# Package 'SurrogateTest' 

March 29, 2020
Type Package
Title Early Testing for a Treatment Effect using Surrogate Marker Information

Version 1.2
Date 2020-03-26

## Author Layla Parast

Maintainer Layla Parast [parast@rand.org](mailto:parast@rand.org)
Description Provides functions to test for a treatment effect in terms of the difference in survival between a treatment group and a control group using surrogate marker information obtained at some early time point in a time-to-event outcome setting. Nonparametric kernel estimation is used to estimate the test statistic and perturbation resampling is used for variance estimation. More details will be available in the future in: Parast L, Cai T, Tian L (2019). Using a Surrogate Marker for Early Testing of a Treatment Effect. Biometrics, In press.

License GPL
Imports stats, survival
NeedsCompilation no
Repository CRAN
Date/Publication 2020-03-29 11:00:26 UTC

## $R$ topics documented:

dataA ..... 2
dataB ..... 3
delta.estimate ..... 4
design.study ..... 5
early.delta.test ..... 8
recover.B ..... 11
Index ..... 14
dataA Hypothetical Study A data

## Description

Hypothetical Study A data to be used in examples; $\mathrm{t}=1$ and the landmark time $=0.50$.

## Usage

data(dataA)

## Format

A list with 6 elements representing 1000 observations from a control group and 1000 observations from a treatment group:
s1 Surrogate marker measurement for treated observations; this marker is measured at time $=0.5$. For observations that experience the primary outcome or are censored before 0.5 , this value is NA.
$x 1$ The observed event or censoring time for treated observations; $X=\min (T, C)$ where $T$ is the time of the primary outcome and C is the censoring time.
delta1 The indicator identifying whether the treated observation was observed to have the event or was censored; $\mathrm{D}=1^{*}(\mathrm{~T}<\mathrm{C})$ where T is the time of the primary outcome and C is the censoring time.
s0 Surrogate marker measurement for control observations; this marker is measured at time $=0.5$. For observations that experience the primary outcome or are censored before 0.5 , this value is NA.
$x 0$ The observed event or censoring time for control observations; $X=\min (T, C)$ where $T$ is the time of the primary outcome and C is the censoring time.
delta0 The indicator identifying whether the control observation was observed to have the event or was censored; $\mathrm{D}=1^{*}(\mathrm{~T}<\mathrm{C})$ where T is the time of the primary outcome and C is the censoring time.

## Details

Note that if the observation is censored or experienced the primary outcome before the landmark time of 0.50 , the surrogate marker measurement is not observed and coded NA.

## Examples

```
data(dataA)
```

names(dataA)

```
dataB Hypothetical Study B data
```


## Description

Hypothetical Study B data to be used in examples; landmark time $=0.50$.

## Usage

```
    data(dataB)
```


## Format

A list with 6 elements representing 800 observations from a control group and 800 observations from a treatment group:
s1 Surrogate marker measurement for treated observations; this marker is measured at time $=0.5$. For observations that experience the primary outcome or are censored before 0.5 , this value is NA.
$x 1$ The observed event or censoring time for treated observations; $\mathrm{X}=\min (\mathrm{T}, \mathrm{C})$ where T is the time of the primary outcome and C is the censoring time. This time is administratively censored at 0.55 (see details).
delta1 The indicator identifying whether the treated observation was observed to have the event or was censored; $\mathrm{D}=1^{*}(\mathrm{~T}<\mathrm{C})$ where T is the time of the primary outcome and C is the censoring time.
$s 0$ Surrogate marker measurement for control observations; this marker is measured at time $=0.5$. For observations that experience the primary outcome or are censored before 0.5 , this value is NA.
$x 0$ The observed event or censoring time for control observations; $X=\min (T, C)$ where $T$ is the time of the primary outcome and C is the censoring time. This time is administratively censored at 0.55 (see details).
delta0 The indicator identifying whether the control observation was observed to have the event or was censored; $\mathrm{D}=1 *(\mathrm{~T}<\mathrm{C})$ where T is the time of the primary outcome and C is the censoring time.

## Details

Note that if the observation is censored or experienced the primary outcome before the landmark time of 0.50 , the surrogate marker measurement is not observed and coded NA. In addition, Study B data is only observed up to the landmark time plus some epsilon, here epsilon=0.05 such that all observations are essentially adminstratively censored at time $=0.55$.

## Examples

```
data(dataB)
names(dataB)
```


## Description

This function calculates the treatment effect in the survival setting i.e. the difference in survival at time $t$ between the treatment group and the control group. The inverse probability of censoring weighted estimate of survival within each treatment group is used; there is an option to use the Kaplan-Meier estimate instead. This function is generally not expected to be used directly by the user, it is called by the recover.B function.

## Usage

delta.estimate(xone, xzero, deltaone, deltazero, $t$, weight = NULL, KM = FALSE)

## Arguments

xone numeric vector, the observed event times in the treatment group, $X=\min (T, C)$ where T is the time of the primary outcome and C is the censoring time.
xzero numeric vector, the observed event times in the control group, $X=\min (T, C)$ where T is the time of the primary outcome and C is the censoring time.
deltaone numeric vector, the event indicators for the treatment group, $\mathrm{D}=\mathrm{I}(\mathrm{T}<\mathrm{C})$ where T is the time of the primary outcome and C is the censoring time.
deltazero numeric vector, the event indicators for the control group, $\mathrm{D}=\mathrm{I}(\mathrm{T}<\mathrm{C})$ where T is the time of the primary outcome and C is the censoring time.
t
the time of interest.
weight a $\quad$ n $+n_{0}$ by $x$ matrix of weights where $n_{1}=$ sample size in the treatment group and $n_{0}=$ sample size in the control group, default is null; generally not supplied by user, only used by other functions.

KM true or false, indicating whether the Kaplan-Meier estimate of survival should be used instead of the inverse probability of censoring weighted estimate

## Value

the difference in survival at time $t$ (treatment group minus control group)

## Author(s)

Layla Parast

## Examples

```
data(dataA)
delta.estimate(xone = dataA$x1, xzero = dataA$x0, deltaone = dataA$delta1, deltazero =
dataA$delta0, t=1)
delta.estimate(xone = dataA$x1, xzero = dataA$x0, deltaone = dataA$delta1, deltazero =
dataA$delta0, t=0.5)
```

design.study Power and sample size calculation for designing a future study

## Description

Power and sample size calculation for designing a future study

## Usage

design.study(Axzero, Adeltazero, Aszero, Axone = NULL, Adeltaone = NULL, Asone = NULL, delta.ea $=$ NULL, psi = NULL, R.A.given = NULL, t, landmark, extrapolate $=$ T, adjustment $=\mathrm{F}, \mathrm{n}=$ NULL, power $=$ NULL, pi. $1=0.5$, pi. $0=0.5$, cens.rate, transform = F)

## Arguments

Axzero observed event times in the control group in Study A
Adeltazero event/censoring indicators in the control group in Study A
Aszero surrogate marker values in the control group in Study A, NA for individuals not observable at the time the surrogate marker was measured
Axone observed event times in the treatment group in Study A; optional (user must provide either (1) data from treatment arm in Study A or (2) hypothesized values for delta.ea (or R.A.given)and psi or (3) data from treatment arm in Study A and hypothesized psi (if different from observed treatment effect at $t$ in Study A))
Adeltaone event/censoring indicators in the treatment group in Study A; optional (user must provide either (1) data from treatment arm in Study A or (2) hypothesized values for delta.ea (or R.A.given)and psi or (3) data from treatment arm in Study A and hypothesized psi (if different from observed treatment effect at $t$ in Study A))
Asone surrogate marker values in the treatment group in Study A, NA for individuals not observable at the time the surrogate marker was measured; optional (user must provide either (1) data from treatment arm in Study A or (2) hypothesized values for delta.ea (or R.A.given) and psi or (3) data from treatment arm in Study A and hypothesized psi (if different from observed treatment effect at t in Study A))
delta.ea hypothesized value for the early treatment effect at time t0; optional (user must provide either (1) data from treatment arm in Study A or (2) hypothesized values for delta.ea (or R.A.given) and psi or (3) data from treatment arm in Study A and
hypothesized psi (if different from observed treatment effect at $t$ in Study A)), if not given then it is assumed that this quantity equals the osberved early treatment effect at time t0 in Study A
psi hypothesized value for the treatment effect at time $t$; optional (user must provide either (1) data from treatment arm in Study A or (2) hypothesized values for delta.ea (or R.A.given) and psi or (3) data from treatment arm in Study A and hypothesized psi (if different from observed treatment effect at $t$ in Study A)), if not given then it is assumed that this quantity equals the osberved treatment effect at time t in Study A
R.A.given hypothesized value for the proportion of treatment effect on the primary outcome explained by surrogate information at t0 in Study A; optional (user must provide either (1) data from treatment arm in Study A or (2) hypothesized values for delta.ea (or R.A.given) and psi or (3) data from treatment arm in Study A and hypothesized psi (if different from observed treatment effect at t in Study A))
t
time of interest
landmark landmark time of interest, t0
extrapolate TRUE or FALSE; indicates whether local constant extrapolation should be used, default is TRUE
adjustment TRUE or FALSE; indicates whether adjustment that is needed when survival past time t is high should be used, default is FALSE if survival past t 0 is $<0.90$ in both arms arm of Study A, otherwise default is true if survival past t 0 is $>=$ 0.90 in either arm of Study A
n total sample size for future study (Study B); optional (user needs to provide either n or power)
power desired power for testing at time t0 for future study (Study B); optional (user needs to provide either n or power)
pi. 1 proportion of total sample size in future study (Study B) that would be assigned to the treatment group, default is 0.5
pi. $0 \quad$ proportion of total sample size in future study (Study B) that would be assigned to the treatment group, default is 0.5
cens.rate censoring in the future study (Study B) is assumed to follow an exponential distribution with censoring rate equal to this specificed value
transform TRUE or FALSE; indicates whether a transformation should be used, default is FALSE.

## Details

Assume information is available on a prior study, Study A, examining the effectiveness of a treatment up to some time of interest, t . The aim is to plan a future study, Study B, that would be conducted only up to time $t_{0}<t$ and a test for a treatment effect would occur at $t_{0}$. In both studies, we assume a surrogate marker is/will be measured at time $t_{0}$ for individuals still observable at $t_{0}$. Let $G$ be the binary treatment indicator with $G=1$ for treatment and $G=0$ for control and we assume throughout that subjects are randomly assigned to a treatment group at baseline. Let $T_{K}^{(1)}$ and $T_{K}^{(0)}$ denote the time of the primary outcome of interest, death for example, under the treatment and
under the control, respectively, in Study K. Let $S_{K}^{(1)}$ and $S_{K}^{(0)}$ denote the surrogate marker measured at time $t_{0}$ under the treatment and the control, respectively, in Study K.
The null and alternative hypotheses of interest are:

$$
\begin{gathered}
H_{0}: \Delta_{B}(t) \equiv P\left(T_{B}^{(1)}>t\right)-P\left(T_{B}^{(0)}>t\right)=0 \\
H_{1}: \Delta_{B}(t)=\psi>0
\end{gathered}
$$

Here, we plan to test $H_{0}$ in Study B using the test statistic

$$
Z_{E B}\left(t, t_{0}\right)=\sqrt{n_{B}} \frac{\hat{\Delta}_{E B}\left(t, t_{0}\right)}{\hat{\sigma}_{E B}\left(t, t_{0}\right)}
$$

(see early.delta.test documentation). The estimated power at a type I error rate of 0.05 is thus

$$
1-\Phi\left\{1.96-\frac{\sqrt{n_{B}} \hat{R}_{S A}\left(t, t_{0}\right) \psi}{\hat{\sigma}_{E B 0}\left(t, t_{0} \mid \hat{r}_{A}^{(0)}, W_{B}^{C}\right)}\right\}
$$

where $\hat{R}_{S A}\left(t, t_{0}\right)=\hat{\Delta}_{E A}\left(t, t_{0}\right) / \hat{\Delta}_{A}(t)$, and

$$
\hat{\Delta}_{A}(t)=n_{A 1}^{-1} \sum_{i=1}^{n_{A 1}} \frac{I\left(X_{A i}^{(1)}>t\right)}{\hat{W}_{A 1}^{C}(t)}-n_{A 0}^{-1} \sum_{i=1}^{n_{A 0}} \frac{I\left(X_{A i}^{(0)}>t\right)}{\hat{W}_{A 0}^{C}(t)}
$$

and $\hat{\Delta}_{E A}\left(t, t_{0}\right)$ is parallel to $\hat{\Delta}_{E B}\left(t, t_{0}\right)$ except replacing $n_{A 0}^{-1} \sum_{i=1}^{n_{A 0}} \hat{r}_{A}^{(0)}\left(t \mid S_{A i}^{(0)}, t_{0}\right) \frac{I\left(X_{A i}^{(0)}>t_{0}\right)}{\hat{W}_{A 0}^{C}\left(t_{0}\right)}$ by $n_{A 0}^{-1} \sum_{i=1}^{n_{A 0}} \hat{W}_{A 0}^{C}(t)^{-1} I\left(X_{A i}^{(0)}>t\right)$, and $\hat{W}_{A g}^{C}(\cdot)$ is the Kaplan-Meier estimator of the survival function for $C_{A}^{(g)}$ for $g=0,1$. In addition, $\hat{\sigma}_{E B 0}\left(t, t_{0} \mid \hat{r}_{A}^{(0)}, W_{B}^{C}\right)^{2}=$

$$
\frac{1}{\pi_{B 0} \pi_{B 1}}\left[\frac{\hat{\mu}_{A B 2}^{(0)}\left(t, t_{0}, \mid \hat{r}_{A}^{(0)}\right)}{W_{B}^{C}\left(t_{0}\right)}-\hat{\mu}_{A B 1}^{(0)}\left(t, t_{0}, \mid \hat{r}_{A}^{(0)}\right)^{2}\left\{1+\int_{0}^{t_{0}} \frac{\lambda_{B}^{C}(u) d u}{\hat{W}_{A 0}^{T}(u) W_{B}^{C}(u)}\right\}\right]
$$

assuming that the survival function of the censoring distribution is $W_{B}^{C}(t)$ in both arms, where $\pi_{B g}=n_{B g} / n_{B}$ and $\hat{W}_{A 0}^{T}(\cdot)$ is the Kaplan-Meier estimator of the survival function of $T_{A}^{(0)}$ based on the observations from Study A, and

$$
\hat{\mu}_{A B m}^{(0)}\left(t, t_{0}, \mid \hat{r}_{A}^{(0)}\right)=n_{A 0}^{-1} \sum_{i=1}^{n_{A 0}} \frac{\hat{r}_{A}^{(0)}\left(t \mid S_{A i}^{(0)}, t_{0}\right)^{m} I\left(X_{A i}^{(0)}>t_{0}\right)}{\hat{W}_{A 0}^{C}\left(t_{0}\right)}
$$

where $\hat{r}_{A}^{(0)}\left(t \mid s, t_{0}\right)$ is provided in the early.delta.test documentation.
This can be re-arranged to calculate the sample size needed in Study B to achieve a power of $100(1-\beta) \%$ :

$$
n_{B}=\left\{\hat{\sigma}_{E B 0}\left(t, t_{0} \mid \hat{r}_{A}^{(0)}, W_{B}^{C}\right)\left(\frac{1.96-\Phi^{-1}(\beta)}{\hat{R}_{S A}\left(t, t_{0}\right) \psi}\right)\right\}^{2}
$$

When the outcome rate is low (i.e., survival rate at t is high), an adjustment to the variance calculation is needed. This is automatically implemented if the survival rate at $t$ in either arm is 0.90 or higher.

## Value

n
Total sample size needed for Study B at the given power (if power is provided by user).
power $\quad$ Estimated power for Study B at the given sample size (if sample size is provided by user).

## Author(s)

Layla Parast

## References

Parast L, Cai T, Tian L (2019). Using a Surrogate Marker for Early Testing of a Treatment Effect. Biometrics, In press.

## Examples

```
data(dataA)
design.study(Axzero = dataA$x0, Adeltazero = dataA$delta0, Aszero = dataA$s0,
Axone = dataA$x1, Adeltaone = dataA$delta1, Asone = dataA$s1, t=1, landmark=0.5,
power = 0.80, cens.rate=0.5)
design.study(Axzero = dataA$x0, Adeltazero = dataA$delta0, Aszero = dataA$s0,
Axone = dataA$x1, Adeltaone = dataA$delta1, Asone = dataA$s1, t=1, landmark=0.5,
n=2500, cens.rate=0.5)
design.study(Axzero = dataA$x0, Adeltazero = dataA$delta0, Aszero = dataA$s0,
Axone = dataA$x1, Adeltaone = dataA$delta1, Asone = dataA$s1, t=1, landmark=0.5,
power = 0.80, cens.rate=0.5, psi = 0.05)
```

early.delta.test Estimate and test the early treatment effect

## Description

Estimates the early treatment effect estimate and provides two versions of the standard error; tests the null hypothesis that this treatment effect is equal to 0

## Usage

early.delta.test(Axzero, Adeltazero, Aszero, Bxzero, Bdeltazero, Bszero, Bxone, Bdeltaone, Bsone, t, landmark, perturb = T, extrapolate = T, transform = F)

## Arguments

| Axzero | observed event times in the control group in Study A |
| :---: | :---: |
| Adeltazero | event/censoring indicators in the control group in Study A |
| Aszero | surrogate marker values in the control group in Study A, NA for individuals not observable at the time the surrogate marker was measured |
| Bxzero | observed event times in the control group in Study B |
| Bdeltazero | event/censoring indicators in the control group in Study B |
| Bszero | surrogate marker values in the control group in Study B, NA for individuals not observable at the time the surrogate marker was measured |
| Bxone | observed event times in the treatment group in Study B |
| Bdeltaone | event/censoring indicators in the treatment group in Study B |
| Bsone | surrogate marker values in the treatment group in Study B, NA for individuals not observable at the time the surrogate marker was measured |
| t | time of interest |
| landmark | landmark time of interest, t0 |
| perturb | TRUE or FALSE; indicates whether the standard error estimate obtained using perturbation resampling should be calculated |
| extrapolate | TRUE or FALSE; indicates whether local constant extrapolation should be used, default is TRUE |
| transform | TRUE or FALSE; indicates whether a transformation should be used, default is FALSE. |

## Details

Assume there are two randomized studies of a treatment effect, a prior study (Study A) and a current study (Study B). Study A was completed up to some time t , while Study B was stopped at time $t_{0}<t$. In both studies, a surrogate marker was measured at time $t_{0}$ for individuals still observable at $t_{0}$. Let $G$ be the binary treatment indicator with $G=1$ for treatment and $G=0$ for control and we assume throughout that subjects are randomly assigned to a treatment group at baseline. Let $T_{K}^{(1)}$ and $T_{K}^{(0)}$ denote the time of the primary outcome of interest, death for example, under the treatment and under the control, respectively, in Study K. Let $S_{K}^{(1)}$ and $S_{K}^{(0)}$ denote the surrogate marker measured at time $t_{0}$ under the treatment and the control, respectively, in Study K.
The treatment effect quantity of interest, $\Delta_{K}(t)$, is the difference in survival rates by time $t$ under treatment versus control,

$$
\Delta_{K}(t)=E\left\{I\left(T_{K}^{(1)}>t\right)\right\}-E\left\{I\left(T_{K}^{(0)}>t\right)\right\}=P\left(T_{K}^{(1)}>t\right)-P\left(T_{K}^{(0)}>t\right)
$$

where $t>t_{0}$. Here, we estimate an early treatment effect quantity using surrogate marker information defined as,

$$
\Delta_{E B}\left(t, t_{0}\right)=P\left(T_{B}^{(1)}>t_{0}\right) \int r\left(t \mid s, t_{0}\right) d F_{B}^{(1)}\left(s \mid t_{0}\right)-P\left(T_{B}^{(0)}>t_{0}\right) \int r\left(t \mid s, t_{0}\right) d F_{B}^{(0)}\left(s \mid t_{0}\right)
$$

where $r\left(t \mid s, t_{0}\right)=P\left(T_{A}^{(0)}>t \mid T_{A}^{(0)}>t_{0}, S_{A}^{(0)}=s\right)$ and $F_{B}^{(g)}\left(s \mid t_{0}\right)=P\left(S_{B}^{(g)} \leq s \mid T_{B}^{(g)}>t_{0}\right)$.

To test the null hypothesis that $\Delta_{B}(t)=0$, we test the null hypothesis $\Delta_{E B}\left(t, t_{0}\right)=0$ using the test statistic

$$
Z_{E B}\left(t, t_{0}\right)=\sqrt{n_{B}} \frac{\hat{\Delta}_{E B}\left(t, t_{0}\right)}{\hat{\sigma}_{E B}\left(t, t_{0}\right)}
$$

where $\hat{\Delta}_{E B}\left(t, t_{0}\right)$ is a consistent estimate of $\Delta_{E B}\left(t, t_{0}\right)$ and $\hat{\sigma}_{E B}\left(t, t_{0}\right)$ is the estimated standard error of $\sqrt{n_{B}}\left\{\hat{\Delta}_{E B}\left(t, t_{0}\right)-\Delta_{E B}\left(t, t_{0}\right)\right\}$. We reject the null hypothesis when $\left|Z_{E B}\left(t, t_{0}\right)\right|>$ $\Phi^{-1}(1-\alpha / 2)$ where $\alpha$ is the Type 1 error rate.
To obtain $\hat{\Delta}_{E B}\left(t, t_{0}\right)$, we use

$$
\hat{\Delta}_{E B}\left(t, t_{0}\right)=n_{B 1}^{-1} \sum_{i=1}^{n_{B 1}} \hat{r}_{A}^{(0)}\left(t \mid S_{B i}^{(1)}, t_{0}\right) \frac{I\left(X_{B i}^{(1)}>t_{0}\right)}{\hat{W}_{B 1}^{C}\left(t_{0}\right)}-n_{B 0}^{-1} \sum_{i=1}^{n_{B 0}} \hat{r}_{A}^{(0)}\left(t \mid S_{B i}^{(0)}, t_{0}\right) \frac{I\left(X_{B i}^{(0)}>t_{0}\right)}{\hat{W}_{B 0}^{C}\left(t_{0}\right)}
$$

where $\hat{W}_{k g}^{C}(u)$ is the Kaplan-Meier estimator of $W_{k g}^{C}(u)=P\left(C_{k}^{(g)}>u\right)$ and $\hat{r}_{A}^{(0)}\left(t \mid s, t_{0}\right)=$ $\exp \left\{-\hat{\Lambda}_{A}^{(0)}\left(t \mid s, t_{0}\right)\right\}$, where

$$
\hat{\Lambda}_{A}^{(0)}\left(t \mid t_{0}, s\right)=\int_{t_{0}}^{t} \frac{\sum_{i=1}^{n_{A 0}} I\left(X_{A i}^{(0)}>t_{0}\right) K_{h}\left\{\gamma\left(S_{A i}^{(0)}\right)-\gamma(s)\right\} d N_{A i}^{(0)}(z)}{\sum_{i=1}^{n_{A 0}} K_{h}\left\{\gamma\left(S_{A i}^{(0)}\right)-\gamma(s)\right\} Y_{A i}^{(0)}(z)}
$$

is a consistent estimate of $\Lambda_{A}^{(0)}\left(t \mid t_{0}, s\right)=-\log \left[r_{A}^{(0)}\left(t \mid t_{0}, s\right)\right], Y_{A i}^{(0)}(t)=I\left(X_{A i}^{(0)} \geq t\right)$, $N_{A i}^{(0)}(t)=I\left(X_{A i}^{(0)} \leq t\right) \delta_{A i}^{(0)}, K(\cdot)$ is a smooth symmetric density function, $K_{h}(x)=K(x / h) / h$ and $\gamma(\cdot)$ is a given monotone transformation function. For the bandwidth $h$, we require the standard undersmoothing assumption of $h=O\left(n_{g}^{-\gamma}\right)$ with $\gamma \in(1 / 4,1 / 2)$ in order to eliminate the impact of the bias of the conditional survival function on the resulting estimator.
The quantity $\hat{\sigma}_{E B}\left(t, t_{0}\right)$ is obtained using either a closed form expression under the null or a perturbation resampling approach. If a confidence interval is desired, perturbation resampling is required.

## Value

| delta.eb | The estimate early treatment effect, $\hat{\Delta}_{E B}\left(t, t_{0}\right)$. |
| :--- | :--- |
| se.closed | The standard error estimate of the early treatment effect using the closed form <br> expression under the null. |
|  | Z.closed |
| p.value.closed | The p-value using the closed form standard error expression. |
| conf.closed.norm |  |

The confidence interval for the early treatment effect, using a normal approximation and using the closed form standard error expression.
se.perturb The standard error estimate of the early treatment effect using perturbation resampling, if perturb $=T$.
Z. perturb $\quad$ The test statistic using the perturbed standard error estimate, if perturb $=T$.
p.value.perturb

The p-value using the perturbed standard error estimate, if perturb $=\mathrm{T}$.
conf.perturb.norm
The confidence interval for the early treatment effect, using a normal approximation and using the perturbed standard error expression, if perturb $=\mathrm{T}$.
delta.eb.CI The confidence interval for the early treatment effect, using the quantiles of the perturbed estimates, if perturb $=\mathrm{T}$.

## Author(s)

Layla Parast

## References

Parast L, Cai T, Tian L (2019). Using a Surrogate Marker for Early Testing of a Treatment Effect. Biometrics, In press.

## Examples

```
data(dataA)
data(dataB)
early.delta.test(Axzero = dataA$x0, Adeltazero = dataA$delta0, Aszero = dataA$s0,
Bxzero = dataB$x0, Bdeltazero = dataB$delta0, Bszero = dataB$s0, Bxone = dataB$x1,
Bdeltaone = dataB$delta1, Bsone = dataB$s1, t=1, landmark=0.5, perturb = FALSE,
extrapolate = TRUE)
early.delta.test(Axzero = dataA$x0, Adeltazero = dataA$delta0, Aszero = dataA$s0,
Bxzero = dataB$\times0, Bdeltazero = dataB$delta0, Bszero = dataB$s0, Bxone = dataB$x1,
Bdeltaone = dataB$delta1, Bsone = dataB$s1, t=0.75, landmark=0.5, perturb = FALSE,
extrapolate = TRUE)
early.delta.test(Axzero = dataA$x0, Adeltazero = dataA$delta0, Aszero = dataA$s0,
Bxzero = dataB$\times0, Bdeltazero = dataB$delta0, Bszero = dataB$s0, Bxone = dataB$x1,
Bdeltaone = dataB$delta1, Bsone = dataB$s1, t=1, landmark=0.5, perturb = TRUE,
extrapolate = TRUE)
```

recover.B

Recover an estimate of the treatment effect at time tin Study B

## Description

Recover an estimate of the treatment effect at time $t$ in Study B

## Usage

recover. B (Axzero, Adeltazero, Aszero, Axone, Adeltaone, Asone, Bxzero, Bdeltazero, Bszero, Bxone, Bdeltaone, Bsone, t, landmark, extrapolate = T, transform = F)

## Arguments

Axzero observed event times in the control group in Study A
Adeltazero
event/censoring indicators in the control group in Study A
Aszero surrogate marker values in the control group in Study A, NA for individuals not observable at the time the surrogate marker was measured

| Axone | observed event times in the treatment group in Study A <br> event/censoring indicators in the treatment group in Study A <br> Adeltaone <br> surrogate marker values in the treatment group in Study A, NA for individuals <br> not observable at the time the surrogate marker was measured <br> observed event times in the control group in Study B |
| :--- | :--- |
| Bxzero |  |
| Bdeltazero | event/censoring indicators in the control group in Study B <br> surrogate marker values in the control group in Study B, NA for individuals not <br> observable at the time the surrogate marker was measured |
| Bszero | observed event times in the treatment group in Study B |
| Bxone |  |
| Bdeltaone | event/censoring indicators in the treatment group in Study B |
| Bsone | surrogate marker values in the treatment group in Study B, NA for individuals <br> not observable at the time the surrogate marker was measured <br> time of interest |
| t landmark | landmark time of interest, t0 |
| extrapolate | TRUE or FALSE; indicates whether local constant extrapolation should be used, <br> default is TRUE |
| transform | TRUE or FALSE; indicates whether a transformation should be used, default is |
| FALSE |  |

## Details

Assume there are two randomized studies of a treatment effect, a prior study (Study A) and a current study (Study B). Study A was completed up to some time t, while Study B was stopped at time $t_{0}<t$. In both studies, a surrogate marker was measured at time $t_{0}$ for individuals still observable at $t_{0}$. Let $G$ be the binary treatment indicator with $G=1$ for treatment and $G=0$ for control and we assume throughout that subjects are randomly assigned to a treatment group at baseline. Let $T_{K}^{(1)}$ and $T_{K}^{(0)}$ denote the time of the primary outcome of interest, death for example, under the treatment and under the control, respectively, in Study K. Let $S_{K}^{(1)}$ and $S_{K}^{(0)}$ denote the surrogate marker measured at time $t_{0}$ under the treatment and the control, respectively, in Study K.
The treatment effect quantity of interest, $\Delta_{K}(t)$, is the difference in survival rates by time $t$ under treatment versus control,

$$
\Delta_{K}(t)=E\left\{I\left(T_{K}^{(1)}>t\right)\right\}-E\left\{I\left(T_{K}^{(0)}>t\right)\right\}=P\left(T_{K}^{(1)}>t\right)-P\left(T_{K}^{(0)}>t\right)
$$

where $t>t_{0}$. Here, we recover an estimate of $\Delta_{B}(t)$ using Study B information (which stopped follow-up at time $t_{0}<t$ ) and Study A information (which has follow-up information through time $\mathrm{t})$. The estimate is obtained as

$$
\hat{\Delta}_{E B}\left(t, t_{0}\right) / \hat{R}_{S A}\left(t, t_{0}\right)
$$

where $\hat{\Delta}_{E B}\left(t, t_{0}\right)$ is the early treatment effect estimate in Study B, described in the early.delta.test documention, and $\hat{R}_{S A}\left(t, t_{0}\right)$ is the proportion of treatment effect explained by the surrogate marker information at $t_{0}$ in Study A. This proportion is calculated as $\hat{R}_{S A}\left(t, t_{0}\right)=\hat{\Delta}_{E A}\left(t, t_{0}\right) / \hat{\Delta}_{A}(t)$ where

$$
\hat{\Delta}_{A}(t)=n_{A 1}^{-1} \sum_{i=1}^{n_{A 1}} \frac{I\left(X_{A i}^{(1)}>t\right)}{\hat{W}_{A 1}^{C}(t)}-n_{A 0}^{-1} \sum_{i=1}^{n_{A 0}} \frac{I\left(X_{A i}^{(0)}>t\right)}{\hat{W}_{A 0}^{C}(t)}
$$

and $\hat{\Delta}_{E A}\left(t, t_{0}\right)$ is parallel to $\hat{\Delta}_{E B}\left(t, t_{0}\right)$ except replacing $n_{A 0}^{-1} \sum_{i=1}^{n_{A 0}} \hat{r}_{A}^{(0)}\left(t \mid S_{A i}^{(0)}, t_{0}\right) \frac{I\left(X_{A i}^{(0)}>t_{0}\right)}{\hat{W}_{A 0}^{C}\left(t_{0}\right)}$ by $n_{A 0}^{-1} \sum_{i=1}^{n_{A 0}} \hat{W}_{A 0}^{C}(t)^{-1} I\left(X_{A i}^{(0)}>t\right)$, and $\hat{W}_{A g}^{C}(\cdot)$ is the Kaplan-Meier estimator of the survival function for $C_{A}^{(g)}$ for $g=0,1$.
Perturbation resampling is used to provide a standard error estimate for the estimate of $\Delta_{B}(t)$ and a confidence interval.

## Value

recovered.deltaB
The recovered estimate of $\Delta_{B}(t)$.
sd.recovered.deltaB
The standard error estimate of the recovered estimate of $\Delta_{B}(t)$.
conf.quantile.recovered.deltaB
A confidence interval for the recovered estimate of $\Delta_{B}(t)$.

## Author(s)

Layla Parast

## References

Parast L, Cai T, Tian L (2019). Using a Surrogate Marker for Early Testing of a Treatment Effect. Biometrics, In press.
Parast L, Cai T and Tian L (2017). Evaluating Surrogate Marker Information using Censored Data. Statistics in Medicine, 36(11): 1767-1782.

## Examples

```
data(dataA)
data(dataB)
recover.B(Axzero = dataA$x0, Adeltazero = dataA$delta0, Aszero = dataA$s0, Axone
= dataA$x1, Adeltaone = dataA$delta1, Asone = dataA$s1, Bxzero = dataB$x0, Bdeltazero
= dataB$delta0, Bszero = dataB$s0, Bxone = dataB$x1, Bdeltaone = dataB$delta1, Bsone
= dataB$s1, t=1, landmark=0.5, extrapolate = TRUE)
recover.B(Axzero = dataA$x0, Adeltazero = dataA$delta0, Aszero = dataA$s0, Axone
= dataA$x1, Adeltaone = dataA$delta1, Asone = dataA$s1, Bxzero = dataB$x0, Bdeltazero
= dataB$delta0, Bszero = dataB$s0, Bxone = dataB$x1, Bdeltaone = dataB$delta1, Bsone
= dataB$s1, t=0.75, landmark=0.5, extrapolate = TRUE)
```


## Index

```
*Topic datasets
    dataA, 2
    dataB, 3
*Topic htest
    design.study,5
    early.delta.test, }
*Topic nonparametric
    delta.estimate, 4
    design.study,5
    early.delta.test, }
    recover.B,11
*Topic Survival
    delta.estimate,4
    design.study,5
    early.delta.test, 8
    recover.B,11
*Topic univar
    delta.estimate,4
dataA, 2
dataB, 3
delta.estimate,4
design.study,5
early.delta.test, }
recover.B,11
```

