effects. For a cluster with two subjects, we let the random.design rows be v_1 and v_2. Such that the random effects for subject 1 is

$$v_1^T(Z_1, ..., Z_d)$$

, for d random effects. Each random effect has an associated parameter $(\lambda_1, ..., \lambda_d)$. By construction subjects 1's random effect are Gamma distributed with mean $\lambda_1/v_1^T\lambda$ and variance $\lambda_1/(v_1^T\lambda)^2$. Note that the random effect $v_1^T(Z_1, ..., Z_d)$ has mean 1 and variance $1/(v_1^T\lambda)$.

The parameters $(\lambda_1, ..., \lambda_d)$ are related to the parameters of the model by a regression construction $pard$ (d x k), that links the d λ parameters with the (k) underlying θ parameters

$$\lambda = pard\theta$$

Value

returns an object of type 'random.cif'. With the following arguments:

theta	estimate of parameters of model.
var.theta	variance for gamma.
hess	the derivative of the used score.
score	scores at final stage.
theta.iid	matrix of iid decomposition of parametric effects.

Author(s)

Thomas Scheike

References

A Semiparametric Random Effects Model for Multivariate Competing Risks Data, Scheike, Zhang, Sun, Jensen (2010), Biometrika.

Cross odds ratio Modelling of dependence for Multivariate Competing Risks Data, Scheike and Sun (2013), Biostatitistics.

Scheike, Holst, Hjelmborg (2014), LIDA, Estimating heritability for cause specific hazards based on twin data

Examples

```
## Reduce Ex.Timings
d <- simnordic.random(5000,delayed=TRUE,
      cordz=1.0,cormz=2,lam0=0.3,country=TRUE)
times <- seq(50,90,by=10)
addm<-comp.risk(Event(time,cause)~-1+factor(country)+cluster(id),data=d,
times=times,cause=1,max.clust=NULL)

### making group indidcator
mm <- model.matrix(~-1+factor(zyg),d)

out1m<-random.cif(addm,data=d,cause1=1,cause2=1,theta=1,
  theta.des=mm,same.cens=TRUE)
summary(out1m)

## this model can also be formulated as a random effects model
## but with different parameters
out2m<-Grandom.cif(addm,data=d,cause1=1,cause2=1,
   theta=c(0.5,1),step=1.0,
   random.design=mm,same.cens=TRUE)
summary(out2m)
1/out2m$theta
out1m$theta

######################################################################
################### ACE modelling of twin data ######################
######################################################################
### assume that zygbin gives the zygosity of mono and dizygotic twins
### 0 for mono and 1 for dizygotic twins. We now formulate and AC model
zygbin <- d$zyg=="DZ"

n <- nrow(d)
### random effects for each cluster
des.rv <- cbind(mm,(zygbin==1)*rep(c(1,0)),(zygbin==1)*rep(c(0,1)),1)
### design making parameters half the variance for dizygotic components
pardes <- rbind(c(1,0), c(0.5,0),c(0.5,0), c(0.5,0), c(0,1))

outacem <-Grandom.cif(addm,data=d,cause1=1,cause2=1,
same.cens=TRUE,theta=c(0.35,0.15),
            step=1.0,theta.des=pardes,random.design=des.rv)
summary(outacem)
```

---

hapfreqs *hapfreqs data set*

Description

hapfreqs data set

Source

Simulated data

`haplo.surv.discrete` *Discrete time to event haplo type analysis*

Description

Can be used for logistic regression when time variable is "1" for all id.

Usage

```
haplo.surv.discrete(X = NULL, y = "y", time.name = "time",
  Haplos = NULL, id = "id", desnames = NULL, designfunc = NULL,
  beta = NULL, no.opt = FALSE, method = "NR", stderr = TRUE,
  designMatrix = NULL, response = NULL, idhap = NULL,
  design.only = FALSE, covnames = NULL, fam = binomial,
  weights = NULL, offsets = NULL, idhapweights = NULL, ...)
```

Arguments

`X`	design matrix data-frame (sorted after id and time variable) with id time response and desnames
`y`	name of response (binary response with logistic link) from X
`time.name`	to sort after time for X
`Haplos`	(data.frame with id, haplo1, haplo2 (haplotypes (h)) and p=P(h|G)) haplotypes given as factor.
`id`	name of id variale from X
`desnames`	names for design matrix
`designfunc`	function that computes design given haplotypes h=(h1,h2) x(h)
`beta`	starting values
`no.opt`	optimization TRUE/FALSE
`method`	NR, nlm
`stderr`	to return only estimate
`designMatrix`	gives response and designMatrix directly not implemented (mush contain: p, id, idhap)
`response`	gives response and design directly designMatrix not implemented
`idhap`	name of id-hap variable to specify different haplotypes for different id
`design.only`	to return only design matrices for haplo-type analyses.
`covnames`	names of covariates to extract from object for regression
`fam`	family of models, now binomial default and only option
`weights`	weights following id for GLM
`offsets`	following id for GLM
`idhapweights`	weights following id-hap for GLM (WIP)
`...`	Additional arguments to lower level funtions lava::NR optimizer or nlm

Details

Cycle-specific logistic regression of haplo-type effects with known haplo-type probabilities. Given observed genotype G and unobserved haplotypes H we here mix out over the possible haplotypes using that P(H|G) is provided.

$$S(t|x,G)) = E(S(t|x,H)|G) = \sum_{h \in G} P(h|G)S(t|z,h)$$

so survival can be computed by mixing out over possible h given g.

Survival is based on logistic regression for the discrete hazard function of the form

$$logit(P(T = t|T \geq t, x, h)) = \alpha_t + x(h)\beta$$

where x(h) is a regression design of x and haplotypes $h = (h_1, h_2)$

Likelihood is maximized and standard errors assumes that P(H|G) is known.

The design over the possible haplotypes is constructed by merging X with Haplos and can be viewed by design.only=TRUE

Author(s)

Thomas Scheike

Examples

```
## some haplotypes of interest
types <- c("DCGCGCTCACG","DTCCGCTGACG","ITCAGTTGACG","ITCCGCTGAGG")

## some haplotypes frequencies for simulations
data(hapfreqs)

www <-which(hapfreqs$haplotype %in% types)
hapfreqs$freq[www]

baseline=hapfreqs$haplotype[9]
baseline

designftypes <- function(x,sm=0) {# {{{
hap1=x[1]
hap2=x[2]
if (sm==0) y <- 1*( (hap1==types) | (hap2==types))
if (sm==1) y <- 1*(hap1==types) + 1*(hap2==types)
return(y)
}# }}}

tcoef=c(-1.93110204,-0.47531630,-0.04118204,-1.57872602,-0.22176426,-0.13836416,
0.88830288,0.60756224,0.39802821,0.32706859)

data(hHaplos)
data(haploX)
```

```
haploX$time <- haploX$times
Xdes <- model.matrix(~factor(time),haploX)
colnames(Xdes) <- paste("X",1:ncol(Xdes),sep="")
X <- dkeep(haploX,~id+y+time)
X <- cbind(X,Xdes)
Haplos <- dkeep(ghaplos,~id+"haplo*"+p)
desnames=paste("X",1:6,sep="")   # six X's related to 6 cycles
out <- haplo.surv.discrete(X=X,y="y",time.name="time",
         Haplos=Haplos,desnames=desnames,designfunc=designftypes)
names(out$coef) <- c(desnames,types)
out$coef
 summary(out)
```

---

haploX *haploX covariates and response for haplo survival discrete survival*

Description

haploX covariates and response for haplo survival discrete survival

Source

Simulated data

---

interval.logitsurv.discrete *## uses HaploSurvival package of github install via devtools ## devtools::install_github("scheike/HaploSurvival") ## this is only used for simulations ## out <- simHaplo(1,100,tcoef,hapfreqs) Discrete time to event interval censored data*

Description

$$logit(P(T > t|x)) = log(G(t)) + x\beta$$

$$P(T > t|x) = \frac{1}{1 + G(t)exp(x\beta)}$$

Usage

```
interval.logitsurv.discrete(formula, data, beta = NULL, no.opt = FALSE,
  method = "NR", stderr = TRUE, weights = NULL, offsets = NULL,
  exp.link = 1, increment = 1, ...)
```

Arguments

formula	formula
data	data
beta	starting values
no.opt	optimization TRUE/FALSE
method	NR, nlm
stderr	to return only estimate
weights	weights following id for GLM
offsets	following id for GLM
exp.link	parametrize increments exp(alpha) > 0
increment	using increments dG(t)=exp(alpha) as parameters
...	Additional arguments to lower level funtions lava::NR optimizer or nlm

Details

Input are intervals given by]t_l,t_r] where t_r can be infinity for right-censored intervals When truly discrete]0,1] will be an observation at 1, and]j,j+1] will be an observation at j+1

Likelihood is maximized:

$$\prod P(T_i > t_{il}|x) - P(T_I > t_{ir}|x)$$

Author(s)

Thomas Scheike

Examples

```
data(ttpd)
out <- interval.logitsurv.discrete(Interval(entry,time2)~X1+X2+X3+X4,ttpd)
summary(out)

n <- 100
Z <- matrix(rbinom(n*4,1,0.5),n,4)
outsim <- simlogitSurvd(out$coef,Z)
outsim <- transform(outsim,left=time,right=time+1)
outsim <- dtransform(outsim,right=Inf,status==0)

outss <- interval.logitsurv.discrete(Interval(left,right)~+X1+X2+X3+X4,outsim)

Z <- matrix(0,5,4)
Z[2:5,1:4] <- diag(4)
pred <- predictlogitSurvd(out,se=FALSE)
plotSurvd(pred)

## simulations
n <- 100
Z <- matrix(rbinom(n*4,1,0.5),n,4)
outsim <- simlogitSurvd(out$coef,Z)
```

```
###
outsim <- transform(outsim,left=time,right=time+1)
outsim <- dtransform(outsim,right=Inf,status==0)

out$coef
outss <- interval.logitsurv.discrete(Interval(left,right)~+X1+X2+X3+X4,outsim)
summary(outss)
```

---

ipw *Inverse Probability of Censoring Weights*

Description

Internal function. Calculates Inverse Probability of Censoring Weights (IPCW) and adds them to a data.frame

Usage

```
ipw(formula, data, cluster, same.cens = FALSE, obs.only = TRUE,
  weight.name = "w", trunc.prob = FALSE, weight.name2 = "wt",
  indi.weight = "pr", cens.model = "aalen", pairs = FALSE,
  theta.formula = ~1, ...)
```

Arguments

formula	Formula specifying the censoring model
data	data frame
cluster	clustering variable
same.cens	For clustered data, should same censoring be assumed (bivariate probability calculated as mininum of the marginal probabilities)
obs.only	Return data with uncensored observations only
weight.name	Name of weight variable in the new data.frame
trunc.prob	If TRUE truncation probabilities are also calculated and stored in 'weight.name2' (based on Clayton-Oakes gamma frailty model)
weight.name2	Name of truncation probabilities
indi.weight	Name of individual censoring weight in the new data.frame
cens.model	Censoring model (default Aalens additive model)
pairs	For paired data (e.g. twins) only the complete pairs are returned (With pairs=TRUE)
theta.formula	Model for the dependence parameter in the Clayton-Oakes model (truncation only)
...	Additional arguments to censoring model

Author(s)

Klaus K. Holst

Examples

```
## Not run:
data("prt",package="mets")
prtw <- ipw(Surv(time,status==0)~country, data=prt[sample(nrow(prt),5000),],
            cluster="id",weight.name="w")
plot(0,type="n",xlim=range(prtw$time),ylim=c(0,1),xlab="Age",ylab="Probability")
count <- 0
for (l in unique(prtw$country)) {
    count <- count+1
    prtw <- prtw[order(prtw$time),]
    with(subset(prtw,country==l),
         lines(time,w,col=count,lwd=2))
}
legend("topright",legend=unique(prtw$country),col=1:4,pch=-1,lty=1)

## End(Not run)
```

---

ipw2 *Inverse Probability of Censoring Weights*

Description

Internal function. Calculates Inverse Probability of Censoring and Truncation Weights and adds them to a data.frame

Usage

```
ipw2(data, times = NULL, entrytime = NULL, time = "time",
  cause = "cause", same.cens = FALSE, cluster = NULL,
  pairs = FALSE, strata = NULL, obs.only = TRUE,
  cens.formula = NULL, cens.code = 0, pair.cweight = "pcw",
  pair.tweight = "ptw", pair.weight = "weights", cname = "cweights",
  tname = "tweights", weight.name = "indi.weights",
  prec.factor = 100, ...)
```

Arguments

data	data frame
times	possible time argument for speciying a maximum value of time tau=max(times), to specify when things are considered censored or not.
entrytime	nam of entry-time for truncation.
time	name of time variable on data frame.
cause	name of cause indicator on data frame.

same.cens	For clustered data, should same censoring be assumed and same truncation (bivariate probability calculated as mininum of the marginal probabilities)
cluster	name of clustering variable
pairs	For paired data (e.g. twins) only the complete pairs are returned (With pairs=TRUE)
strata	name of strata variable to get weights stratified.
obs.only	Return data with uncensored observations only
cens.formula	model for Cox models for truncation and right censoring times.
cens.code	censoring.code
pair.cweight	Name of weight variable in the new data.frame for right censorig of pairs
pair.tweight	Name of weight variable in the new data.frame for left truncation of pairs
pair.weight	Name of weight variable in the new data.frame for right censoring and left truncation of pairs
cname	Name of weight variable in the new data.frame for right censoring of individuals
tname	Name of weight variable in the new data.frame for left truncation of individuals
weight.name	Name of weight variable in the new data.frame for right censoring and left truncation of individuals
prec.factor	To let tied censoring and truncation times come after the death times.
...	Additional arguments to censoring model

Author(s)

Thomas Scheike

Examples

```
library("timereg")
set.seed(1)
d <- simnordic.random(3000,delayed=TRUE,ptrunc=0.7,
      cordz=0.5,cormz=2,lam0=0.3,country=FALSE)
d$strata <- as.numeric(d$country)+(d$zyg=="MZ")*4
times <- seq(60,100,by=10)
c1 <- comp.risk(Event(time,cause)~1+cluster(id),data=d,cause=1,
model="fg",times=times,max.clust=NULL,n.sim=0)
mm=model.matrix(~-1+zyg,data=d)
out1<-random.cif(c1,data=d,cause1=1,cause2=1,same.cens=TRUE,theta.des=mm)
summary(out1)
pc1 <- predict(c1,X=1,se=0)
plot(pc1)

dl <- d[!d$truncated,]
dl <- ipw2(dl,cluster="id",same.cens=TRUE,time="time",entrytime="entry",cause="cause",
           strata="strata",prec.factor=100)
cl <- comp.risk(Event(time,cause)~+1+
cluster(id),
data=dl,cause=1,model="fg",
weights=dl$indi.weights,cens.weights=rep(1,nrow(dl)),
```

```
            times=times,max.clust=NULL,n.sim=0)
pcl <- predict(cl,X=1,se=0)
lines(pcl$time,pcl$P1,col=2)
mm=model.matrix(~-1+factor(zyg),data=dl)
out2<-random.cif(cl,data=dl,cause1=1,cause2=1,theta.des=mm,
                 weights=dl$weights,censoring.weights=rep(1,nrow(dl)))
summary(out2)
```

---

km *Kaplan-Meier with robust standard errors*

Description

Kaplan-Meier with robust standard errors Robust variance is default variance with the summary.

Usage

```
km(formula, data = data, conf.type = "log", conf.int = 0.95,
  robust = TRUE, ...)
```

Arguments

formula	formula with 'Surv' outcome (see coxph)
data	data frame
conf.type	transformation
conf.int	level of confidence intervals
robust	for robust standard errors based on martingales
...	Additional arguments to lower level funtions

Author(s)

Thomas Scheike

Examples

```
data(TRACE)
TRACE$cluster <- sample(1:100,1878,replace=TRUE)
out1 <- km(Surv(time,status==9)~strata(vf,chf),data=TRACE)
out2 <- km(Surv(time,status==9)~strata(vf,chf)+cluster(cluster),data=TRACE)

par(mfrow=c(1,2))
bplot(out1,se=TRUE)
bplot(out2,se=TRUE)
```

`lifecourse` *Life-course plot*

Description

Life-course plot for event life data with recurrent events

Usage

```
lifecourse(formula, data, id = "id", group = NULL, type = "l",
  lty = 1, col = 1:10, alpha = 0.3, lwd = 1,
  recurrent.col = NULL, recurrent.lty = NULL, legend = NULL,
  pchlegend = NULL, by = NULL, status.legend = NULL,
  place.sl = "bottomright", xlab = "Time", ylab = "", add = FALSE,
  ...)
```

Arguments

`formula`	Formula (Event(start,slut,status) ~ ...)
`data`	data.frame
`id`	Id variable
`group`	group variable
`type`	Type (line 'l', stair 's', ...)
`lty`	Line type
`col`	Colour
`alpha`	transparency (0-1)
`lwd`	Line width
`recurrent.col`	col of recurrence type
`recurrent.lty`	lty's of of recurrence type
`legend`	position of optional id legend
`pchlegend`	point type legends
`by`	make separate plot for each level in 'by' (formula, name of column, or vector)
`status.legend`	Status legend
`place.sl`	Placement of status legend
`xlab`	Label of X-axis
`ylab`	Label of Y-axis
`add`	Add to existing device
`...`	Additional arguments to lower level arguments

Author(s)

Thomas Scheike, Klaus K. Holst

Examples

```
data = data.frame(id=c(1,1,1,2,2),start=c(0,1,2,3,4),slut=c(1,2,4,4,7),
                  type=c(1,2,3,2,3),status=c(0,1,2,1,2),group=c(1,1,1,2,2))
ll = lifecourse(Event(start,slut,status)~id,data,id="id")
ll = lifecourse(Event(start,slut,status)~id,data,id="id",recurrent.col="type")

ll = lifecourse(Event(start,slut,status)~id,data,id="id",group=~group,col=1:2)
op <- par(mfrow=c(1,2))
ll = lifecourse(Event(start,slut,status)~id,data,id="id",by=~group)
par(op)
legends=c("censored","pregnant","married")
ll = lifecourse(Event(start,slut,status)~id,data,id="id",group=~group,col=1:2,status.legend=legends)
```

`lifetable.matrix` *Life table*

Description

Create simple life table

Usage

```
## S3 method for class 'matrix'
lifetable(x, strata = list(), breaks = c(),
   weights=NULL, confint = FALSE, ...)

 ## S3 method for class 'formula'
lifetable(x, data=parent.frame(), breaks = c(),
   weights=NULL, confint = FALSE, ...)
```

Arguments

`x`	time formula (Surv) or matrix/data.frame with columns time,status or entry,exit,status
`strata`	strata
`breaks`	time intervals
`weights`	weights variable
`confint`	if TRUE 95% confidence limits are calculated
`...`	additional arguments to lower level functions
`data`	data.frame

Author(s)

Klaus K. Holst

Examples

```
library(timereg)
data(TRACE)

d <- with(TRACE,lifetable(Surv(time,status==9)~sex+vf,breaks=c(0,0.2,0.5,8.5)))
summary(glm(events ~ offset(log(atrisk))+factor(int.end)*vf + sex*vf,
            data=d,poisson))
```

LinSpline *Simple linear spline*

Description

Simple linear spline

Usage

```
LinSpline(x, knots, num = TRUE, name = "Spline")
```

Arguments

x	variable to make into spline
knots	cut points
num	to give names x1 x2 and so forth
name	name of spline expansion name.1 name.2 and so forth

Author(s)

Thomas Scheike

---

logitSurv *Proportional odds survival model*

Description

Semiparametric Proportional odds model, that has the advantage that

$$logit(S(t|x)) = \log(\Lambda(t)) + x\beta$$

so covariate effects give OR of survival.

Usage

```
logitSurv(formula, data, offset = NULL, weights = NULL, ...)
```

Arguments

formula	formula with 'Surv' outcome (see coxph)
data	data frame
offset	offsets for exp(x beta) terms
weights	weights for score equations
...	Additional arguments to lower level funtions

Details

This is equivalent to using a hazards model

$$Z\lambda(t)\exp(x\beta)$$

where Z is gamma distributed with mean and variance 1.

Author(s)

Thomas Scheike

Examples

```
data(TRACE)
dcut(TRACE) <- ~.
out1 <- logitSurv(Surv(time,status==9)~vf+chf+strata(wmicat.4),data=TRACE)
summary(out1)
gof(out1)
bplot(out1)
```

---

mena *Menarche data set*

Description

Menarche data set

Source

Simulated data

mets.options *Set global options for* mets

Description

Extract and set global parameters of mets.

Usage

```
mets.options(...)
```

Arguments

...	Arguments

Details

- regex: If TRUE character vectors will be interpreted as regular expressions (dby, dcut, ...)
- silent: Set to FALSE to disable various output messages

Value

list of parameters

Examples

```
## Not run:
mets.options(regex=TRUE)

## End(Not run)
```

migr *Migraine data*

Description

Migraine data

mlogit *Multinomial regression based on phreg regression*

Description

Fits multinomial regression model

$$P_i = \frac{\exp(X_i^\beta)}{\sum_{j=1}^{K} \exp(X_j^\beta)}$$

for

$$i = 1, .., K$$

where

$$\beta_1 = 0$$

, such that

$$\sum_j P_j = 1$$

using phreg function. Thefore the ratio

$$\frac{P_i}{P_1} = \exp(X_i^\beta)$$

Usage

```
mlogit(formula, data, offset = NULL, weights = NULL, ...)
```

Arguments

formula	formula with outcome (see coxph)
data	data frame
offset	offsets for partial likelihood
weights	for score equations
...	Additional arguments to lower level funtions

Details

Coefficients give log-Relative-Risk relative to baseline group (first level of factor, so that it can reset by relevel command). Standard errors computed based on sandwhich form

$$DU^-1 \sum U_i^2 DU^-1$$

.

Can also get influence functions (possibly robust) via iid() function, response should be a factor.

Author(s)

Thomas Scheike

Examples

```
data(bmt)
dfactor(bmt) <- cause1f~cause
drelevel(bmt,ref=3) <- cause3f~cause
dlevels(bmt)

mreg <- mlogit(cause1f~tcell+platelet,bmt)
summary(mreg)

mreg3 <- mlogit(cause3f~tcell+platelet,bmt)
summary(mreg3)

## inverse information standard errors
estimate(coef=mreg3$coef,vcov=mreg3$II)
```

---

multcif *Multivariate Cumulative Incidence Function example data set*

Description

Multivariate Cumulative Incidence Function example data set

Source

Simulated data

---

np *np data set*

Description

np data set

Source

Simulated data

npc *For internal use*

Description

For internal use

Author(s)

Klaus K. Holst

phreg *Fast Cox PH regression*

Description

Fast Cox PH regression Robust variance is default variance with the summary.

Usage

```
phreg(formula, data, offset = NULL, weights = NULL, ...)
```

Arguments

formula	formula with 'Surv' outcome (see coxph)
data	data frame
offset	offsets for cox model
weights	weights for Cox score equations
...	Additional arguments to lower level funtions

Details

influence functions (iid) will follow numerical order of given cluster variable so ordering after $id will give iid in order of data-set.

Author(s)

Klaus K. Holst, Thomas Scheike

Examples

```
data(TRACE)
dcut(TRACE) <- ~.
out1 <- phreg(Surv(time,status==9)~vf+chf+strata(wmicat.4),data=TRACE)
## tracesim <- timereg::sim.cox(out1,1000)
## sout1 <- phreg(Surv(time,status==1)~vf+chf+strata(wmicat.4),data=tracesim)
## robust standard errors default
summary(out1)

par(mfrow=c(1,2))
bplot(out1)
## bplot(sout1,se=TRUE)

## computing robust variance for baseline
rob1 <- robust.phreg(out1)
bplot(rob1,se=TRUE,robust=TRUE)

## making iid decomposition of regression parameters
betaiiid <- iid(out1)
```

`phregR` *Fast Cox PH regression and calculations done in R to make play and adjustments easy*

Description

Fast Cox PH regression with R implementation to play and adjust in R function: FastCoxPLstrataR

Usage

```
phregR(formula, data, offset = NULL, weights = NULL, ...)
```

Arguments

`formula`	formula with 'Surv' outcome (see `coxph`)
`data`	data frame
`offset`	offsets for cox model
`weights`	weights for Cox score equations
`...`	Additional arguments to lower level funtions

Details

Robust variance is default variance with the summary.

influence functions (iid) will follow numerical order of given cluster variable so ordering after $id will give iid in order of data-set.

Author(s)

Klaus K. Holst, Thomas Scheike

`plack.cif` *plack Computes concordance for or.cif based model, that is Plackett random effects model*

Description

.. content for description (no empty lines) ..

Usage

```
plack.cif(cif1, cif2, object)
```

Arguments

`cif1`	Cumulative incidence of first argument.
`cif2`	Cumulative incidence of second argument.
`object`	or.cif object with dependence parameters.

Author(s)

Thomas Scheike

`pmvn` *Multivariate normal distribution function*

Description

Multivariate normal distribution function

Usage

```
pmvn(lower, upper, mu, sigma, cor = FALSE)
```

Arguments

`lower`	lower limits
`upper`	upper limits
`mu`	mean vector
`sigma`	variance matrix or vector of correlation coefficients
`cor`	if TRUE sigma is treated as standardized (correlation matrix)

Examples

```
lower <- rbind(c(0,-Inf),c(-Inf,0))
upper <- rbind(c(Inf,0),c(0,Inf))
mu <- rbind(c(1,1),c(-1,1))
sigma <- diag(2)+1
pmvn(lower=lower,upper=upper,mu=mu,sigma=sigma)
```

`predict.phreg` *Predictions from proportional hazards model*

Description

Predictions from proportional hazards model

Usage

```
## S3 method for class 'phreg'
predict(object, newdata, times = NULL,
  individual.time = FALSE, tminus = FALSE, se = TRUE,
  robust = FALSE, conf.type = "log", conf.int = 0.95, km = FALSE,
  ...)
```

Arguments

`object`	phreg object
`newdata`	data.frame
`times`	Time where to predict variable, default is all time-points from the object sorted
`individual.time`	when TRUE then newdata and times have same length and makes only predictions for these individual times.
`tminus`	to make predictions in T- that is just before, useful for IPCW techniques
`se`	with standard errors and upper and lower confidence intervals.
`robust`	to get robust se's.
`conf.type`	transformation for suvival estimates, default is log
`conf.int`	significance level
`km`	to use Kaplan-Meier for baseline $$S_{s0}(t) = (1 - dA_{s0}(t))$$ where s is strata.
`...`	Additional arguments to plot functions

print.casewise *prints Concordance test*

Description

prints Concordance test

Usage

```
## S3 method for class 'casewise'
print(x, digits = 3, ...)
```

Arguments

x	output from casewise.test
digits	number of digits
...	Additional arguments to lower level functions

Author(s)

Thomas Scheike

prob.exceed.recurrent *Estimation of probability of more that k events for recurrent events process*

Description

Estimation of probability of more that k events for recurrent events process where there is terminal event, based on this also estimate of variance of recurrent events. The estimator is based on cumulative incidence of exceeding "k" events. In contrast the probability of exceeding k events can also be computed as a counting process integral, and this is implemented in prob.exceedRecurrent

Usage

```
prob.exceed.recurrent(data, type, status = "status", death = "death",
  start = "start", stop = "stop", id = "id", times = NULL,
  exceed = NULL, cifmets = FALSE, strata = NULL, all.cifs = FALSE,
  ...)
```

Arguments

data	data-frame
type	type of evnent (code) related to status
status	name of status
death	name of death indicator
start	start stop call of Hist() of prodlim
stop	start stop call of Hist() of prodlim
id	id
times	time at which to get probabilites P(N1(t) >= n)
exceed	n's for which which to compute probabilites P(N1(t) >= n)
cifmets	if true uses cif of mets package rather than prodlim
strata	to stratify according to variable, only for cifmets=TRUE, when strata is given then only consider the output in the all.cifs
all.cifs	if true then returns list of all fitted objects in cif.exceed
...	Additional arguments to lower level funtions

Author(s)

Thomas Scheike

References

Scheike, Eriksson, Tribler (2019) The mean, variance and correlation for bivariate recurrent events with a terminal event, JRSS-C

Examples

```
########################################
## getting some rates to mimick
########################################

data(base1cumhaz)
data(base4cumhaz)
data(drcumhaz)
dr <- drcumhaz
base1 <- base1cumhaz
base4 <- base4cumhaz

cor.mat <- corM <- rbind(c(1.0, 0.6, 0.9), c(0.6, 1.0, 0.5), c(0.9, 0.5, 1.0))
rr <- simRecurrent(1000,base1,cumhaz2=base4,death.cumhaz=dr)
rr <-  count.history(rr)
dtable(rr,~death+status)

oo <- prob.exceedRecurrent(rr,1)
bplot(oo)
```

```
par(mfrow=c(1,2))
with(oo,plot(time,mu,col=2,type="l"))
###
with(oo,plot(time,varN,type="l"))


### Bivariate probability of exceeding
oo <- prob.exceedBiRecurrent(rr,1,2,exceed1=c(1,5,10),exceed2=c(1,2,3))
with(oo, matplot(time,pe1e2,type="s"))
nc <- ncol(oo$pe1e2)
legend("topleft",legend=colnames(oo$pe1e2),lty=1:nc,col=1:nc)



### do not test to avoid dependence on prodlim
### now estimation based on cumualative incidence, but do not test to avoid dependence on prodlim
library(prodlim)
pp <- prob.exceed.recurrent(rr,1,status="status",death="death",start="entry",stop="time",id="id")
with(pp, matplot(times,prob,type="s"))
###
with(pp, matlines(times,se.lower,type="s"))
with(pp, matlines(times,se.upper,type="s"))
```

`prt` *Prostate data set*

Description

Prostate data set

Source

Simulated data

`random.cif` *Random effects model for competing risks data*

Description

Fits a random effects model describing the dependence in the cumulative incidence curves for subjects within a cluster. Given the gamma distributed random effects it is assumed that the cumulative incidence curves are indpendent, and that the marginal cumulative incidence curves are on the form

$$P(T \leq t, cause = 1|x, z) = P_1(t, x, z) = 1 - exp(-x^T A(t) exp(z^T \beta))$$

We allow a regression structure for the random effects variances that may depend on cluster covariates.

Usage

```
random.cif(cif, data, cause = NULL, cif2 = NULL, cause1 = 1,
  cause2 = 1, cens.code = NULL, cens.model = "KM", Nit = 40,
  detail = 0, clusters = NULL, theta = NULL, theta.des = NULL,
  sym = 1, step = 1, same.cens = FALSE, var.link = 0,
  score.method = "fisher.scoring", entry = NULL, trunkp = 1, ...)
```

Arguments

cif	a model object from the comp.risk function with the marginal cumulative incidence of cause2, i.e., the event that is conditioned on, and whose odds the comparision is made with respect to
data	a data.frame with the variables.
cause	specifies the causes related to the death times, the value cens.code is the censoring value.
cif2	specificies model for cause2 if different from cause1.
cause1	cause of first coordinate.
cause2	cause of second coordinate.
cens.code	specificies the code for the censoring if NULL then uses the one from the marginal cif model.
cens.model	specified which model to use for the ICPW, KM is Kaplan-Meier alternatively it may be "cox"
Nit	number of iterations for Newton-Raphson algorithm.
detail	if 0 no details are printed during iterations, if 1 details are given.
clusters	specifies the cluster structure.
theta	specifies starting values for the cross-odds-ratio parameters of the model.
theta.des	specifies a regression design for the cross-odds-ratio parameters.
sym	1 for symmetry 0 otherwise
step	specifies the step size for the Newton-Raphson algorith.m
same.cens	if true then censoring within clusters are assumed to be the same variable, default is independent censoring.
var.link	if var.link=1 then var is on log-scale.
score.method	default uses "nlminb" optimzer, alternatively, use the "fisher-scoring" algorithm.
entry	entry-age in case of delayed entry. Then two causes must be given.
trunkp	gives probability of survival for delayed entry, and related to entry-ages given above.
...	extra arguments.

Value

returns an object of type 'cor'. With the following arguments:

theta	estimate of proportional odds parameters of model.
var.theta	variance for gamma.
hess	the derivative of the used score.
score	scores at final stage.
score	scores at final stage.
theta.iid	matrix of iid decomposition of parametric effects.

Author(s)

Thomas Scheike

References

A Semiparametric Random Effects Model for Multivariate Competing Risks Data, Scheike, Zhang, Sun, Jensen (2010), Biometrika.

Cross odds ratio Modelling of dependence for Multivariate Competing Risks Data, Scheike and Sun (2012), work in progress.

Examples

```
## Reduce Ex.Timings
library("timereg")
d <- simnordic.random(4000,delayed=TRUE,
      cordz=0.5,cormz=2,lam0=0.3,country=TRUE)
times <- seq(50,90,by=10)
add1<-comp.risk(Event(time,cause)~-1+factor(country)+cluster(id),data=d,
times=times,cause=1,max.clust=NULL)

### making group indidcator
mm <- model.matrix(~-1+factor(zyg),d)

out1<-random.cif(add1,data=d,cause1=1,cause2=1,theta=1,same.cens=TRUE)
summary(out1)

out2<-random.cif(add1,data=d,cause1=1,cause2=1,theta=1,
  theta.des=mm,same.cens=TRUE)
summary(out2)

########################################
##### 2 different causes
########################################

add2<-comp.risk(Event(time,cause)~const(country)+cluster(id),data=d,
                 times=times,cause=2,max.clust=NULL)
out3<-random.cif(add1,data=d,cause1=1,cause2=2,cif2=add2,sym=1,same.cens=TRUE)
summary(out3) ## negative dependence
```

```
out4<-random.cif(add1,data=d,cause1=1,cause2=2,cif2=add2,theta.des=mm,sym=1,same.cens=TRUE)
summary(out4) ## negative dependence
```

`recurrentMarginal` *Fast recurrent marginal mean when death is possible*

Description

Fast Marginal means of recurrent events. Using the Lin and Ghosh (2000) standard errors. Fitting two models for death and recurent events these are combined to prducte the estimator

$$\int_0^t S(u|x=0)dR(u|x=0)$$

the mean number of recurrent events, here

$$S(u|x=0)$$

is the probability of survival for the baseline group, and

$$dR(u|x=0)$$

is the hazard rate of an event among survivors for the baseline. Here

$$S(u|x=0)$$

is estimated by

$$exp(-\Lambda_d(u|x=0)$$

with

$$\Lambda_d(u|x=0)$$

being the cumulative baseline for death.

Usage

```
recurrentMarginal(recurrent, death, fixbeta = NULL, km = TRUE, ...)
```

Arguments

`recurrent`	phreg object with recurrent events
`death`	phreg object with deaths
`fixbeta`	to force the estimation of standard errors to think of regression coefficients as known/fixed
`km`	if true then uses Kaplan-Meier for death, otherwise exp(- Nelson-Aalen)
`...`	Additional arguments to lower level funtions

Details

Assumes no ties in the sense that jump times needs to be unique, this is particularly so for the stratified version.

Author(s)

Thomas Scheike

References

Ghosh and Lin (2002) Nonparametric Analysis of Recurrent events and death, Biometrics, 554–562.

Examples

```
data(base1cumhaz)
data(base4cumhaz)
data(drcumhaz)
dr <- drcumhaz
base1 <- base1cumhaz
base4 <- base4cumhaz
rr <- simRecurrent(1000,base1,death.cumhaz=dr)
rr$x <- rnorm(nrow(rr))
rr$strata <- floor((rr$id-0.01)/500)

##  to fit non-parametric models with just a baseline
xr <- phreg(Surv(entry,time,status)~cluster(id),data=rr)
dr <- phreg(Surv(entry,time,death)~cluster(id),data=rr)
par(mfrow=c(1,3))
bplot(dr,se=TRUE)
title(main="death")
bplot(xr,se=TRUE)
### robust standard errors
rxr <-   robust.phreg(xr,fixbeta=1)
bplot(rxr,se=TRUE,robust=TRUE,add=TRUE,col=4)

## marginal mean of expected number of recurrent events
out <- recurrentMarginal(xr,dr)
bplot(out,se=TRUE,ylab="marginal mean",col=2)

#########################################################################
###   with strata     #################################################
#########################################################################
xr <- phreg(Surv(entry,time,status)~strata(strata)+cluster(id),data=rr)
dr <- phreg(Surv(entry,time,death)~strata(strata)+cluster(id),data=rr)
par(mfrow=c(1,3))
bplot(dr,se=TRUE)
title(main="death")
bplot(xr,se=TRUE)
rxr <-   robust.phreg(xr,fixbeta=1)
bplot(rxr,se=TRUE,robust=TRUE,add=TRUE,col=1:2)
```

```
out <- recurrentMarginal(xr,dr)
bplot(out,se=TRUE,ylab="marginal mean",col=1:2)

#########################################################################
###   cox case        ###################################################
#########################################################################
xr <- phreg(Surv(entry,time,status)~x+cluster(id),data=rr)
dr <- phreg(Surv(entry,time,death)~x+cluster(id),data=rr)
par(mfrow=c(1,3))
bplot(dr,se=TRUE)
title(main="death")
bplot(xr,se=TRUE)
rxr <-   robust.phreg(xr)
bplot(rxr,se=TRUE,robust=TRUE,add=TRUE,col=1:2)

out <- recurrentMarginal(xr,dr)
bplot(out,se=TRUE,ylab="marginal mean",col=1:2)

#########################################################################
###   CIF  ##############################################################
#########################################################################
### use of function to compute cumulative incidence (cif) with robust standard errors
 data(bmt)
 bmt$id <- 1:nrow(bmt)
 xr  <- phreg(Surv(time,cause==1)~cluster(id),data=bmt)
 dr  <- phreg(Surv(time,cause!=0)~cluster(id),data=bmt)

 out <- recurrentMarginal(xr,dr,km=TRUE)
 bplot(out,se=TRUE,ylab="cumulative incidence")
```

---

rpch *Piecewise constant hazard distribution*

Description

Piecewise constant hazard distribution

Usage

```
rpch(n, lambda = 1, breaks = c(0, Inf))
```

Arguments

n	sample size
lambda	rate parameters
breaks	time cut-points

simAalenFrailty *Simulate from the Aalen Frailty model*

Description

Simulate observations from Aalen Frailty model with Gamma distributed frailty and constant intensity.

Usage

```
simAalenFrailty(n = 5000, theta = 0.3, K = 2, beta0 = 1.5,
  beta = 1, cens = 1.5, cuts = 0, ...)
```

Arguments

n	Number of observations in each cluster
theta	Dependence paramter (variance of frailty)
K	Number of clusters
beta0	Baseline (intercept)
beta	Effect (log hazard ratio) of covariate
cens	Censoring rate
cuts	time cuts
...	Additional arguments

Author(s)

Klaus K. Holst

simClaytonOakes *Simulate from the Clayton-Oakes frailty model*

Description

Simulate observations from the Clayton-Oakes copula model with piecewise constant marginals.

Usage

```
simClaytonOakes(K, n, eta, beta, stoptime, lam = 1, left = 0,
  pairleft = 0, trunc.prob = 0.5, same = 0)
```

Arguments

K	Number of clusters
n	Number of observations in each cluster
eta	variance
beta	Effect (log hazard ratio) of covariate
stoptime	Stopping time
lam	constant hazard
left	Left truncation
pairleft	pairwise (1) left truncation or individual (0)
trunc.prob	Truncation probability
same	if 1 then left-truncation is same also for univariate truncation

Author(s)

Thomas Scheike and Klaus K. Holst

---

simClaytonOakesWei *Simulate from the Clayton-Oakes frailty model*

Description

Simulate observations from the Clayton-Oakes copula model with Weibull type baseline and Cox marginals.

Usage

```
simClaytonOakesWei(K, n, eta, beta, stoptime, weiscale = 1,
  weishape = 2, left = 0, pairleft = 0)
```

Arguments

K	Number of clusters
n	Number of observations in each cluster
eta	1/variance
beta	Effect (log hazard ratio) of covariate
stoptime	Stopping time
weiscale	weibull scale parameter
weishape	weibull shape parameter
left	Left truncation
pairleft	pairwise (1) left truncation or individual (0)

Author(s)

Klaus K. Holst

`simMultistate` *Simulation of illness-death model*

Description

Simulation of illness-death model

Usage

```
simMultistate(n, cumhaz, cumhaz2, death.cumhaz, death.cumhaz2, rr = NULL,
  rr2 = NULL, rd = NULL, rd2 = NULL, gap.time = FALSE,
  max.recurrent = 100, dependence = 0, var.z = 0.22,
  cor.mat = NULL, cens = NULL, ...)
```

Arguments

`n`	number of id's
`cumhaz`	cumulative hazard of recurrent events
`cumhaz2`	cumulative hazard of recurrent events of type 2
`death.cumhaz`	cumulative hazard of death from state 1
`death.cumhaz2`	cumulative hazard of death from state 2
`rr`	relative risk adjustment for cumhaz
`rr2`	relative risk adjustment for cumhaz2
`rd`	relative risk adjustment for death.cumhaz
`rd2`	relative risk adjustment for death.cumhaz2
`gap.time`	if true simulates gap-times with specified cumulative hazard
`max.recurrent`	limits number recurrent events to 100
`dependence`	0:independence; 1:all share same random effect with variance var.z; 2:random effect exp(normal) with correlation structure from cor.mat; 3:additive gamma distributed random effects, z1= (z11+ z12)/2 such that mean is 1 , z2= (z11^cor.mat(1,2)+ z13)/2, z3= (z12^(cor.mat(2,3)+z13^cor.mat(1,3))/2, with z11 z12 z13 are gamma with mean and variance 1 , first random effect is z1 and for N1 second random effect is z2 and for N2 third random effect is for death
`var.z`	variance of random effects
`cor.mat`	correlation matrix for var.z variance of random effects
`cens`	rate of censoring exponential distribution
`...`	Additional arguments to lower level funtions

Details

simMultistate with same death intensity from states 1 and 2 simMultistateII with different death intensities from states 1 and 2

Must give cumulative hazards on some time-range

Author(s)

Thomas Scheike

Examples

```
########################################
## getting some rates to mimick
########################################
data(base1cumhaz)
data(base4cumhaz)
data(drcumhaz)
dr <- drcumhaz
dr2 <- drcumhaz
dr2[,2] <- 1.5*drcumhaz[,2]
base1 <- base1cumhaz
base4 <- base4cumhaz
cens <- rbind(c(0,0),c(2000,0.5),c(5110,3))

iddata <- simMultistate(100,base1,base1,dr,dr2,cens=cens)
dlist(iddata,.~id|id<3,n=0)

### estimating rates from simulated data
c0 <- phreg(Surv(start,stop,status==0)~+1,iddata)
c3 <- phreg(Surv(start,stop,status==3)~+strata(from),iddata)
c1 <- phreg(Surv(start,stop,status==1)~+1,subset(iddata,from==2))
c2 <- phreg(Surv(start,stop,status==2)~+1,subset(iddata,from==1))
###
par(mfrow=c(2,3))
bplot(c0)
lines(cens,col=2)
bplot(c3,main="rates 1-> 3 , 2->3")
lines(dr,col=1,lwd=2)
lines(dr2,col=2,lwd=2)
###
bplot(c1,main="rate 1->2")
lines(base1,lwd=2)
###
bplot(c2,main="rate 2->1")
lines(base1,lwd=2)
```

---

simRecurrent *Simulation of recurrent events data based on cumulative hazards*

Description

Simulation of recurrent events data based on cumulative hazards

Usage

```
simRecurrent(n, cumhaz, death.cumhaz = NULL, cumhaz2 = NULL,
  gap.time = FALSE, max.recurrent = 100, dhaz = NULL, haz2 = NULL,
  dependence = 0, var.z = 2, cor.mat = NULL, ...)
```

Arguments

n	number of id's
cumhaz	cumulative hazard of recurrent events
death.cumhaz	cumulative hazard of death
cumhaz2	cumulative hazard of recurrent events of type 2
gap.time	if true simulates gap-times with specified cumulative hazard
max.recurrent	limits number recurrent events to 100
dhaz	rate for death hazard if it is extended to time-range of first event
haz2	rate of second cause if it is extended to time-range of first event
dependence	=0 independence, =1 all share same random effect with variance var.z =2 random effect exp(normal) with correlation structure from cor.mat, first random effect is z1 and shared for a possible second cause, second random effect is for death
var.z	variance of random effects
cor.mat	correlation matrix for var.z variance of random effects
...	Additional arguments to lower level funtions

Details

Must give hazard of death and recurrent events. Possible with two event types and their dependence can be specified but the two recurrent events need to have the same random effect, simRecurrentII more flexible !

Author(s)

Thomas Scheike

Examples

```
######################################
## getting some rates to mimick
######################################

data(base1cumhaz)
data(base4cumhaz)
data(drcumhaz)
dr <- drcumhaz
base1 <- base1cumhaz
base4 <- base4cumhaz

 ######################################################################
```

```
### simulating simple model that mimicks data
######################################################################
rr <- simRecurrent(5,base1,death.cumhaz=dr)
dlist(rr,.~id,n=0)

rr <- simRecurrent(1000,base1,death.cumhaz=dr)
par(mfrow=c(1,3))
showfitsim(causes=1,rr,dr,base1,base1)

######################################################################
### simulating simple model that mimicks data
### now with two event types and second type has same rate as death rate
######################################################################

rr <- simRecurrent(1000,base1,death.cumhaz=dr,cumhaz2=base4)
dtable(rr,~death+status)
par(mfrow=c(2,2))
showfitsim(causes=2,rr,dr,base1,base4)

######################################################################
### simulating simple model
### random effect for all causes (Z shared for death and recurrent)
######################################################################

rr <- simRecurrent(1000,base1,death.cumhaz=dr,dependence=1,var.gamma=0.4)
### marginals do fit after input after integrating out
par(mfrow=c(2,2))
showfitsim(causes=1,rr,dr,base1,base1)
```

`simRecurrentII` *Simulation of recurrent events data based on cumulative hazards II*

Description

Simulation of recurrent events data based on cumulative hazards

Usage

```
simRecurrentII(n, cumhaz, cumhaz2, death.cumhaz = NULL,
  gap.time = FALSE, max.recurrent = 100, dhaz = NULL, haz2 = NULL,
  dependence = 0, var.z = 0.22, cor.mat = NULL, cens = NULL, ...)
```

Arguments

`n`	number of id's
`cumhaz`	cumulative hazard of recurrent events
`cumhaz2`	cumulative hazard of recurrent events of type 2
`death.cumhaz`	cumulative hazard of death

gap.time	if true simulates gap-times with specified cumulative hazard
max.recurrent	limits number recurrent events to 100
dhaz	rate for death hazard if it is extended to time-range of first event
haz2	rate of second cause if it is extended to time-range of first event
dependence	0:independence; 1:all share same random effect with variance var.z; 2:random effect exp(normal) with correlation structure from cor.mat; 3:additive gamma distributed random effects, z1= (z11+ z12)/2 such that mean is 1 , z2= (z11^cor.mat(1,2)+ z13)/2, z3= (z12^(cor.mat(2,3)+z13^cor.mat(1,3))/2, with z11 z12 z13 are gamma with mean and variance 1 , first random effect is z1 and for N1 second random effect is z2 and for N2 third random effect is for death
var.z	variance of random effects
cor.mat	correlation matrix for var.z variance of random effects
cens	rate of censoring exponential distribution
...	Additional arguments to lower level funtions

Details

Must give hazard of death and two recurrent events. Possible with two event types and their dependence can be specified but the two recurrent events need to share random effect. Based on drawing the from cumhaz and cumhaz2 and taking the first event rather the cumulative and then distributing it out. Key advantage of this is that there is more flexibility wrt random effects

Author(s)

Thomas Scheike

Examples

```
########################################
## getting some rates to mimick
########################################

data(base1cumhaz)
data(base4cumhaz)
data(drcumhaz)
dr <- drcumhaz
base1 <- base1cumhaz
base4 <- base4cumhaz

 cor.mat <- corM <- rbind(c(1.0, 0.6, 0.9), c(0.6, 1.0, 0.5), c(0.9, 0.5, 1.0))

######################################################################
### simulating simple model that mimicks data
### now with two event types and second type has same rate as death rate
######################################################################

rr <- simRecurrentII(1000,base1,base4,death.cumhaz=dr)
dtable(rr,~death+status)
```

```
par(mfrow=c(2,2))
showfitsim(causes=2,rr,dr,base1,base4)
```

`simRecurrentTS` *Simulation of recurrent events data based on cumulative hazards: Two-stage model*

Description

Simulation of recurrent events data based on cumulative hazards

Usage

```
simRecurrentTS(n, cumhaz, cumhaz2, death.cumhaz = NULL, nu = rep(1, 3),
  share1 = 0.3, vargamD = 2, vargam12 = 0.5, gap.time = FALSE,
  max.recurrent = 100, cens = NULL, ...)
```

Arguments

`n`	number of id's
`cumhaz`	cumulative hazard of recurrent events
`cumhaz2`	cumulative hazard of recurrent events of type 2
`death.cumhaz`	cumulative hazard of death
`nu`	powers of random effects where nu > -1/shape
`share1`	how random effect for death splits into two parts
`vargamD`	variance of random effect for death
`vargam12`	shared random effect for N1 and N2
`gap.time`	if true simulates gap-times with specified cumulative hazard
`max.recurrent`	limits number recurrent events to 100
`cens`	rate of censoring exponential distribution
`...`	Additional arguments to lower level funtions

Details

Model is constructed such that marginals are on specified form by linear approximations of cumulative hazards that are on a specific form to make them equivalent to marginals after integrating out.

Must give hazard of death and two recurrent events. Possible with two event types and their dependence can be specified but the two recurrent events need to share random effect.

Random effect to death Z.death=(Zd1+Zd2), Z1=(Zd1^nu1) Z12, Z2=(Zd2^nu2) Z12^nu3

$$Z.death = Zd1 + Zd2$$

gamma distributions

$$Zdj$$

gamma distribution with mean parameters (sharej), vargamD, share2=1-share1

$$Z12$$

gamma distribution with mean 1 and variance vargam12

Author(s)

Thomas Scheike

Examples

```
########################################
## getting some rates to mimick
########################################

data(base1cumhaz)
data(base4cumhaz)
data(drcumhaz)
dr <- drcumhaz
base1 <- base1cumhaz
base4 <- base4cumhaz

rr <- simRecurrentTS(1000,base1,base4,death.cumhaz=dr)
dtable(rr,~death+status)
showfitsim(causes=2,rr,dr,base1,base4)
```

`summary.cor` *Summary for dependence models for competing risks*

Description

Computes concordance and casewise concordance for dependence models for competing risks models of the type cor.cif, rr.cif or or.cif for the given cumulative incidences and the different dependence measures in the object.

Usage

```
## S3 method for class 'cor'
summary(object, marg.cif = NULL, marg.cif2 = NULL,
  digits = 3, ...)
```

Arguments

object	object from cor.cif rr.cif or or.cif for dependence between competing risks data for two causes.
marg.cif	a number that gives the cumulative incidence in one time point for which concordance and casewise concordance are computed.
marg.cif2	the cumulative incidence for cause 2 for concordance and casewise concordance are computed. Default is that it is the same as marg.cif.
digits	digits in output.
...	Additional arguments.

Value

prints summary for dependence model.

casewise	gives casewise concordance that is, probability of cause 2 (related to cif2) given that cause 1 (related to cif1) has occured.
concordance	gives concordance that is, probability of cause 2 (related to cif2) and cause 1 (related to cif1).
cif1	cumulative incidence for cause1.
cif2	cumulative incidence for cause1.

Author(s)

Thomas Scheike

References

Cross odds ratio Modelling of dependence for Multivariate Competing Risks Data, Scheike and Sun (2012), Biostatistics.

A Semiparametric Random Effects Model for Multivariate Competing Risks Data, Scheike, Zhang, Sun, Jensen (2010), Biometrika.

Examples

```
library("timereg")
data("multcif",package="mets") # simulated data
multcif$cause[multcif$cause==0] <- 2

times=seq(0.1,3,by=0.1) # to speed up computations use only these time-points
add<-comp.risk(Event(time,cause)~+1+cluster(id),data=multcif,
               n.sim=0,times=times,cause=1)
###
out1<-cor.cif(add,data=multcif,cause1=1,cause2=1,theta=log(2+1))
summary(out1)

pad <- predict(add,X=1,se=0,uniform=0)
summary(out1,marg.cif=pad)
```

`survival.iterative` *Survival model for multivariate survival data*

Description

Fits additive gamma frailty model with additive hazard condtional on the random effects

$$\lambda_{ij} = (V_{ij}^T Z)(X_{ij}^T \alpha(t))$$

The baseline $\alpha(t)$ is profiled out using marginal modelling adjusted for the random effects structure as in Eriksson and Scheike (2015). One advantage of the standard frailty model is that one can deal with competing risks for this model.

For all models the standard errors do not reflect this uncertainty of the baseline estimates, and might therefore be a bit to small. To remedy this one can do bootstrapping or use survival.twostage.fullse function when possible.

If clusters contain more than two times, we use a composite likelihood based on the pairwise bivariate models. Can also fit a additive gamma random effects model described in detail below.

We allow a regression structure for the indenpendent gamma distributed random effects and their variances that may depend on cluster covariates. So

$$\theta = z_j^T \alpha$$

where z is specified by theta.des The reported standard errors are based on the estimated information from the likelihood assuming that the marginals are known.

Can also fit a structured additive gamma random effects model, such as the ACE, ADE model for survival data.

Now random.design specificies the random effects for each subject within a cluster. This is a matrix of 1's and 0's with dimension n x d. With d random effects. For a cluster with two subjects, we let the random.design rows be v_1 and v_2. Such that the random effects for subject 1 is

$$v_1^T (Z_1, ..., Z_d)$$

, for d random effects. Each random effect has an associated parameter $(\lambda_1, ..., \lambda_d)$. By construction subjects 1's random effect are Gamma distributed with mean $\lambda_j / v_1^T \lambda$ and variance $\lambda_j / (v_1^T \lambda)^2$. Note that the random effect $v_1^T (Z_1, ..., Z_d)$ has mean 1 and variance $1/(v_1^T \lambda)$. It is here asssumed that $lamtot = v_1^T \lambda$ is fixed over all clusters as it would be for the ACE model below. The lamtot parameter may be specified separately for some sets of the parameter is the additive.gamma.sum (ags) matrix is specified and then lamtot for the j'th random effect is $ags_j^T \lambda$.

Based on these parameters the relative contribution (the heritability, h) is equivalent to the expected values of the random effects $\lambda_j / v_1^T \lambda$

The DEFAULT parametrization uses the variances of the random effecs

$$\theta_j = \lambda_j / (v_1^T \lambda)^2$$

For alternative parametrizations one can specify how the parameters relate to λ_j with the function

Given the random effects the survival distributions with a cluster are independent and on the form

$$P(T > t|x, z) = exp(-ZA(t)\exp(Z^t beta))$$

The parameters $(\lambda_1, ..., \lambda_d)$ are related to the parameters of the model by a regression construction $pard$ (d x k), that links the d λ parameters with the (k) underlying θ parameters

$$\lambda = theta.des\theta$$

here using theta.des to specify these low-dimension association. Default is a diagonal matrix.

The case.control option that can be used with the pair specification of the pairwise parts of the estimating equations. Here it is assumed that the second subject of each pair is the proband.

Usage

```
survival.iterative(margsurv, data = sys.parent(),
  score.method = "fisher.scoring", Nit = 60, detail = 0,
  clusters = NULL, silent = 1, weights = NULL, control = list(),
  theta = NULL, theta.des = NULL, var.link = 1, iid = 1,
  step = 0.5, model = "clayton.oakes", marginal.trunc = NULL,
  marginal.survival = NULL, marginal.status = NULL, strata = NULL,
  se.clusters = NULL, max.clust = NULL, numDeriv = 0,
  random.design = NULL, pairs = NULL, pairs.rvs = NULL,
  numDeriv.method = "simple", additive.gamma.sum = NULL, var.par = 1,
  cr.models = NULL, case.control = 0, ascertained = 0, shut.up = 0)
```

Arguments

margsurv	Marginal model
data	data frame
score.method	Scoring method "fisher.scoring", "nlminb", "optimize", "nlm"
Nit	Number of iterations
detail	Detail
clusters	Cluster variable
silent	Debug information
weights	Weights
control	Optimization arguments
theta	Starting values for variance components
theta.des	design for dependence parameters, when pairs are given this is could be a (pairs) x (numer of parameters) x (max number random effects) matrix
var.link	Link function for variance
iid	Calculate i.i.d. decomposition
step	Step size
model	model
marginal.trunc	marginal left truncation probabilities

`marginal.survival`
: optional vector of marginal survival probabilities

`marginal.status`
: related to marginal survival probabilities

`strata`
: strata for fitting, see example

`se.clusters`
: for clusters for se calculation with iid

`max.clust`
: max se.clusters for se calculation with iid

`numDeriv`
: to get numDeriv version of second derivative, otherwise uses sum of squared score

`random.design`
: random effect design for additive gamma model, when pairs are given this is a (pairs) x (2) x (max number random effects) matrix, see pairs.rvs below

`pairs`
: matrix with rows of indeces (two-columns) for the pairs considered in the pairwise composite score, useful for case-control sampling when marginal is known.

`pairs.rvs`
: for additive gamma model and random.design and theta.des are given as arrays, this specifice number of random effects for each pair.

`numDeriv.method`
: uses simple to speed up things and second derivative not so important.

`additive.gamma.sum`
: for two.stage=0, this is specification of the lamtot in the models via a matrix that is multiplied onto the parameters theta (dimensions=(number random effects x number of theta parameters), when null then sums all parameters.

`var.par`
: is 1 for the default parametrization with the variances of the random effects, var.par=0 specifies that the λ_j's are used as parameters.

`cr.models`
: competing risks models for two.stage=0, should be given as a list with models for each cause

`case.control`
: assumes case control structure for "pairs" with second column being the probands, when this options is used the twostage model is profiled out via the paired estimating equations for the survival model.

`ascertained`
: if the pair are sampled only when there is an event. This is in contrast to case.control sampling where a proband is given. This can be combined with control probands. Pair-call of twostage is needed and second column of pairs are the first jump time with an event for ascertained pairs, or time of control proband.

`shut.up`
: to make the program more silent in the context of iterative procedures for case-control and ascertained sampling

Author(s)

Thomas Scheike

`survival.twostage` *Twostage survival model for multivariate survival data*

Description

Fits Clayton-Oakes or bivariate Plackett models for bivariate survival data using marginals that are on Cox form. The dependence can be modelled via

1. Regression design on dependence parameter.
2. Random effects, additive gamma model.

If clusters contain more than two subjects, we use a composite likelihood based on the pairwise bivariate models, for MLE see twostageMLE.

The two-stage model is constructed such that given the gamma distributed random effects it is assumed that the survival functions are indpendent, and that the marginal survival functions are on Cox form (or additive form)

$$P(T > t|x) = S(t|x) = exp(-exp(x^T\beta)A_0(t))$$

One possibility is to model the variance within clusters via a regression design, and then one can specify a regression structure for the indenpendent gamma distributed random effect for each cluster, such that the variance is given by

$$\theta = z_j^T\alpha$$

where z is specified by theta.des The reported standard errors are based on the estimated information from the likelihood assuming that the marginals are known.

Can also fit a structured additive gamma random effects model, such as the ACE, ADE model for survival data. In this case the random.design specificies the random effects for each subject within a cluster. This is a matrix of 1's and 0's with dimension n x d. With d random effects. For a cluster with two subjects, we let the random.design rows be v_1 and v_2. Such that the random effects for subject 1 is

$$v_1^T(Z_1, ..., Z_d)$$

, for d random effects. Each random effect has an associated parameter $(\lambda_1, ..., \lambda_d)$. By construction subjects 1's random effect are Gamma distributed with mean $\lambda_j/v_1^T\lambda$ and variance $\lambda_j/(v_1^T\lambda)^2$. Note that the random effect $v_1^T(Z_1, ..., Z_d)$ has mean 1 and variance $1/(v_1^T\lambda)$. It is here asssumed that $lamtot = v_1^T\lambda$ is fixed within clusters as it would be for the ACE model below.

Based on these parameters the relative contribution (the heritability, h) is equivalent to the expected values of the random effects: $\lambda_j/v_1^T\lambda$

The DEFAULT parametrization (var.par=1) uses the variances of the random effecs

$$\theta_j = \lambda_j/(v_1^T\lambda)^2$$

For alternative parametrizations one can specify how the parameters relate to λ_j with the argument var.par=0.

For both types of models the basic model assumptions are that given the random effects of the clusters the survival distributions within a cluster are independent and ' on the form

$$P(T > t|x, z) = exp(-Z \cdot Laplace^{-1}(lamtot, lamtot, S(t|x)))$$

with the inverse laplace of the gamma distribution with mean 1 and variance 1/lamtot.

The parameters $(\lambda_1, ..., \lambda_d)$ are related to the parameters of the model by a regression construction $pard$ (d x k), that links the d λ parameters with the (k) underlying θ parameters

$$\lambda = theta.des\theta$$

here using theta.des to specify these low-dimension association. Default is a diagonal matrix. This can be used to make structural assumptions about the variances of the random-effects as is needed for the ACE model for example.

The case.control option that can be used with the pair specification of the pairwise parts of the estimating equations. Here it is assumed that the second subject of each pair is the proband.

Usage

```
survival.twostage(margsurv, data = sys.parent(),
  score.method = "fisher.scoring", Nit = 60, detail = 0,
  clusters = NULL, silent = 1, weights = NULL, control = list(),
  theta = NULL, theta.des = NULL, var.link = 1, iid = 1,
  step = 0.5, model = "clayton.oakes", marginal.trunc = NULL,
  marginal.survival = NULL, marginal.status = NULL, strata = NULL,
  se.clusters = NULL, numDeriv = 0, random.design = NULL,
  pairs = NULL, pairs.rvs = NULL, numDeriv.method = "simple",
  additive.gamma.sum = NULL, var.par = 1, cr.models = NULL,
  case.control = 0, ascertained = 0, shut.up = 0)
```

Arguments

margsurv	Marginal model
data	data frame
score.method	Scoring method "fisher.scoring", "nlminb", "optimize", "nlm"
Nit	Number of iterations
detail	Detail
clusters	Cluster variable
silent	Debug information
weights	Weights
control	Optimization arguments
theta	Starting values for variance components
theta.des	design for dependence parameters, when pairs are given this is could be a (pairs) x (numer of parameters) x (max number random effects) matrix
var.link	Link function for variance

`iid`	Calculate i.i.d. decomposition
`step`	Step size
`model`	model
`marginal.trunc`	marginal left truncation probabilities
`marginal.survival`	optional vector of marginal survival probabilities
`marginal.status`	related to marginal survival probabilities
`strata`	strata for fitting, see example
`se.clusters`	for clusters for se calculation with iid
`numDeriv`	to get numDeriv version of second derivative, otherwise uses sum of squared score
`random.design`	random effect design for additive gamma model, when pairs are given this is a (pairs) x (2) x (max number random effects) matrix, see pairs.rvs below
`pairs`	matrix with rows of indeces (two-columns) for the pairs considered in the pairwise composite score, useful for case-control sampling when marginal is known.
`pairs.rvs`	for additive gamma model and random.design and theta.des are given as arrays, this specifice number of random effects for each pair.
`numDeriv.method`	uses simple to speed up things and second derivative not so important.
`additive.gamma.sum`	for two.stage=0, this is specification of the lamtot in the models via a matrix that is multiplied onto the parameters theta (dimensions=(number random effects x number of theta parameters), when null then sums all parameters.
`var.par`	is 1 for the default parametrization with the variances of the random effects, var.par=0 specifies that the λ_j's are used as parameters.
`cr.models`	competing risks models for two.stage=0, should be given as a list with models for each cause
`case.control`	assumes case control structure for "pairs" with second column being the probands, when this options is used the twostage model is profiled out via the paired estimating equations for the survival model.
`ascertained`	if the pair are sampled only when there is an event. This is in contrast to case.control sampling where a proband is given. This can be combined with control probands. Pair-call of twostage is needed and second column of pairs are the first jump time with an event for ascertained pairs, or time of control proband.
`shut.up`	to make the program more silent in the context of iterative procedures for case-control and ascertained sampling

Author(s)

Thomas Scheike

References

Twostage estimation of additive gamma frailty models for survival data. Scheike (2019), work in progress

Shih and Louis (1995) Inference on the association parameter in copula models for bivariate survival data, Biometrics, (1995).

Glidden (2000), A Two-Stage estimator of the dependence parameter for the Clayton Oakes model, LIDA, (2000).

Measuring early or late dependence for bivariate twin data Scheike, Holst, Hjelmborg (2015), LIDA

Estimating heritability for cause specific mortality based on twins studies Scheike, Holst, Hjelmborg (2014), LIDA

Additive Gamma frailty models for competing risks data, Biometrics (2015) Eriksson and Scheike (2015),

Examples

```
data(diabetes)

# Marginal Cox model  with treat as covariate
margph <- phreg(Surv(time,status)~treat+cluster(id),data=diabetes)
### Clayton-Oakes, MLE
fitco1<-twostageMLE(margph,data=diabetes,theta=1.0)
summary(fitco1)

### Plackett model
mph <- phreg(Surv(time,status)~treat+cluster(id),data=diabetes)
fitp <- survival.twostage(mph,data=diabetes,theta=3.0,Nit=40,
               clusters=diabetes$id,var.link=1,model="plackett")
summary(fitp)

### Clayton-Oakes
fitco2 <- survival.twostage(mph,data=diabetes,theta=0.0,detail=0,
                 clusters=diabetes$id,var.link=1,model="clayton.oakes")
summary(fitco2)
fitco3 <- survival.twostage(margph,data=diabetes,theta=1.0,detail=0,
                 clusters=diabetes$id,var.link=0,model="clayton.oakes")
summary(fitco3)

### without covariates but with stratafied
marg <- phreg(Surv(time,status)~+strata(treat)+cluster(id),data=diabetes)
fitpa <- survival.twostage(marg,data=diabetes,theta=1.0,
                clusters=diabetes$id,score.method="optimize")
summary(fitpa)

fitcoa <- survival.twostage(marg,data=diabetes,theta=1.0,clusters=diabetes$id,
                 model="clayton.oakes")
summary(fitcoa)

### Piecewise constant cross hazards ratio modelling
#########################################################
```

```
d <- subset(simClaytonOakes(2000,2,0.5,0,stoptime=2,left=0),!truncated)
udp <- piecewise.twostage(c(0,0.5,2),data=d,score.method="optimize",
                          id="cluster",timevar="time",
                          status="status",model="clayton.oakes",silent=0)
summary(udp)

 ## Reduce Ex.Timings
### Same model using the strata option, a bit slower
########################################################
## makes the survival pieces for different areas in the plane
##ud1=surv.boxarea(c(0,0),c(0.5,0.5),data=d,id="cluster",timevar="time",status="status")
##ud2=surv.boxarea(c(0,0.5),c(0.5,2),data=d,id="cluster",timevar="time",status="status")
##ud3=surv.boxarea(c(0.5,0),c(2,0.5),data=d,id="cluster",timevar="time",status="status")
##ud4=surv.boxarea(c(0.5,0.5),c(2,2),data=d,id="cluster",timevar="time",status="status")

## everything done in one call
ud <- piecewise.data(c(0,0.5,2),data=d,timevar="time",status="status",id="cluster")
ud$strata <- factor(ud$strata);
ud$intstrata <- factor(ud$intstrata)

## makes strata specific id variable to identify pairs within strata
## se's computed based on the id variable across strata "cluster"
ud$idstrata <- ud$id+(as.numeric(ud$strata)-1)*2000

marg2 <- aalen(Surv(boxtime,status)~-1+factor(num):factor(intstrata),
               data=ud,n.sim=0,robust=0)
tdes <- model.matrix(~-1+factor(strata),data=ud)
fitp2 <- survival.twostage(marg2,data=ud,se.clusters=ud$cluster,clusters=ud$idstrata,
                score.method="fisher.scoring",model="clayton.oakes",
                theta.des=tdes,step=0.5)
summary(fitp2)

### now fitting the model with symmetry, i.e. strata 2 and 3 same effect
ud$stratas <- ud$strata;
ud$stratas[ud$strata=="0.5-2,0-0.5"] <- "0-0.5,0.5-2"
tdes2 <- model.matrix(~-1+factor(stratas),data=ud)
fitp3 <- survival.twostage(marg2,data=ud,clusters=ud$idstrata,se.cluster=ud$cluster,
                score.method="fisher.scoring",model="clayton.oakes",
                theta.des=tdes2,step=0.5)
summary(fitp3)

### same model using strata option, a bit slower
fitp4 <- survival.twostage(marg2,data=ud,clusters=ud$cluster,se.cluster=ud$cluster,
                score.method="fisher.scoring",model="clayton.oakes",
                theta.des=tdes2,step=0.5,strata=ud$strata)
summary(fitp4)


 ## Reduce Ex.Timings
### structured random effects model additive gamma ACE
### simulate structured two-stage additive gamma ACE model
data <- simClaytonOakes.twin.ace(4000,2,1,0,3)
```

```
out <- twin.polygen.design(data,id="cluster")
pardes <- out$pardes
pardes
des.rv <- out$des.rv
head(des.rv)
aa <- phreg(Surv(time,status)~x+cluster(cluster),data=data,robust=0)
ts <- survival.twostage(aa,data=data,clusters=data$cluster,detail=0,
       theta=c(2,1),var.link=0,step=0.5,
       random.design=des.rv,theta.des=pardes)
summary(ts)
```

---

test.conc *Concordance test Compares two concordance estimates*

Description

.. content for description (no empty lines) ..

Usage

```
test.conc(conc1, conc2, same.cluster = FALSE)
```

Arguments

conc1	Concordance estimate of group 1
conc2	Concordance estimate of group 2
same.cluster	if FALSE then groups are independent, otherwise estimates are based on same data.

Author(s)

Thomas Scheike

---

tetrachoric *Estimate parameters from odds-ratio*

Description

Calculate tetrachoric correlation of probabilities from odds-ratio

Usage

```
tetrachoric(P, OR, approx = 0, ...)
```

Arguments

P	Joint probabilities or marginals (if OR is given)
OR	Odds-ratio
approx	If TRUE an approximation of the tetrachoric correlation is used
...	Additional arguments

Examples

```
tetrachoric(0.3,1.25) # Marginal p1=p2=0.3, OR=2
P <- matrix(c(0.1,0.2,0.2,0.5),2)
prod(diag(P))/prod(lava::revdiag(P))
##mets:::assoc(P)
tetrachoric(P)
or2prob(2,0.1)
or2prob(2,c(0.1,0.2))
```

---

ttpd *ttpd discrete survival data on interval form*

Description

ttpd discrete survival data on interval form

Source

Simulated data

---

twin.clustertrunc *Estimation of twostage model with cluster truncation in bivariate situation*

Description

Estimation of twostage model with cluster truncation in bivariate situation

Usage

```
twin.clustertrunc(survformula, data = sys.parent(), theta.des = NULL,
  clusters = NULL, var.link = 1, Nit = 10, final.fitting = FALSE,
  ...)
```

Arguments

survformula	Formula with survival model aalen or cox.aalen, some limitiation on model specification due to call of fast.reshape (so for example interactions and * and : do not work here, expand prior to call)
data	Data frame
theta.des	design for dependence parameters in two-stage model
clusters	clustering variable for twins
var.link	exp link for theta
Nit	number of iteration
final.fitting	TRUE to do final estimation with SE and ... arguments for marginal models
...	Additional arguments to lower level functions

Author(s)

Thomas Scheike

Examples

```
library("timereg")
data(diabetes)
v <- diabetes$time*runif(nrow(diabetes))*rbinom(nrow(diabetes),1,0.5)
diabetes$v <- v

aout <- twin.clustertrunc(Surv(v,time,status)~1+treat+adult,
 data=diabetes,clusters="id")
aout$two        ## twostage output
par(mfrow=c(2,2))
plot(aout$marg) ## marginal model output

out <- twin.clustertrunc(Surv(v,time,status)~1+prop(treat)+prop(adult),
 data=diabetes,clusters="id")
out$two        ## twostage output
plot(out$marg) ## marginal model output
```

---

twinbmi *BMI data set*

Description

BMI data set

Format

Self-reported BMI-values on 11,411 subjects

tvparnr: twin id bmi: BMI (m/kg^2) age: Age gender: (male/female) zyg: zygosity, MZ:=mz, DZ(same sex):=dz, DZ(opposite sex):=os

twinlm *Classic twin model for quantitative traits*

Description

Fits a classical twin model for quantitative traits.

Usage

```
twinlm(formula, data, id, zyg, DZ, group = NULL, group.equal = FALSE,
  strata = NULL, weights = NULL, type = c("ace"),
  twinnum = "twinnum", binary = FALSE, ordinal = 0, keep = weights,
  estimator = NULL, constrain = TRUE, control = list(),
  messages = 1, ...)
```

Arguments

formula	Formula specifying effects of covariates on the response
data	data.frame with one observation pr row. In addition a column with the zygosity (DZ or MZ given as a factor) of each individual much be specified as well as a twin id variable giving a unique pair of numbers/factors to each twin pair
id	The name of the column in the dataset containing the twin-id variable.
zyg	The name of the column in the dataset containing the zygosity variable
DZ	Character defining the level in the zyg variable corresponding to the dyzogitic twins. If this argument is missing, the reference level (i.e. the first level) will be interpreted as the dyzogitic twins
group	Optional. Variable name defining group for interaction analysis (e.g., gender)
group.equal	If TRUE marginals of groups are asummed to be the same
strata	Strata variable name
weights	Weights matrix if needed by the chosen estimator. For use with Inverse Probability Weights
type	Character defining the type of analysis to be performed. Should be a subset of "aced" (additive genetic factors, common environmental factors, unique environmental factors, dominant genetic factors).
twinnum	The name of the column in the dataset numbering the twins (1,2). If it does not exist in data it will automatically be created.
binary	If TRUE a liability model is fitted. Note that if the right-hand-side of the formula is a factor, character vector, og logical variable, then the liability model is automatically chosen (wrapper of the bptwin function).
ordinal	If non-zero (number of bins) a liability model is fitted.
keep	Vector of variables from data that are not specified in formula, to be added to data.frame of the SEM
estimator	Choice of estimator/model

constrain	Development argument
control	Control argument parsed on to the optimization routine
messages	Control amount of messages shown
...	Additional arguments parsed on to lower-level functions

Value

Returns an object of class twinlm.

Author(s)

Klaus K. Holst

See Also

bptwin, twinlm.time, twinlm.strata, twinsim

Examples

```
## Simulate data
set.seed(1)
d <- twinsim(1000,b1=c(1,-1),b2=c(),acde=c(1,1,0,1))
## E(y|z1,z2) = z1 - z2. var(A) = var(C) = var(E) = 1

## E.g to fit the data to an ACE-model without any confounders we simply write
ace <- twinlm(y ~ 1, data=d, DZ="DZ", zyg="zyg", id="id")
ace
## An AE-model could be fitted as
ae <- twinlm(y ~ 1, data=d, DZ="DZ", zyg="zyg", id="id", type="ae")
## LRT:
lava::compare(ae,ace)
## AIC
AIC(ae)-AIC(ace)
## To adjust for the covariates we simply alter the formula statement
ace2 <- twinlm(y ~ x1+x2, data=d, DZ="DZ", zyg="zyg", id="id", type="ace")
## Summary/GOF
summary(ace2)
 ## Reduce Ex.Timings
## An interaction could be analyzed as:
ace3 <- twinlm(y ~ x1+x2 + x1:I(x2<0), data=d, DZ="DZ", zyg="zyg", id="id", type="ace")
ace3
## Categorical variables are also supported
d2 <- transform(d,x2cat=cut(x2,3,labels=c("Low","Med","High")))
ace4 <- twinlm(y ~ x1+x2cat, data=d2, DZ="DZ", zyg="zyg", id="id", type="ace")
```

`twinsim` *Simulate twin data*

Description

Simulate twin data from a linear normal ACE/ADE/AE model.

Usage

```
twinsim(nMZ = 100, nDZ = nMZ, b1 = c(), b2 = c(), mu = 0,
  acde = c(1, 1, 0, 1), randomslope = NULL, threshold = 0,
  cens = FALSE, wide = FALSE, ...)
```

Arguments

`nMZ`	Number of monozygotic twin pairs
`nDZ`	Number of dizygotic twin pairs
`b1`	Effect of covariates (labelled x1,x2,...) of type 1. One distinct covariate value for each twin/individual.
`b2`	Effect of covariates (labelled g1,g2,...) of type 2. One covariate value for each twin pair.
`mu`	Intercept parameter.
`acde`	Variance of random effects (in the order A,C,D,E)
`randomslope`	Logical indicating wether to include random slopes of the variance components w.r.t. x1,x2,...
`threshold`	Treshold used to define binary outcome y0
`cens`	Logical variable indicating whether to censor outcome
`wide`	Logical indicating if wide data format should be returned
`...`	Additional arguments parsed on to lower-level functions

Author(s)

Klaus K. Holst

See Also

`twinlm`

twinstut *Stutter data set*

Description

Based on nation-wide questionnaire answers from 33,317 Danish twins

Format

tvparnr: twin-pair id zyg: zygosity, MZ:=mz, DZ(same sex):=dz, DZ(opposite sex):=os stutter: stutter status (yes/no) age: age nr: number within twin-pair

twostageMLE *Twostage survival model fitted by pseudo MLE*

Description

Fits Clayton-Oakes clustered survival data using marginals that are on Cox form in the likelihood for the dependence parameter as in Glidden (2000). The dependence can be modelled via a

1. Regression design on dependence parameter.

We allow a regression structure for the indenpendent gamma distributed random effects and their variances that may depend on cluster covariates. So

$$\theta = h(z_j^T \alpha)$$

where z is specified by theta.des . The link function can be the exp when var.link=1

Usage

```
twostageMLE(margsurv, data = sys.parent(), theta = NULL,
  theta.des = NULL, var.link = 0, method = "NR", no.opt = FALSE,
  weights = NULL, ...)
```

Arguments

margsurv	Marginal model from phreg
data	data frame
theta	Starting values for variance components
theta.des	design for dependence parameters, when pairs are given this is could be a (pairs) x (numer of parameters) x (max number random effects) matrix
var.link	Link function for variance if 1 then uses exp link
method	type of opitmizer, default is Newton-Raphson "NR"
no.opt	to not optimize, for example to get score and iid for specific theta
weights	cluster specific weights, but given with length equivalent to data-set, weights for score equations
...	arguments to be passed to optimizer

Author(s)

Thomas Scheike

References

Measuring early or late dependence for bivariate twin data Scheike, Holst, Hjelmborg (2015), LIDA

Twostage modelling of additive gamma frailty models for survival data. Scheike and Holst, working paper

Shih and Louis (1995) Inference on the association parameter in copula models for bivariate survival data, Biometrics, (1995).

Glidden (2000), A Two-Stage estimator of the dependence parameter for the Clayton Oakes model, LIDA, (2000).

Examples

```
data(diabetes)
dd <- phreg(Surv(time,status==1)~treat+cluster(id),diabetes)
oo <- twostageMLE(dd,data=diabetes)
summary(oo)

theta.des <- model.matrix(~-1+factor(adult),diabetes)

oo <-twostageMLE(dd,data=diabetes,theta.des=theta.des)
summary(oo)
```

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